

1 **Title (a maximum of 20 words)**

2 Examining the bias of pitfall traps with enclosure experiments and removal sampling

3

4 **Running title (a maximum of 40 characters including spaces)**

5 Bias of pitfall traps

6

7 **Abstract (249/250 words)**

- 8 1. Pitfall trapping is among the most widely used sampling methods for ground-  
9 dwelling beetles. However, the sampled ground-dwelling beetle abundance in  
10 pitfall traps may be biased, which can lead to difficulties when comparing the  
11 sampled abundance between different habitats.
- 12 2. To better understand the comparability of the sampled abundance, we focused on  
13 two sampling processes: a temporary migration process (referring to the individual  
14 movements into and out of the effective trap area) and a removal process  
15 (determining the number of individuals collected given the number of individuals  
16 within the trap area).
- 17 3. To examine the effects of the temporary migration process, we compared the  
18 sampled abundance of ten ground-dwelling beetle species inside and outside  
19 enclosures in two habitats (clear-cut and forested sites) and found that installing the  
20 enclosure decreased the sampled abundance in both habitats, but the effects did not  
21 differ among habitats.
- 22 4. To examine the effects of the removal process, we examined the effects of daily soil  
23 temperature on the daily sampled abundance (removal probability) of *Pterostichus*  
24 *thunbergi* using removal sampling within enclosures. We found that soil

25 temperature increased the removal probability, but this effect was limited and not  
26 consistent across survey periods.

27 5. The results suggest that the effects of the temporary migration process may not be  
28 habitat dependent and that the effects of the removal process may be small.

29 Therefore, the traditional use of pitfall traps is, to some extent, a robust and  
30 comparable measure of the sampled abundance of ground-dwelling beetles among  
31 different habitats.

32

33 **Keywords (maximum of 10 keywords)**

34 activity density, Carabidae, catchability, detection probability, forest, population density,  
35 sampling bias, soil temperature, trap efficiency

36

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52 **5551 Words (abstract: 249, references: 1345)**

53 (Max 7000 words without Abstract and References - Insect Conservation and Diversity)

54

55 **Introduction**

56 Pitfall trapping is one of the most widely used techniques for ground  
57 invertebrate sampling (Woodcock 2005; Henderson & Southwood 2016). As such,  
58 pitfall trap sampling has been used in various studies of ground-dwelling beetles,  
59 including the examination of their regional community and habitat preferences (e.g.,  
60 Niemela & Halme 1992; Jouveau *et al.* 2019) and the anthropogenic impact on this  
61 group (e.g., Koivula *et al.* 2019; Yamanaka *et al.* 2021). However, the abundance  
62 sampled using pitfall traps is determined through the sampling process, which is  
63 influenced by several factors, including the sampling design and targeted site- and  
64 species-specific characteristics. Sampled abundance, therefore, does not equal the “true”  
65 abundance existing in a certain area during an arbitrary time period (hereafter, initial  
66 abundance; see Table 1 for glossary), and it should be interpreted with caution when  
67 comparing sampled abundance among different sampling designs and sites and periods  
68 with different environmental conditions (Adis 1979; Lang 2000; Woodcock 2005). For  
69 example, if the catchability of pitfall traps differs greatly between sampling designs,  
70 sites, or periods, the sampled abundance will be determined by the differences in the  
71 catchability rather than differences in the initial abundance, making it difficult to infer  
72 differences in the initial abundance between sites and periods.

73 The factors influencing the catchability of pitfall traps can be divided into two  
74 groups: (1) sampling design and (2) site- and species-specific characteristics that affect  
75 the movement activity of target species. There have been many studies on the effects of  
76 sampling design on catchability, for example, trap design (Luff 1975; Koivula *et al.*  
77 2003), preservation liquid types (Koivula *et al.* 2003), the number and arrangement of  
78 traps installed (Lövei & Magura 2011; Engel *et al.* 2017), and the sampling period

79 (Schirmel *et al.* 2010; Lövei & Magura 2011). The differences in the sampling design  
80 can lead to differences in the sampled abundance even if the initial abundance is the  
81 same between sites. However, this problem can be addressed by using the same  
82 sampling method across target sites, that is, by disseminating a standard design (Brown  
83 & Matthews 2016).

84 For factors influencing catchability through changes in the movement activity  
85 of ground-dwelling beetles, there are extrinsic factors, such as temperature (Honek  
86 1988; Engel *et al.* 2017), vegetation complexity (Honek 1988; Lang 2000; Koivula *et*  
87 *al.* 2003), and soil disturbances (e.g., digging-in effects; Schirmel *et al.* 2010). In  
88 addition, there are intrinsic factors, such as body size (Lang 2000; Engel *et al.* 2017),  
89 habitat preferences (Koivula *et al.* 2003; Růžicková *et al.* 2021), seasonal  
90 abundance/behavior changes (Baars 1979), and the hunger level of individuals (Raworth  
91 & Choi 2001). The influence of these factors can be expected to vary widely among  
92 target sites and species. Therefore, it is difficult to control differences in sampled  
93 abundance caused by changes in movement activity, which can become a serious  
94 problem, particularly when comparing the sampled abundance between different sites  
95 and/or periods. For example, even if the initial abundance is the same in two habitats,  
96 such as forest and grassland, the sampled abundance may be different between the  
97 habitats if the movement activity is higher in one habitat than the other. Similarly,  
98 differences in the movement activity may mask differences in the initial abundance even  
99 if the abundance truly differs between the two habitats.

100 In this study, to better understand the influence of differences in the movement  
101 activity of ground-dwelling beetles on the comparability of the sampled abundance, we  
102 deconstruct the pitfall trap sampling process into two subprocesses, namely, (i)

103 temporary migration and (ii) removal processes (Iknayan *et al.* 2014; Yamaura & Royle  
104 2017). The temporary migration process is defined as individual movement into and out  
105 of the effective trap area. In this process, the movement activity increases/decreases the  
106 number of individuals moving into/out of the trap area. Differences in the movement  
107 activity between sites can lead to differences in the potential number of individuals who  
108 can be caught by traps between sites. The removal process determines the number of  
109 individuals collected based on the removal probability (Table 1) of the traps given the  
110 number of individuals exposed to the sampling within the trapping area. In this process,  
111 the movement activity increases/decreases the removal probability, and differences in  
112 the movement activity between sites can change the sampled abundance of each site.

113         The purpose of this study is to address the two sampling processes (temporary  
114 emigration and removal process) separately and examine whether sampled abundance  
115 can be compared among different habitats. First, to examine the effect of the temporary  
116 migration process on the sampled abundance, we installed enclosure fences (hereafter,  
117 fences) preventing the emigration of ground-dwelling beetles into/out of the trapping  
118 area and compared their sampled abundance among areas with and without fences and  
119 different habitats (clear-cut and forested areas). We hypothesized that the fences prevent  
120 individuals from moving into the effective trap areas, and thus, the sampled abundance  
121 within fences decreases more than that without fences. We also examined whether the  
122 fence effect differed between habitats. If the effect differs between habitats, the  
123 comparison of sampled abundance may be difficult because initial abundances are likely  
124 to differ between sites, even if the sampled abundances are the same. In fact, the  
125 movement activity of ground-dwelling beetles can change depending on its habitat (e.g.,  
126 Niemela & Halme 1992; Allema *et al.* 2015; Růžicková *et al.* 2021).

127           Second, to examine the effect of the removal sampling process on sampled  
128 abundance, we conducted removal sampling within fences and examined the  
129 relationships between daily removal probability and sampled abundance. By installing  
130 fences, we excluded the effects of the temporary migration process on the sampled  
131 abundance and focused on the removal process. In this study, we focused on the effect  
132 of soil temperature on removal probability because temperature is one of the  
133 fundamental factors determining insect activity (e.g., Messenger 1959). High  
134 temperature can increase movement activity and sampled abundance of ground-  
135 dwelling beetles (Honek 1988; Honek 1997; Raworth & Choi 2001; Saska *et al.* 2013),  
136 and temperature is highly variable on a short time scale and is easy to monitor. If soil  
137 temperature affects removal probability, it may be problematic because sampled  
138 abundances would differ between sites and/or periods with different soil temperatures  
139 even if initial abundances are the same. We hypothesized that a high daily soil  
140 temperature increases the daily removal probability and indirectly increases the sampled  
141 abundance on each sampling day.

142

## 143 **Materials and methods**

### 144 Study area and experimental design

145           The study area is in the Irumukeppu Mountain area of Hokkaido in northern  
146 Japan (43°34'37" to 39°26'N, 142°05'27" to 09°33'E). The mean annual temperature  
147 and precipitation in this area are 7.4°C and 1141.6 mm, respectively, according to the  
148 data obtained from 1991 to 2020 by the Ashibetsu Observatory, which is the closest to  
149 the study area (Japan Meteorological Agency 2021). This area is dominated by conifer  
150 plantations, which mostly consist of Sakhalin fir (*Abies sachalinensis* [F. Schmidt]

151 Mast.). Harvesting significantly changes forest structure, including forest floor  
152 vegetation and soil conditions, and these changes can affect the initial abundance of  
153 each species and may change their movement activity. Therefore, we surveyed two  
154 forest stands, a clear-cut plantation site and an unharvested plantation site (Figure 1a;  
155 Appendix 1), to examine the effects of the two sampling processes on sampled  
156 abundance. At the clear-cut site, a mature plantation of Sakhalin fir was last harvested in  
157 2015; hence, three to four years had passed by the time of the survey. The unharvested  
158 plantation site mainly consisted of mature Sakhalin fir, and its forest age is  
159 approximately 55 years (Akashi *et al.* 2017).

160 We installed three fences with galvanized iron panels on each site (clear-cut  
161 site: C1, C2, and C3; plantation site: P1, P2, and P3). The fences were established as a 5  
162 × 5 m square with a height of 0.8 m above ground level, and their footings were buried  
163 0.1 m deep under the surface (Figure 1b). We installed a grid of nine pitfall traps (3 × 3  
164 traps) within each fence. Each pitfall trap consisted of a plastic cup with a diameter of  
165 95 mm and a depth of 124 mm and contained 100% propylene glycol for preservation.  
166 The distances among traps and between traps and the nearest fence were 1 m and 1.5 m,  
167 respectively (Figure 1c). We also installed the same trap grid outside of the fences at  
168 least 5 m away from each fence. In all, we set 12 grids (2 sites × 3 fences × 2 grids;  
169 Figure 1a). Fences were installed in May 2018, and sampling was conducted four times  
170 during the following 14 months (period 1 [from June 28 or 29 to July 5 or 6, 2018],  
171 period 2 [from July 25 or 26 to August 1 or 2, 2018], period 3 [from June 21 to July 10,  
172 2019], and period 4 [from July 31 to August 21, 2019]: Appendix 5). In each period, we  
173 collected ground-dwelling beetles each day for seven days. In periods 3 and 4, we  
174 conducted additional sampling after day seven. We collected samples on days 10, 12,

175 and 19 in period 3 and on days 14 and 21 in period 4 to examine the number of  
176 individuals who could be collected beyond the formal 7-day sampling period. Before  
177 each sampling period, we removed two of the panels, each 1.5 m in length, from two  
178 sides of each fence and allowed ground-dwelling beetles to move in and out of the  
179 fences to eliminate the effects of the fences on the initial abundance in each grid. In  
180 2018 (periods 1 and 2) and 2019 (periods 3 and 4), there were 18 and 20 days between  
181 the two sampling periods, respectively (Appendix 5).

182 To examine the soil temperature across sites and periods, we installed a  
183 temperature data logger (Onset, HOBO UA-001-64) 1 cm underground at the center of  
184 each fenced area. Then, to examine whether soil temperature is correlated with air  
185 temperature, another type of temperature data logger (Onset, HOBO U23-004) with a  
186 solar radiation shield (Onset, HOBO, RS3-B) was set to observe air temperature. The  
187 logger was set in the corner of the fences 1 m above the ground surface to avoid the  
188 lower airflow caused by fences. The loggers were set at fences C1, C2, P2, and P3 in  
189 periods 1 and 2; C2, C3, P1, P2, and P3 in period 3; and C2, C3, P1, and P2 in period 4.

190 We identified the species in the collected samples based on the work of Ueno *et*  
191 *al.* (1985) and counted the number of individuals of each species collected on each  
192 sampling day and at each grid. In this study, we categorized each species as a  
193 macropterous, apterous, or brachypterous species by checking their wing types  
194 according to Shibuya *et al.* (2018) and used only apterous and brachypterous species  
195 (species having no, very short, or rudimentary wings) in the analysis. This is because  
196 our fences had no roof, and we could not exclude the possibility of movement of  
197 individuals in species that can fly over fences.

198

199 Data Analysis

200 Effects of the temporary migration process on sampled abundance

201 To examine the effects of the temporary migration process (i.e., installing  
202 fences) on the sampled abundance of pitfall traps, we examined whether the sampled  
203 abundance of ground-dwelling beetle species differed between areas with and without  
204 fences and whether the effect of fences differed between the two habitats. Here, we  
205 constructed abundance-based hierarchical community models (HCMs) according to the  
206 method of Yamaura *et al.* (2016). HCMs can be used to estimate multiple hierarchical  
207 parameters, allowing for simultaneous estimation of both community-level responses  
208 (common responses of targeted species) and species-level responses.

209 In this analysis, we focused on ten ground-dwelling beetles (nine carabid  
210 species and one carrion species: see Results section). For the HCMs, we pooled the  
211 sampled abundance for each species from the whole period (7 sampling days  $\times$  4  
212 periods) and used the total sampled abundance of each species as response variables  
213 following a negative binomial distribution. As explanatory variables for the species-  
214 level response variables, we used habitat type (binary variable where 0 denotes clear-cut  
215 and 1 denotes plantation), installed fence (binary variable where 0 denotes absent and 1  
216 denotes present), and the interaction of the habitat type and installed fence variables. We  
217 also considered the log-transformed number of valid traps as an offset term because  
218 several traps without fences were disturbed by mammals. In addition, for the intercept  
219 and coefficients of each species, we set hyperparameters that were shared by species  
220 parameters to estimate species- and community-level responses (see details in Appendix  
221 2).

222 To build the HCMs, we used R version 4.0.5 (R Core Team 2021) and the R

223 package “RStan” version 2.21.2 (Stan Development Team 2020). For the model, we  
224 used four chains that had 20,000 iterations following 5,000 burn-in with 20 thinning  
225 rate. We used the “R-hat” and “N\_eff/N” statistics ( $R\text{-hat} < 1.1$  and  $N\text{-eff}/N > 0.1$ ) to  
226 assess the convergence of the model according to Gabry and Modrák (2019).

227

## 228 Effects of the removal process on sampled abundance

229 Removal sampling is a method in which individuals are sequentially removed  
230 from a population to estimate population size, assuming that if the population is closed  
231 and the removal probability remains at a certain level, the number of individuals  
232 removed during each sampling event will gradually decrease. While classic removal  
233 sampling models assume that the probability is constant over a given period and/or  
234 across different sampling sites (Rodriguez de Rivera & McCrea 2021), recently, several  
235 models have been presented that can take into account the changes in removal  
236 probability with site-dependent factors (Kéry & Royle 2015; Kellner *et al.* 2022).

237 To examine the effects of the removal process (i.e., the changes in the removal  
238 probability) on the sampled abundance, we examined whether the daily soil temperature  
239 increases the daily removal probability using “unmarked” removal sampling models  
240 (Fiske & Chandler 2011). In the analysis, we focused on the most dominant species,  
241 *Pterostichus thunbergi*. We used the daily number of individuals collected within the  
242 fences during the four survey periods consisting of seven continuous sampling days. We  
243 used two types of data: (1) the whole period data, which covered the four survey  
244 periods, and (2) each period data, which included data from each survey period because  
245 the effects of soil temperature may differ across survey periods, i.e., we built four  
246 separate models for the four periods.

247           In the `unmarked` models, we used the daily number of individuals collected  
248 as the response variable, assuming a negative binomial distribution, and categorical  
249 habitat type (clear-cut vs. unharvested plantation) and standardized daily mean soil  
250 temperature as explanatory variables for initial abundance and removal probability,  
251 respectively. That is, the models assumed that the initial abundance of each grid was  
252 determined by habitat type and that the daily removal probability (and the daily number  
253 of individuals collected) depended on daily soil temperature. For the construction of the  
254 models, we used the “`gmultmix`” function in the R package “`unmarked`” (Fiske &  
255 Chandler 2011). We compared (1) the effects of soil temperature on removal probability  
256 across survey periods, (2) the estimated removal probability across survey periods, and  
257 (3) the initial abundance across survey periods and between habitat types.

258           Several removal sampling models have been presented in addition to the  
259 `unmarked` models, including `ubms` models (Kellner *et al.* 2022). Therefore, we  
260 additionally constructed several models, including the classic models that assume the  
261 removal probability remains constant across survey sites and periods, with different  
262 model-fitting methods and/or different distributions that were followed by response  
263 variables (`unmarked` Poisson, `ubms`, and `Rstan` models; Appendix 3). We compared  
264 the results with those of the abovementioned `unmarked` models.

265           Moreover, we constructed generalized linear models (GLMs) that assumed that  
266 the response variables follow a negative binomial distribution. In the models, we used  
267 the total sampled abundance of *P. thunbergi* in each period and at each grid as response  
268 variables and habitat type as explanatory variables. Unlike the abovementioned  
269 `unmarked` models, the GLMs do not account for the removal process, and therefore,  
270 the abundance estimates would include the bias caused by the removal process. We

271 compared the effects of habitat type on the abundance estimates between GLMs and the  
272 unmarked models to examine whether potential bias can impede the comparison of  
273 abundance estimates between different habitats. If the effects of habitat type on the  
274 abundance estimates were consistent between the GLMs and unmarked models, we  
275 would interpret this as evidence that models that do not account for the removal process  
276 could be used to obtain differences in the initial abundance between habitat types. For  
277 the GLMs, the “`glm.nb`” function of the R package MASS was used.

278           To build the removal sampling models and GLMs, we used R version 4.0.5 (R  
279 Core Team 2021). In all analyses, we used the 95% confidence interval or credible  
280 interval (which is used in the Bayesian approach instead of a confidence interval) to  
281 judge whether each variable was significant (hereafter, 95% CI). When the 95% CI of  
282 the coefficient of each variable or estimated value did not contain 0, we considered the  
283 variables to be significant.

284

285 **Results**

286 We collected a total of 606 individuals, consisting of ten apterous and  
287 brachypterous ground-dwelling beetle species (nine forest species and one nonforest  
288 species; Table 2), in four survey periods composed of seven continuous sampling days  
289 (not including extended surveys in periods 3 and 4). We collected 203 individuals of the  
290 most dominant species, *Pterostichus thunbergi*, within fences in four survey periods  
291 composed of seven continuous sampling days. We collected 111, 25, 33, and 34  
292 individuals of *P. thunbergi* in period 1, period 2, period 3, and period 4, respectively  
293 (Appendix 4). Extension sampling in periods 3 and 4 collected an additional 5 and 42  
294 individuals of *P. thunbergi*, respectively (Appendix 4). The daily mean values of soil  
295 temperature were relatively higher in the clear-cut site than in the unharvested  
296 plantation site, and the daily air temperature of each grid exhibited a trend similar to that  
297 of soil temperature (Appendix 5).

298

299 Effects of the temporary migration process on sampled abundance

300 We constructed HCMs using the sampled abundance of ten species (Table 2).  
301 For community-level hyperparameters, the coefficient of fence was significantly  
302 negative, and the 95% CI of the coefficients of habitat type and the interaction between  
303 habitat type and fences included zero (Figure 2). This suggested that, as the average  
304 response, the sampled abundance did not differ between habitat types but decreased  
305 within fences and that the fence effect did not differ between habitat types. At the  
306 species level, the coefficients of habitat type and fences varied among species. The  
307 sampled abundance of four species significantly differed between habitat types.  
308 Specifically, the abundance of *Carabus blaptoides rugipennis* (SP2) and *Xestagonum*

309 *daisetsuzanus* (SP10) was higher in the unharvested plantation site, while the abundance  
310 of *Carabus arcensis hokkaidoensis* (SP1) and *Pterostichus microcephalus* (SP6) was  
311 higher in the clear-cut site (Figure 2). The coefficients of fences were significantly  
312 negative for four species (*C. arcensis hokkaidoensis* [SP1], *Cychrus morawitzi morawitzi*  
313 [SP5], *Pterostichus orientalis antiquus* [SP7], and *Silpha perforata* [SP9]; Figure 2).  
314 Similar to the hyperparameter, the interaction term for habitat type and fences was not  
315 significant for any species (Figure 2). In addition, there was no apparent relationship  
316 between the coefficients of fences and the body size of each species (Appendix 6).

317

318 Effects of the removal process on sampled abundance

319 The estimated coefficients of removal sampling models were similar among the  
320 different model-fitting methods and the model with a different distribution (Appendix  
321 7). In addition, in periods 3 and 4, the mean values and ranges of most coefficients were  
322 overly high and wide, respectively (Figure 3c and Appendix 7), indicating that the  
323 models in these periods could not estimate the removal probability and initial  
324 abundance precisely.

325 In unmarked models, positive effects of soil temperature on removal  
326 probability were found only in period 1; that is, such positive effects were not observed  
327 in other periods (Figure 3a). In addition, the estimated daily removal probability  
328 differed among the survey periods (Figure 3b). During the whole period and during  
329 periods 1 and 2, the daily removal probability was relatively high, and the range of  
330 mean values was from 10 to 50% (Figure 3b). Conversely, the estimated probability was  
331 overly low for periods 3 and 4, and the mean values were close to zero (Figure 3b). The  
332 95% CI of the estimated initial abundance of *P. thunbergi* at clear-cut and plantation

333 sites overlapped in all periods (Figure 3c), suggesting that the initial abundance did not  
334 differ between habitat types.

335           Comparing the effects of habitat types on the abundance estimate between  
336 unmarked models and GLMs, in most cases, the effects were constant, that is, the  
337 abundance estimates of *P. thunbergi* were slightly higher at the plantation site than at the  
338 clear-cut site (Figure 3c). Thus, we concluded that the GLMs that do not account for the  
339 removal process could also show the differences in the initial abundance between  
340 habitat types. In addition, for the classic removal sampling models that do not account  
341 for the daily changes in removal probability, the estimated coefficients did not largely  
342 differ from those of unmarked models with the whole period data (Appendix 7). This  
343 also suggests that the effect of the removal process (effect of soil temperature on  
344 removal probability) was not substantial, especially when considered over the entire  
345 survey period.

346

347 **Discussion**

348 Effects of the temporary migration process on sampled abundance

349           We found that the sampled abundance of ground-dwelling beetles in pitfall  
350 traps decreased within fences. This result is consistent with those of previous studies  
351 demonstrating that the movement of carabids into and out of effective trap areas (i.e.,  
352 the temporary migration process) can increase their sampled abundance (Desender &  
353 Maelfait 1986; Andersen 1995; Holland & Smith 1999). Furthermore, we did not find  
354 that the effects were dependent on habitat type. Therefore, it could be argued that the  
355 temporary migration process did not impede the comparison of the sampled abundance  
356 among different habitats.

357           In addition, the results presented above also suggest that the movement activity  
358 of each species did not differ between clear-cut and unharvested forests, although  
359 previous studies demonstrated that the movement activity of forest carabids was  
360 generally larger in open land than in woodland (Allema *et al.* 2015; Růžicková *et al.*  
361 2021). At our clear-cut site, there was a high degree of logging residue (slash)  
362 (Appendix 1). Slash increases the microhabitat complexity of harvested areas, which  
363 potentially provides a hiding site for forest ground-dwelling beetles and thus could  
364 maintain the community composition (e.g., Gunnarsson *et al.* 2004; Nittérus *et al.*  
365 2007). Therefore, harvested areas with plenty of slash could be a suitable habitat for  
366 forest ground-dwelling beetles compared to other open lands, such as arable lands, and  
367 this may result in the small difference in the movement activity between the two  
368 habitats.

369           Previous studies suggested or assumed that large species have higher  
370 movement activity and are likely to be captured more than small species (Lang 2000;

371 Engel *et al.* 2017). However, our results also revealed that installing fences decreased  
372 the sampled abundance of each species, but the degree of the decrease was not  
373 correlated with the body size of the species (Appendix 5). Therefore, the negative fence  
374 effects were not explained by differences in body size. This was inconsistent with the  
375 findings of previous studies that pitfall trap bias can be corrected by using the body size  
376 of each species (Hancock & Legg 2012; Engel *et al.* 2017). This may be because our  
377 targeted species were relatively large (mean size: 9.75-34.5 mm), and most of the  
378 species had high movement activity rates.

379

380 Effects of daily soil temperature on the removal probability

381         Saska *et al.* (2013) conducted a meta-analysis on the effect of weather on the  
382 sampled abundance of carabids in several regions and demonstrated that while  
383 temperature generally increased the sampled abundance, the effects varied among  
384 species. We found that soil temperature increased the removal probability of pitfall traps  
385 and indirectly increased the sampled abundance of *Pterostichus thunbergi* but that  
386 positive effects were found in only one survey period (period 1: June 2018). With  
387 respect to *P. thunbergi*, therefore, it could be argued that the effects of soil temperature  
388 on the removal probability were not consistent throughout the survey periods. In period  
389 1, the daily mean values of soil temperature and their ranges were not substantially  
390 different from those of the other periods (Appendix 5), but the precipitation was higher  
391 than that in the other periods (Appendix 5). Although other weather factors, such as  
392 precipitation and soil humidity, can affect the removal probability (Mitchell 1963; Saska  
393 *et al.* 2013), the effect of soil temperature appeared to be small at least in our study,  
394 which focused on a restricted area and season (i.e., hot summer season). In fact, Saska *et*

395 *al.* (2013) suggested that temperature bias may be small on a short time scale and/or in  
396 the absence of distinct temporal trends in temperature.

397 In addition, because we did not examine differences in vegetation structure  
398 among sites, which can affect soil temperature, we could not clearly distinguish between  
399 the effects of soil temperature and vegetation structures on the removal probability.  
400 However, we dealt with daily changes in soil temperature, which can change greatly on  
401 a short time scale compared to vegetation structure; therefore, we concluded that our  
402 analysis mainly examined the effects of soil temperature changes on sampled abundance  
403 rather than those of vegetation.

404 We could not estimate the removal probability and initial abundance precisely  
405 for periods 3 and 4 (June and July 2019; Figure 3bc). In these periods, the total number  
406 of collected individuals of *P. thunbergi* was relatively small (Appendix 4). Moreover,  
407 the number of collected individuals was highest on the third day of periods 3 and 4  
408 (Appendix 4), despite the general expectation that the daily number of collected  
409 individuals would be high on the first day and decrease on subsequent days. It is  
410 possible that the small total number and irregular patterns of collected individuals in  
411 these periods made estimating the coefficients imprecise. Comparing the results of  
412 several removal sampling models, we did not find substantial differences in the  
413 coefficients among the models. Developing customized complex models, such as the  
414 RStan models (Appendix 3), required additional skill, knowledge, and effort to write  
415 code. Thus, `unmarked` and `ubms` models are recommended for estimating the initial  
416 abundance and removal probability, at least in our case.

417

418 Comparing abundance estimates among different habitats

419 Comparing the abundance estimates between the GLMs and unmarked  
420 models, we found that the habitat type effects on the abundance of *P. thunbergi* were  
421 consistent (Figure 3c), suggesting that the effects of the removal process on sampled  
422 abundance were small. In addition, we found that the temporary migration process did  
423 not depend on habitat differences (Figure 2). Therefore, we could compare the sampled  
424 abundance without considering the removal process to represent the difference in the  
425 initial abundance of *P. thunbergi*.

426 Furthermore, our results infer abundance changes in ground-dwelling beetles  
427 after harvest. Studies reported that the sampled abundance of some forest carabids did  
428 not decrease for several years after harvest (e.g., Work *et al.* 2010; Yamanaka *et al.*  
429 2021). There are two possible hypotheses for this finding. One, the changes in the  
430 movement activity of carabids in harvested areas masked the decreased initial  
431 abundance, and two, the initial abundance did not decrease immediately after the  
432 harvest. Our results support the latter hypothesis, suggesting that the initial abundance  
433 of these forest carabids did not change after harvest and that the beetles could survive in  
434 clear-cut areas for at least a few years.

435

## 436 Conclusion

437 We focused on two processes potentially affecting pitfall trap sampling and  
438 found that the sampled abundance of each species tended to increase in the temporary  
439 migration process, but its effect did not differ among habitat types. We also found that  
440 in the removal process, soil temperature could affect the removal probability and  
441 sampled abundance, but its effect was inconsistent and relatively small over the entire  
442 survey period. Therefore, we concluded that the traditional use of pitfall traps is, to

443 some extent, a robust measure, particularly when comparing the sampled abundance of  
444 each species among neighboring habitats, such as in the case of our study.

445 Our study also has several limitations. First, in our removal sampling models,  
446 we only examined one species' response and the effect of soil temperature. Removal  
447 probability and its influencing factors can vary among species (Saska *et al.* 2013;  
448 Růžičková *et al.* 2021) and may depend on individual and habitat conditions, such as  
449 starvation and prey density distribution (Raworth & Choi 2001). Second, we could not  
450 account for the possibility that vegetation density could hamper the movement of  
451 ground-dwelling beetles and decrease the removal probability (Melbourne 1999;  
452 Thomas *et al.* 2006). Third, we could not examine the reason that the removal  
453 probability and effect of soil temperature changed across the survey periods. Future  
454 studies should focus on multispecies responses, the effects of structural complexity, and  
455 different spatial-temporal ranges to compensate for our study's limitations and to gain a  
456 full understanding of the limitations and effective usage of pitfall traps.

457 **Author contributions**

458 Conceptualization: SY, KO; Methodology: SY; Investigation: SY; Formal analysis: SY,  
459 YY; Writing - original draft: SY, YY; Writing - review & editing: SY, YY, KO; Funding  
460 acquisition: SY, KO.

461

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467 18H04154, and 20K15561.

468

469 **Conflict of Interest**

470 The authors declare that they have no known competing financial interests or personal  
471 relationships that could have appeared to influence the work reported in this paper.

472

473 **Figure captions**

474

475 Figure 1 Study sites and enclosure experiment

476 (a) Location of study sites and sampling grids. C1-3 and P1-3 indicate sampling  
477 grids in clear-cut and unharvested plantation sites, respectively. (b) Enclosure fences  
478 installed. (c) Spatial arrangement of a sampling grid.

479

480 Figure 2 Coefficients of hyperparameters and species parameters of the HCMs

481 Black circles and horizontal bars indicate mean values and 95% CIs of  
482 coefficients, respectively. Dashed lines indicate zero values. Intercept: coefficient  
483 indicating log-scale sampled abundance at clear-cut site; Habitat type: changes in  
484 sampled abundance at unharvested plantation compared to clear-cut site; Fence: the  
485 changes in sampled abundance due to installing fences; Habitat type  $\times$  Fence: the  
486 interaction between habitat type and fences. When the 95% CIs of the coefficients of  
487 Habitat types, Fence, and Habitat  $\times$  Fence were above (below) zero, we considered that  
488 the abundance significantly increased (decreased) in the plantation site, within fences,  
489 and within fences in the plantation site, respectively. Hyperparameter: the community-  
490 level response, SP01 CAAR: *Carabus arcensis hokkaidoensis*, SP02 CABL: *Carabus*  
491 *blaptoides rugipennis*, SP03 CAGE: *Carabus gehinii gehinii*, SP04 CAOP: *Carabus*  
492 *opaculus opaculus*, SP05 CYMO: *Cychrus morawitzi morawitzi*, SP06 PTMI:  
493 *Pterostichus microcephalus*, SP07 PTOR: *Pterostichus orientalis antiquus*, SP08  
494 PPTH: *Pterostichus thunbergi*, SP09 SIPE: *Silpha perforata*, SP10 XEDA: *Xestagonum*  
495 *daisetsuzanus*.

496

497 Figure 3 Results of removal sampling models (unmarked model) with the whole  
498 period data (Whole period) and each period data (Periods 1, 2, 3, and 4) in the analysis  
499 of *P. thunbergi*

500 The mean values of estimates are denoted by circles or triangles, and their 95%  
501 CIs are denoted by vertical bars. (a) Effects of daily soil temperature on removal  
502 probability, (b) estimated per-day removal probability (given the mean soil  
503 temperature), and (c) abundance estimates are shown separately for different survey  
504 periods. In (c), we compared the abundance estimates between unmarked models  
505 (unmarked) and generalized linear models that do not consider the removal process  
506 (GLM).

507

#### 508 **Table captions**

509 Table 1 Glossary of terms used in this study

510

511 Table 2 Mean value of the sampled abundance of each species per day at clear-cut and  
512 unharvested plantation sites

513 Values indicate the mean value of sampled abundance per day and per 1 pitfall  
514 trap. Values with parentheses indicate the standard deviations. These values show the  
515 sampled abundance of 7 continuous days in four survey periods (not including extended  
516 sampling in periods 3 and 4). Mean values were calculated with the data of whole  
517 periods (each 7-sampling day in June and July 2018 and 2019). The scientific name of  
518 each species was in accordance with Suzuki (2022). Carabid species are categorized  
519 into forest and nonforest species according to the literature (Hori 2001; Hori 2003; Hori  
520 2012; Yamanaka *et al.* 2021). Silphid species (*Silpha perforata*) are categorized as

521 forest species according to Katakura *et al.* (1986). *Pterostichus microcephalus* is a wing  
522 polymorphic species (Shibuya *et al.* 2018), but most of the individuals we captured  
523 were brachypterous (only 2 of 18 individuals were macropterous). Thus, *P.*  
524 *microcephalus* was included in the analysis.  
525

526 **References**

- 527 Adis, J. (1979) Problems of interpreting arthropod-sampling with pitfall traps.  
528 *Zoologischer Anzeiger Jena*, **202**, 177-184.
- 529 Akashi, N., Tsushima, T., Unno, A., Nagasaka, A., Nagasaka, Y., Ohno, Y., Nitta, N.,  
530 Watanabe, I., Minamino, K., Yamada, K., Ishihama, N., Takiya, M., Tsuda, T.,  
531 Takeuchi, F., Ishizuka, W., Fukuchi, M., Yamaura, Y., Ozaki, K., Hironaka, Y. &  
532 Inari, N. (2017) Composition of trees before harvesting in the Retention  
533 Experiment for plantation FoREstry in Sorachi, Hokkaido (REFRESH) sites (in  
534 Japanese with English abstract). *Bulletin of the Hokkaido Forestry Research*  
535 *Institute*, **54**, 31-45.
- 536 Allema, A.B., van der Werf, W., Groot, J.C., Hemerik, L., Gort, G., Rossing, W.A. &  
537 van Lenteren, J.C. (2015) Quantification of motility of carabid beetles in  
538 farmland. *Bulletin of Entomological Research*, **105**, 234-244.
- 539 Andersen, J. (1995) A comparison of pitfall trapping and quadrat sampling of Carabidae  
540 (Coleoptera) on river banks. *Entomologica Fennica*, **6**, 65-77.
- 541 Baars, M.A. (1979) Catches in pitfall traps in relation to mean densities of carabid  
542 beetles. *Oecologia*, **41**, 25-46.
- 543 Brown, G.R. & Matthews, I.M. (2016) A review of extensive variation in the design of  
544 pitfall traps and a proposal for a standard pitfall trap design for monitoring  
545 ground-active arthropod biodiversity. *Ecology and Evolution*, **6**, 3953-3964.
- 546 Desender, K. & Maelfait, J.-P. (1986) Pitfall trapping within enclosures: a method for  
547 estimating the relationship between the abundances of coexisting carabid species  
548 (Coleoptera: Carabidae). *Ecography*, **9**, 245-250.
- 549 Engel, J., Hertzog, L., Tiede, J., Wagg, C., Ebeling, A., Briesen, H. & Weisser, W.W.

550 (2017) Pitfall trap sampling bias depends on body mass, temperature, and trap  
551 number: insights from an individual-based model. *Ecosphere*, **8**, e01790.

552 Fiske, I.J. & Chandler, R.B. (2011) Unmarked: An R Package for Fitting Hierarchical  
553 Models of Wildlife Occurrence and Abundance. *Journal of Statistical Software*,  
554 **43**, 1-23.

555 Gabry, J. & Modrák, M. (2019) Visual MCMC diagnostics using the bayesplot package,  
556 [https://cran.r-project.org/web/packages/bayesplot/vignettes/visual-mcmc-  
558 diagnostics.html](https://cran.r-project.org/web/packages/bayesplot/vignettes/visual-mcmc-<br/>557 diagnostics.html).

558 Gunnarsson, B., Nittérus, K. & Wirdenäs, P. (2004) Effects of logging residue removal  
559 on ground-active beetles in temperate forests. *Forest Ecology and Management*,  
560 **201**, 229-239.

561 Hancock, M.H. & Legg, C.J. (2012) Pitfall trapping bias and arthropod body mass.  
562 *Insect Conservation and Diversity*, **5**, 312-318.

563 Henderson, P. & Southwood, T. (2016) *Ecological Methods 4th Edition*.

564 Holland, J.M. & Smith, S. (1999) Sampling epigeal arthropods: an evaluation of fenced  
565 pitfall traps using mark-release-recapture and comparisons to unfenced pitfall  
566 traps in arable crops. *Entomologia Experimentalis et Applicata*, **91**, 347-357.

567 Honek, A. (1988) The Effect of Crop Density and Microclimate on Pitfall Trap Catches  
568 of Carabidae, Staphylinidae (Coleoptera), and Lycosidae (Araneae) in Cereal  
569 Fields. *Pedobiologia*, **32**, 233-242.

570 Honek, A. (1997) The effect of temperature on the activity of Carabidae (Coleoptera) in  
571 a fallow field. *European Journal of Entomology*, **94**, 97-104.

572 Hori, S. (2001) Edge effect of a forest viewed from ground beetle community  
573 (Coleoptera, Carabidae) (in Japanese with English abstract) *Bulletin of the*

574 *Historical Museum of Hokkaido*, **29**, 51-58.

575 Hori, S. (2003) Characteristics of carabid beetles inhabiting in the isolated forest (in  
576 Japanese with English abstract). *Bulletin of the Historical Museum of Hokkaido*,  
577 **31**, 15–28.

578 Hori, S. (2012) Forest environment monitoring using communities of ground beetles (in  
579 Japanese with English abstract). *Bulletin of the Historical Museum of Hokkaido*,  
580 **40**.

581 Iknayan, K.J., Tingley, M.W., Furnas, B.J. & Beissinger, S.R. (2014) Detecting  
582 diversity: emerging methods to estimate species diversity. *Trends in Ecology &*  
583 *Evolution*, **29**, 97-106.

584 Japan Meteorological Agency (2021) Historical weather data,  
585 <http://www.data.jma.go.jp/obd/stats/etrn/index.php>.

586 Jouveau, S., Toïgo, M., Giffard, B., Castagneyrol, B., Halder, I., Vétillard, F. & Jactel,  
587 H. (2019) Carabid activity-density increases with forest vegetation diversity at  
588 different spatial scales. *Insect Conservation and Diversity*, **13**, 36-46.

589 Katakura, H., Sonoda, M. & Yoshida, N. (1986) Carrion beetle (Coleoptera, Silphidae)  
590 fauna of Hokkaido University Tomakomai experiment forest, Northern Japan,  
591 with a note on the habitat preference of a geotrupine species, *Geotrupes*  
592 *laevistriatus* (Coleoptera, Scarabaeidae). *Research Bulletin of the Hokkaido*  
593 *University Forests*, **43**, 43-55.

594 Kellner, K.F., Fowler, N.L., Petroelje, T.R., Kautz, T.M., Beyer, D.E. & Belant, J.L.  
595 (2022) ubms: An R package for fitting hierarchical occupancy and N-mixture  
596 abundance models in a Bayesian framework. *Methods in Ecology and Evolution*,  
597 **13**, 577-584.

- 598 Kéry, M. & Royle, J.A. (2015) *Applied Hierarchical Modeling in Ecology: Analysis of*  
599 *distribution, abundance and species richness in R and BUGS: Volume 1:*  
600 *Prelude and Static Models*. Academic Press, Boston.
- 601 Koivula, M., Kotze, D., Hiisivuori, L. & Rita, H. (2003) Pitfall trap efficiency: do trap  
602 size, collecting fluid and vegetation structure matter? *Entomologica Fennica*, **14**,  
603 1-14.
- 604 Koivula, M.J., Venn, S., Hakola, P. & Niemela, J. (2019) Responses of boreal ground  
605 beetles (Coleoptera, Carabidae) to different logging regimes ten years post  
606 harvest. *Forest Ecology and Management*, **436**, 27-38.
- 607 Lang, A. (2000) The pitfalls of pitfalls: a comparison of pitfall trap catches and absolute  
608 density estimates of epigeal invertebrate predators in arable land. *Anzeiger Fur*  
609 *Schadlingskunde-Journal of Pest Science*, **73**, 99-106.
- 610 Lövei, G. & Magura, T. (2011) Can carabidologists spot a pitfall? The non-equivalence  
611 of two components of sampling effort in pitfall-trapped ground beetles  
612 (Carabidae). *Community Ecology*, **12**, 18-22.
- 613 Luff, M.L. (1975) Some features influencing the efficiency of pitfall traps. *Oecologia*,  
614 **19**, 345-357.
- 615 Melbourne, B.A. (1999) Bias in the effect of habitat structure on pitfall traps: An  
616 experimental evaluation. *Australian Journal of Ecology*, **24**, 228-239.
- 617 Messenger, P.S. (1959) Bioclimatic Studies with Insects. *Annual Review of Entomology*,  
618 **4**, 183-206.
- 619 Mitchell, B. (1963) Ecology of Two Carabid Beetles, *Bembidion lampros* (Herbst) and  
620 *Trechus quadristriatus* (Schrank). *The Journal of Animal Ecology*, **32**, 377-392.
- 621 Niemela, J. & Halme, E. (1992) Habitat associations of carabid beetles in fields and

622 forests on the Aland Islands, SW Finland. *Ecography*, **15**, 3-11.

623 Nittérus, K., Åström, M. & Gunnarsson, B. (2007) Commercial harvest of logging  
624 residue in clear-cuts affects the diversity and community composition of ground  
625 beetles (Coleoptera: Carabidae). *Scandinavian Journal of Forest Research*, **22**,  
626 231-240.

627 R Core Team (2021) R: A Language and Environment for Statistical Computing. R  
628 Foundation for Statistical Computing, Vienna, Austria.

629 Raworth, D.A. & Choi, M.Y. (2001) Determining numbers of active carabid beetles per  
630 unit area from pitfall-trap data. *Entomologia Experimentalis et Applicata*, **98**,  
631 95-108.

632 Rodriguez de Rivera, O. & McCrea, R. (2021) Removal modelling in ecology: A  
633 systematic review. *Plos One*, **16**, e0229965.

634 Růžicková, J., Bérces, S., Ackov, S. & Elek, Z. (2021) Individual movement of large  
635 carabids as a link for activity density patterns in various forestry treatments. *Acta*  
636 *Zoologica Academiae Scientiarum Hungaricae*, **67**, 77-86.

637 Saska, P., van der Werf, W., Hemerik, L., Luff, M.L., Hatten, T.D., Honek, A. & Pocock,  
638 M. (2013) Temperature effects on pitfall catches of epigeal arthropods: a model  
639 and method for bias correction. *Journal of Applied Ecology*, **50**, 181-189.

640 Schirmel, J., Lenze, S., Katzmann, D. & Buchholz, S. (2010) Capture efficiency of  
641 pitfall traps is highly affected by sampling interval. *Entomologia Experimentalis*  
642 *et Applicata*, **136**, 206-210.

643 Shibuya, S., Kiritani, K. & Fukuda, K. (2018) Hind wings in ground beetles  
644 (Coleoptera: Carabidae and Brachinidae) – morphology, length, and  
645 characteristics of each subfamily (Japanese with English abstract). *Japanese*

646 *Journal of Ecology*, **68**, 19-41.

647 Stan Development Team (2020) "Rstan: the R interface to Stan." R package version  
648 2.21.2.

649 Suzuki, S. (2022) Catalogue of the Insects of the Japanese archipelago,  
650 <https://japanesebeetles.jimdofree.com/>.

651 Thomas, C.F.G., Brown, N.J. & Kendall, D.A. (2006) Carabid movement and  
652 vegetation density: Implications for interpreting pitfall trap data from split-field  
653 trials. *Agriculture, Ecosystems & Environment*, **113**, 51-61.

654 Ueno, S., Kurosawa, Y. & Sato, M. (1985) *The Coleoptera of Japan in Color, Vol. II (in*  
655 *Japanese)*. Hoikusha, Osaka.

656 Woodcock, B.A. (2005) Pitfall Trapping in Ecological Studies. *Insect sampling in forest*  
657 *ecosystems* (ed. S. Leather), pp. 37-57. Blackwell Science Ltd, Malden.

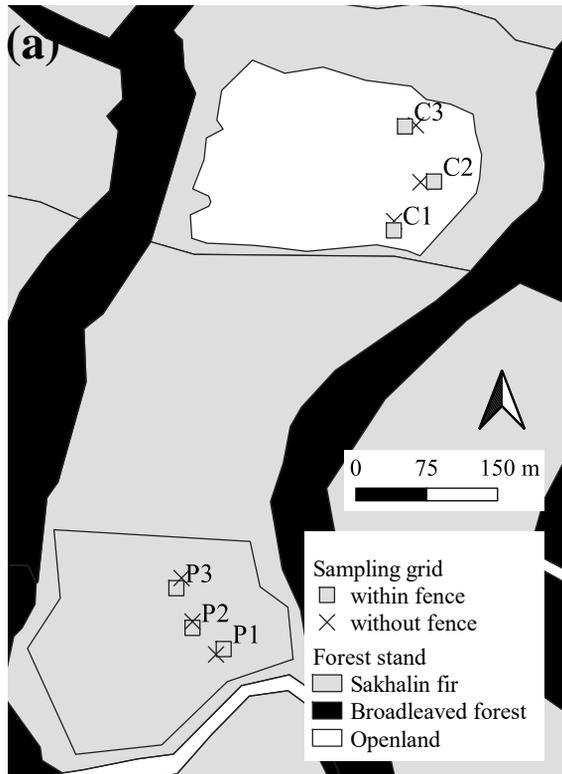
658 Work, T.T., Jacobs, J.M., Spence, J.R. & Volney, W.J. (2010) High levels of green-tree  
659 retention are required to preserve ground beetle biodiversity in boreal  
660 mixedwood forests. *Ecological Applications*, **20**, 741-751.

661 Yamanaka, S., Yamaura, Y., Sayama, K., Sato, S. & Ozaki, K. (2021) Effects of  
662 dispersed broadleaved and aggregated conifer tree retention on ground beetles in  
663 conifer plantations. *Forest Ecology and Management*, **489**, 119073.

664 Yamaura, Y., Kery, M. & Royle, J.A. (2016) Study of biological communities subject to  
665 imperfect detection: bias and precision of community N-mixture abundance  
666 models in small-sample situations. *Ecological Research*, **31**, 289-305.

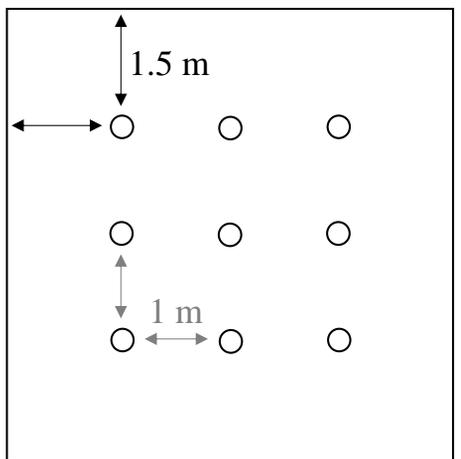
667 Yamaura, Y. & Royle, J.A. (2017) Community distance sampling models allowing for  
668 imperfect detection and temporary emigration. *Ecosphere*, **8**, e02028.

669



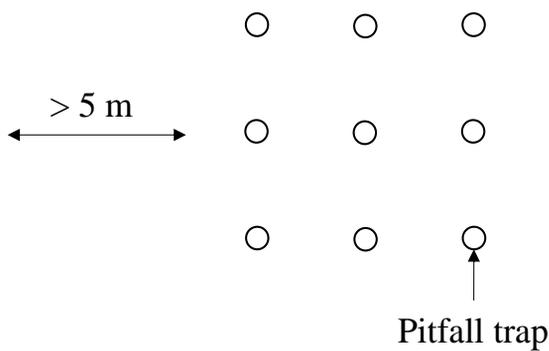
**(c)**

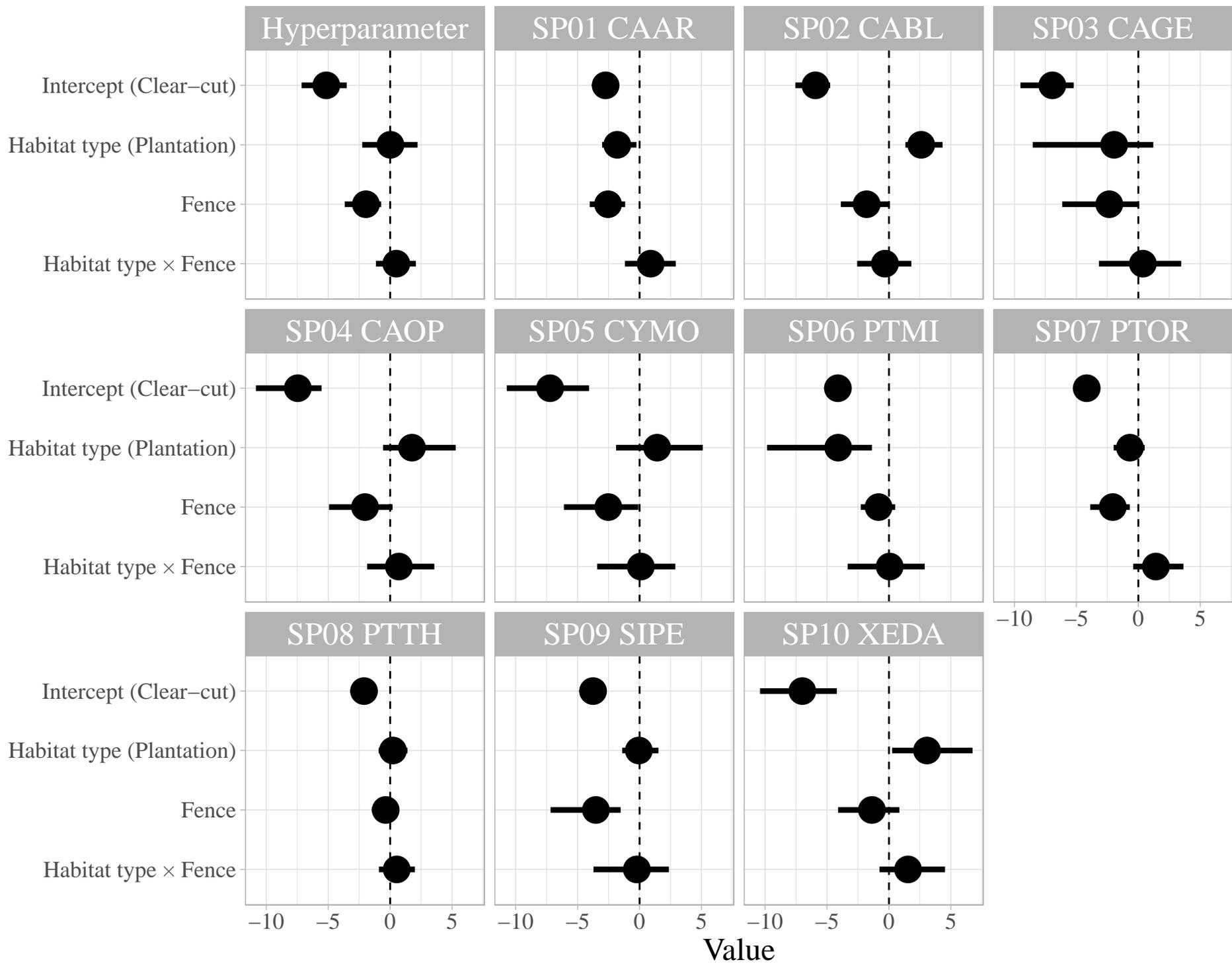
Sampling grid within fence



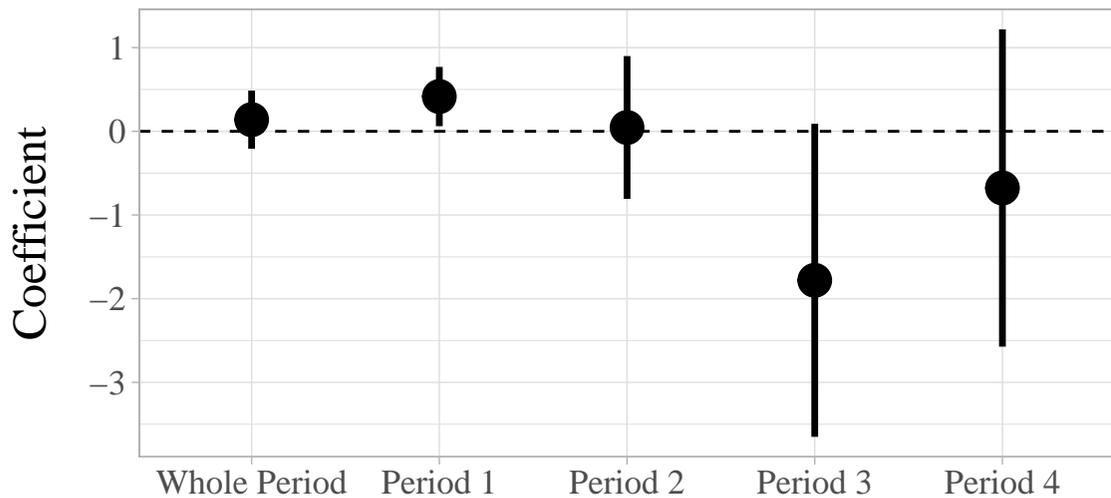
5 × 5 m fence

Sampling grid without fence

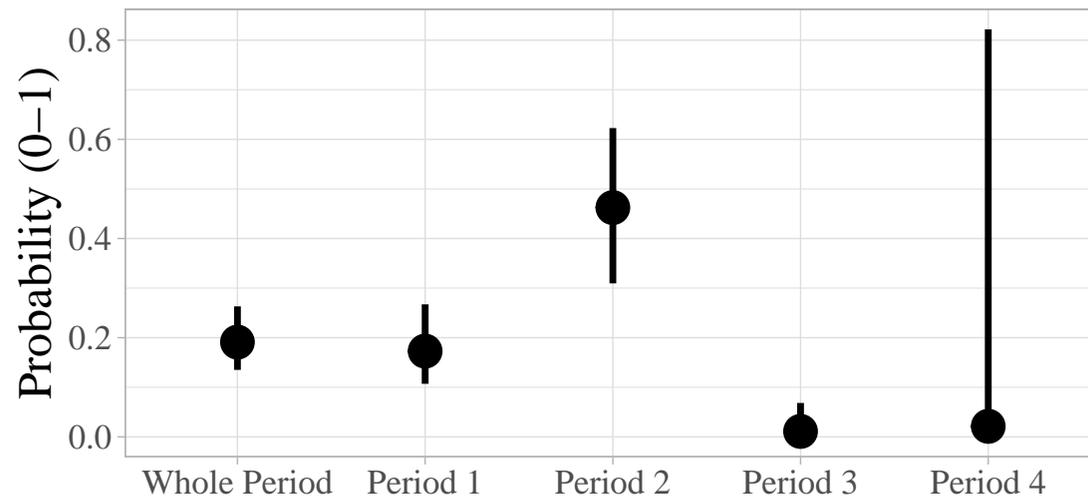




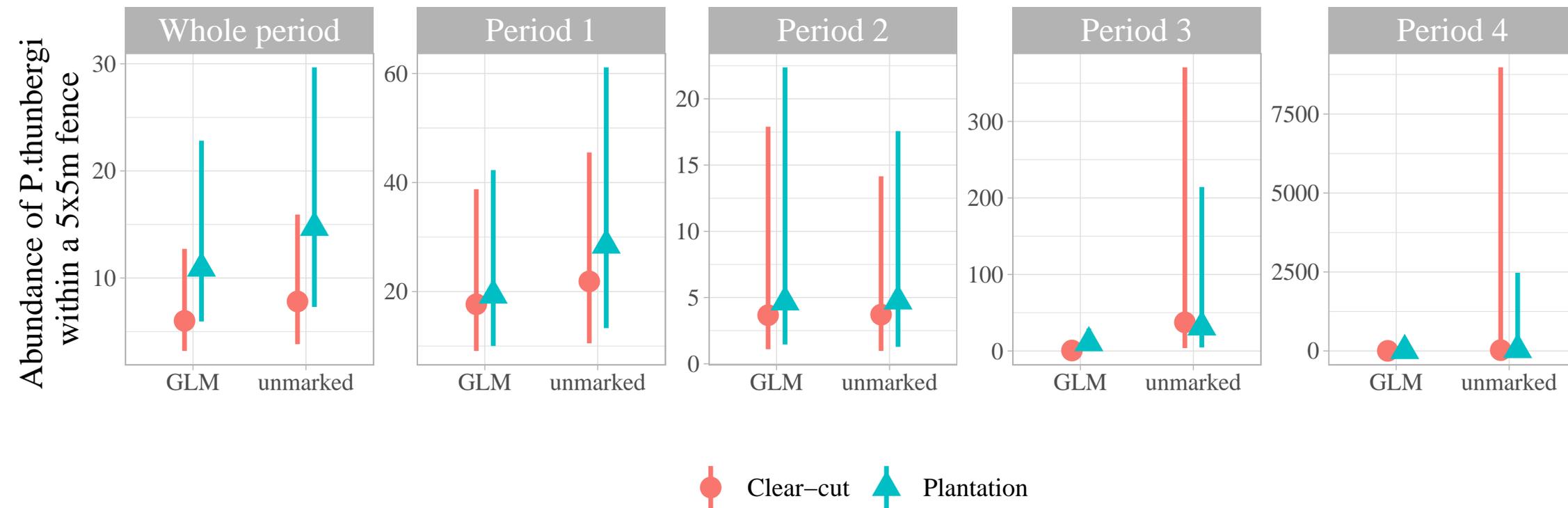
**(a) Effect of soil temperature on removal probability**



**(b) Estimated removal probability**



**(c) Abundance estimates**



1 **Table**

2 Table 1 Glossary of terms used in this study.

Terms	Explanation
Initial abundance	<ul style="list-style-type: none"><li>• The number of individuals present in a certain area during an arbitrary time.</li><li>• In this study, we regarded initial abundance as the number of individuals present within the fences when the fences were shut down, and we estimated it using removal sampling.</li></ul>
Sampled abundance	<ul style="list-style-type: none"><li>• The number of individuals actually collected by pitfall traps.</li><li>• It is determined through the sampling process of pitfall traps, which is influenced by various factors (sampling design and targeted site- and species-specific characteristics).</li></ul>
Removal probability	<ul style="list-style-type: none"><li>• The probability that individuals within effective trap area are caught and removed by pitfall traps during a certain period.</li><li>• It depends on a sampling design and the movement activity of each species.</li><li>• In this study, we regard the removal probability as the probability that an individual is collected by traps per day in an enclosure experiment in the absence of a temporary migration process.</li></ul>
Movement activity	<ul style="list-style-type: none"><li>• It represents the degree of activity of each species, which is dependent on intrinsic factors (e.g., their body size, species-specific movement distance or frequency) and extrinsic factors (e.g., soil temperature, vegetation structures).</li><li>• It affects both the temporary migration and removal processes.</li></ul>
Sampled abundance within fence	<ul style="list-style-type: none"><li>• The total number of individuals collected within fences.</li><li>• It is determined by the initial abundance and removal probability within fences, and thus it is expected to be lower than the initial abundance but close to the initial abundance with a longer survey period.</li></ul>

- 
- The movement of individuals over fences is prevented, and thus sampled abundance within fence is not affected by temporary migration process.

- 
- Sampled abundance without fence
- The total number of individuals collected without fences.
  - It is the common measure of comparison in pitfall trap survey.
-

4 Table 2 Mean value of the sampled abundance of each species per day at clear-cut and unharvested plantation sites

5

ID	Species name	Abbreviation	Habitat	Taxa	Clear-cut		Plantation	
					Within fence	Without fence	Within fence	Without fence
1	<i>Carabus arcensis</i> <i>hokkaidoensis</i>	CAAR	Forest	Carabid	0.004 (0.0072)	0.0757 (0.0625)	0.0026 (0.0092)	0.0093 (0.0143)
2	<i>Carabus blaptoides</i> <i>rugipennis</i>	CABL	Forest	Carabid	0.0013 (0.0046)	0.0013 (0.0046)	0.004 (0.0072)	0.0384 (0.0386)
3	<i>Carabus gehinii gehinii</i>	CAGE	Forest	Carabid	0 (0)	0.0013 (0.0047)	0 (0)	0 (0)
4	<i>Carabus opaculus opaculus</i>	CAOP	Forest	Carabid	0 (0)	0 (0)	0.0013 (0.0046)	0.004 (0.0072)
5	<i>Cychrus morawitzi morawitzi</i>	CYMO	Forest	Carabid	0 (0)	0 (0)	0 (0)	0.004 (0.0137)
6	<i>Pterostichus microcephalus</i>	PTMI	NonForest	Carabid	0.0093 (0.0197)	0.0146 (0.023)	0 (0)	0 (0)
7	<i>Pterostichus orientalis</i> <i>antiquus</i>	PTOR	Forest	Carabid	0.0013 (0.0046)	0.0186 (0.0309)	0.0053 (0.0078)	0.0066 (0.0106)
8	<i>Pterostichus thunbergi</i>	PTTH	Forest	Carabid	0.0952 (0.1463)	0.1067 (0.1492)	0.1733 (0.1401)	0.1455 (0.0593)
9	<i>Silpha perforata</i>	SIPE	Forest	Silphid	0 (0)	0.0274 (0.0448)	0 (0)	0.0238 (0.0441)

10	<i>Xestagonum daisetsuzanus</i>	XEDA	Forest	Carabid	0 (0)	0 (0)	0.0278 (0.0479)	0.0132 (0.0223)
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6 Values indicate the mean value of sampled abundance per day and per 1 pitfall trap. Values with parentheses indicate the standard  
7 deviations. These values show the sampled abundance of 7 continuous days in four survey periods (not including extended sampling in  
8 periods 3 and 4). Mean values were calculated with the data of whole periods (each 7-sampling day in June and July 2018 and 2019).  
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10 species according to the literature (Hori 2001; Hori 2003; Hori 2012; Yamanaka et al. 2021). Silphid species (*Silpha perforata*) are  
11 categorized as forest species according to Katakura et al. (1986). *Pterostichus microcephalus* is a wing polymorphic species (Shibuya et  
12 al. 2018), but most of the individuals we captured were brachypterous (only 2 of 18 individuals were macropterous). Thus, P.  
13 microcephalus was included in the analysis.

# Appendix

## Examining the bias of pitfall traps with enclosure experiments and removal sampling

Satoshi YAMANAKA, Yuichi YAMAURA, Kenichi OZAKI

Contents

### **Materials and methods**

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Appendix 4 Sampled abundance of *Pterostichus thunbergi* within fences

Appendix 5 Weather factors (soil and air temperature and precipitation)

Appendix 6 Relationship between body size and coefficients of the effects of fences on sampled abundance

Appendix 7 Results of removal sampling models

### **Data sources and scripts for analysis**

(stored in different files)

Data source

1. Data of sampled abundance of each species (“Data1.txt”)
2. Data of sampled abundance of *Pterostichus thunbergi* and soil temperature during

- 26           7 days (“Data2.txt”)
- 27        3. Data of body size of each species (“Splist.txt”)
- 28
- 29   Script for analysis
- 30        1. Script for HCMs (“Rcode1.R”)
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- 33        4. Stan code for HCMs (“Stancode1.stan”)
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- 35           (“Stancode2.stan”)
- 36        6. Stan code for removal sampling models with each period data
- 37           (“Stancode3.stan”)
- 38        7. Stan code for classic removal sampling models (“Stancode4.stan”)
- 39

40 **Appendix 1 Study site**

41 Our survey was conducted a clear-cut site and a plantation site (Figure S1). At  
42 the clear-cut site, a mature plantation of Sakhalin fir was harvested in 2015, and 3 or 4  
43 years had passed at the time of the survey. In clear-cut site, logging residue produced by  
44 the harvesting remained (Figure S1a). The unharvested plantation site mainly consists of  
45 mature Sakhalin fir.



47 Figure S1 Study sites

48 (a) clear-cut site. (b) plantation site.

49

## 50 **Appendix 2 Abundance-based hierarchical community models**

51 In order to examine the effects of temporary emigration process (i.e., fence  
52 installation) on sampled abundance of each species, we constructed the abundance-  
53 based hierarchical community models (HCMs) according to Yamaura *et al.* (2016).  
54 HCMs can estimate a community-level response based on species-level responses.

55 Here, we pooled the sampled abundance of each species at 4 survey periods.  
56 We assumed that the total sampled abundance of species  $i$  at grid  $j$  ( $Y_{ij}$ ) follows a  
57 negative binomial distribution:

$$58 \quad Y_{ij} \sim \text{NegativeBinomial}(\lambda_{ij}, \sigma_i), \quad (1)$$

59 where  $\lambda_{ij}$  is an expected sampled abundance of species  $i$  in grid  $j$ , and  $\sigma_i$  is a  
60 dispersion parameter of species  $i$ . We modeled  $\lambda_{ij}$  as a function of an intercept ( $\alpha_i$ ),  
61 habitat type (binary variable  $H_j$ ; 0: clear-cut, 1: plantation), installed fence (binary  
62 variable  $F_j$ ; 0: absent, 1: present), and the interaction of habitat type and installed fence  
63 variables. We also included the log-transformed number of the valid traps as an offset  
64 term because several traps without fences were disturbed by mammals:

$$65 \quad \log(\lambda_{ij}) = \alpha_i + \beta_{h_i} \times H_j + \beta_{f_i} \times F_j + \beta_{hf_i} \times H_j \times F_j + \log(T_j), \quad (2)$$

66 where  $\beta_{h_i}$ ,  $\beta_{f_i}$ , and  $\beta_{hf_i}$  are coefficients of habitat type, installed fence, and the  
67 interaction of habitat type and installed fence variables, respectively. These coefficients  
68 indicate the changes in sampled abundance in unharvested plantation site compared to  
69 clear-cut site, abundance changes due to fence installation, and abundance changes due  
70 to fence installation in plantation site, respectively.  $T_j$  is the number of valid traps in  
71 grid  $j$ , and  $\log(T_j)$  is the offset term.

72 In order to estimate species- and community-level responses, we assumed that the  
73 species level parameters ( $\alpha_i$ ,  $\beta_{h_i}$ ,  $\beta_{f_i}$ , and  $\beta_{hf_i}$ ) follow normal distributions shared by  
74 all species,

75  $\alpha_i \sim \text{Normal}[\mu_\alpha, \sigma_\alpha^2],$   
76  $\beta_{h_i} \sim \text{Normal}[\mu_{\beta_h}, \sigma_{\beta_h}^2], \dots \quad (3)$

77 where,  $\mu_\alpha$  is the mean value of  $\alpha_i$ , showing the community-level response (i.e.,  
78 hyperparameter), and  $\sigma_\alpha$  is its standard deviation.

79 We used normal distributions  $(0, 100^2)$  for  $\mu_\alpha$ ,  $\mu_{\beta_h}$ ,  $\mu_{\beta_f}$  and  $\mu_{\beta_{hf}}$ , and half-  
80 Cauchy distributions  $(0, 5)$  for the standard deviation of each coefficient ( $\sigma_\alpha$ ,  $\sigma_{\beta_h}$ ,  $\sigma_{\beta_f}$ ,  
81 and  $\sigma_{\beta_{hf}}$ ) and dispersion parameter ( $\sigma_i$ ) as weakly informative prior distributions  
82 (Gelman 2006). We used four chains that had 20,000 iterations following 5,000 burn-in  
83 with 20 thinning rate.

84

## 85 **References**

- 86 Gelman, A. (2006) Prior distributions for variance parameters in hierarchical models  
87 (Comment on an Article by Browne and Draper). *Bayesian Analysis*, **1**, 515-533.  
88 Yamaura, Y., Kery, M. & Royle, J.A. (2016) Study of biological communities subject to  
89 imperfect detection: bias and precision of community N-mixture abundance  
90 models in small-sample situations. *Ecological Research*, **31**, 289-305.

91

## 92 **Appendix 3 Explanation of removal sampling models**

93

### 94 1. Removal sampling models considering the changes in removal probability

95 We constructed and fitted removal sampling models using three methods: (1) the R  
96 package “unmarked” version 1.0.1 (Fiske & Chandler 2011) with a Poisson  
97 distribution (hereafter, unmarked Poisson model), (2) the R package “ubms” version  
98 1.0.2 (Kellner et al. 2022), and (3) the custom model code in the R package “RStan”  
99 (Stan Development Team 2020). As well as unmarked models in the main text, we  
100 used two types of data: (1) the whole period data, which covered the four survey  
101 periods, and (2) each period’s data.

102

#### 103 1-1. unmarked Poisson models

104 We used the daily number of individuals collected as the response variable,  
105 following Poisson distributions, and categorical habitat type (clear-cut vs. unharvested  
106 plantation) and standardized daily mean soil temperature as explanatory variables for  
107 initial abundance and removal probability, respectively. We used the “gmultmix”  
108 function of the R package “unmarked” using maximum likelihood estimation (Fiske  
109 & Chandler 2011).

110

#### 111 1-2. ubms models

112 We used the “stan\_multinomPois” function of the R package “ubms”  
113 using Bayesian estimation and included random effects in the models (Kellner et al.  
114 2022). The response variables and explanatory variables are the same as unmarked  
115 models; however, as random effects, we added grid ID and period ID to explanatory  
116 variables for the initial abundance in the model with the whole period data and grid ID

117 to an explanatory variable in the models with each period data. We assumed that the  
 118 response variables followed a Poisson distribution because “ubms” does not support a  
 119 negative binomial distribution for removal sampling models. For each model, we used  
 120 four chains that had 4,000 iterations with no burn-in and one thinning rate.

121

### 122 1-3. RStan model

123 We developed custom codes according to Kéry and Royle (2015). In the  
 124 models, the response variables followed a Poisson distribution, and we included grid  
 125 and/or period IDs as in the ubms models. It is noted that the RStan models had the  
 126 same structure as the ubms models. For the model using the data set of the whole  
 127 survey period, we modeled expected value of the initial abundance of *Pterostichus*  
 128 *thunbergi* at grid  $i$  and period  $j$  ( $\lambda_{ij}$ ) as a function of habitat types and random effects  
 129 of sampling grids and survey periods:

$$130 \quad \log(\lambda_{ij}) = \alpha_i + \varepsilon_{Grid_i} + \varepsilon_{Period_j}, \quad (1)$$

131 where,  $\alpha_i$  is a categorical variable showing the initial abundance in clear-cut or  
 132 plantation sites.  $\varepsilon_{Transect_i}$  and  $\varepsilon_{Period_j}$  are random effects to account for the variation  
 133 among sampling grids and survey periods and were assumed to follow a normal  
 134 distribution:

$$135 \quad \varepsilon_{Grid_i} \sim \text{Normal}(0, \sigma_{Grid}^2),$$

$$136 \quad \varepsilon_{Period_j} \sim \text{Normal}(0, \sigma_{Period}^2), \quad (2)$$

137 where,  $\sigma$  indicates the standard deviation.

138 We assumed that the removal probability on day  $t$  in grid  $i$  at period  $j$  is  
 139 determined by the standardized daily mean value of soil temperature on day  $t$  in grid  $i$  at  
 140 period  $j$  ( $S_{ijt}$ ):

$$141 \quad \text{logit}(p_{ijt}) = \alpha_p + \beta_p \times S_{ijt}, \quad (3)$$

142 where,  $p_{ijt}$  is the removal probability on day  $t$  in grid  $i$  at period  $j$ , and its value is  
 143 ranged from 0 to 1.  $\alpha_p$  and  $\beta_p$  indicate an intercept and a coefficient of soil  
 144 temperature, respectively. We calculated the conditional cell probabilities of  
 145 multinomial models on each sampling day ( $\pi_t$ ). The conditional cell probability on a  
 146 certain day depends on the probability that the individual was not removed before that  
 147 day and the removal probability on that day. For example, the conditional cell  
 148 probability on day 2 ( $\pi_{ij2}$ ) depends on the probability that the individuals were not  
 149 removed on day 1 ( $1 - p_{ij1}$ ) and the removal probability on day 2 ( $p_{ij2}$ ):

$$\begin{aligned}
 150 \quad & \pi_{ij1} = p_{ij1}, \\
 151 \quad & \pi_{ij2} = (1 - p_{ij1}) \times p_{ij2}, \dots, \\
 152 \quad & \pi_{ij7} = (1 - p_{ij1}) \times (1 - p_{ij2}) \times (1 - p_{ij3}) \times (1 - p_{ij4}) \times (1 - p_{ij5}) \times \\
 153 \quad & (1 - p_{ij6}) \times p_{ij7}. \quad (4)
 \end{aligned}$$

154 We summed the conditional cell probability of each sampling day to calculate the  
 155 total removal probability ( $\sum_{t=1}^7 \pi_{ijt}$ ), which indicates the net probability that the  
 156 individual in grid  $i$  at period  $j$  was removed in any of 7 sampling days. In addition,  $\pi_{ij0}$   
 157 is the probability that the individual in grid  $i$  at period  $j$  was not removed during  
 158 sampling days and is shown as  $1 - \sum_{t=1}^7 \pi_{ijt}$ .

159 We assumed that sampled abundance on day  $t$  in grid  $i$  at period  $j$  ( $Y_{ijt}$ ) is  
 160 determined by the total sampled abundance ( $n_{ij}$ ) and rate of collected individuals on  
 161 day  $t$  out of the total sampled abundance ( $\pi_{ijt}^c = \pi_{ijt}/(1 - \pi_{ij0})$ ). We also assume that  
 162 total sampled abundance ( $n_{ij}$ ) depends on the initial abundance ( $N_{ij}$ ) and total removal  
 163 probability ( $1 - \pi_{ij0}$ ), and that the initial abundance follows a Poisson distribution:

$$\begin{aligned}
 164 \quad & Y_{ijt}|n_{ij} \sim \text{Multinom}(n_{ij}, \pi_{ijt}^c), \\
 165 \quad & n_{ij} \sim \text{Binomial}(N_{ij}, 1 - \pi_{ij0}), \\
 166 \quad & N_{ij} \sim \text{Poisson}(\lambda_{ij}). \quad (5)
 \end{aligned}$$

167 As prior distributions, we used normal distributions  $(0, 100^2)$  for  $\alpha_i$ ,  $\alpha_p$ , and  
168  $\beta_p$ . We also used half-Cauchy distributions  $(0, 5)$  for any variance of coefficients as  
169 weakly informative prior distributions (Gelman 2006). For each model, we used four  
170 chains that had 50,000 iterations following 25,000 burn-in with 20 thinning rate. We  
171 used the “R-hat” and “N\_eff/N” statistics (R-hat < 1.1 and N\_eff/N > 0.1) to assess  
172 model convergence.

173 We also constructed models using the data set of each survey period. In the  
174 models, we assumed that the expected initial abundance of *P. thunbergi* at grid  $i$  ( $\lambda_i$ )  
175 was a function of habitat types and random grid effects:

$$176 \log(\lambda_i) = \alpha_i + \varepsilon_{\text{Transect}_i} \quad (1)$$

177

## 178 2. Classic removal sampling models

179 We also constructed classic removal sampling models with four model-fitting  
180 methods (unmarked with a Poisson and negative-binomial distribution, ubms, and  
181 RStan models). These models were simpler than the above-mentioned models because  
182 the classic models do not consider effects of daily soil temperature on removal  
183 probability (“Rcode3.R” and “Stancode4.stan” in Script for analysis), meaning that in  
184 these models the daily removal probability is assumed to be constant through all survey  
185 periods and among sites. We pooled the sampled abundance across survey periods, and  
186 the total of sampled abundance on each sampling day and at each grid was used as the  
187 response variables.

188

## 189 References

190 Gelman, A. (2006) Prior distributions for variance parameters in hierarchical models  
191 (Comment on an Article by Browne and Draper). *Bayesian Analysis*, **1**, 515-533.

192 Fiske, I.J. & Chandler, R.B. (2011) Unmarked: An R Package for Fitting Hierarchical  
193 Models of Wildlife Occurrence and Abundance. *Journal of Statistical*  
194 *Software*, 43, 1-23.

195 Kellner, K.F., Fowler, N.L., Petroelje, T.R., Kautz, T.M., Beyer, D.E. & Belant, J.L.  
196 (2022) ubms: An R package for fitting hierarchical occupancy and N-mixture  
197 abundance models in a Bayesian framework. *Methods in Ecology and*  
198 *Evolution*, 13, 577-584.

199 Kéry, M. & Royle, J.A. (2015) *Applied Hierarchical Modeling in Ecology: Analysis of*  
200 *distribution, abundance and species richness in R and BUGS: Volume 1:*  
201 *Prelude and Static Models*. Academic Press, Boston.

202 Stan Development Team (2020) "Rstan: the R interface to Stan." R package version  
203 2.21.2.  
204  
205

206 **Appendix 4 Sampled abundance of *Pterostichus thunbergi* within fences**

207 We collected totally 203 individuals of *P. thunbergi* within fences during each  
 208 7-sampling day within 4 periods (Figure S4-1). The sampled abundance was highest at  
 209 period 1, June 2018 (111 individuals), and the abundance at the other periods was  
 210 relatively small. In fact, 34, 33 and 25 individuals were caught at period 4, period 3, and  
 211 period 2, respectively (Figure S4-1). At period 1 and 2, the daily sampled abundance  
 212 was highest on the first day tended to decrease after the following days. On the other  
 213 hand, at period 3 and 4, the sampled abundance was highest on the third day and  
 214 maintained after the following days (Table S4 and Figure S4-1).

215 In period 3 and period 4, we conducted extended sampling after day 7 and  
 216 additionally collected a total of 5 and 42 individuals of *P. thunbergi*, respectively  
 217 (Figure S4-2), suggesting that in these periods (especially period 4), the 7-day sampling  
 218 did not remove all the individuals within fences. These results were consistent with the  
 219 results of the removal sampling models showing that the estimated removal probability  
 220 was overly low in periods 3 and 4 (see Figure 3b and S7-1b in Appendix 7).

221

222 Table S4 Mean value of daily sampled abundance of *Pterostichus thunbergi* at each  
 223 period

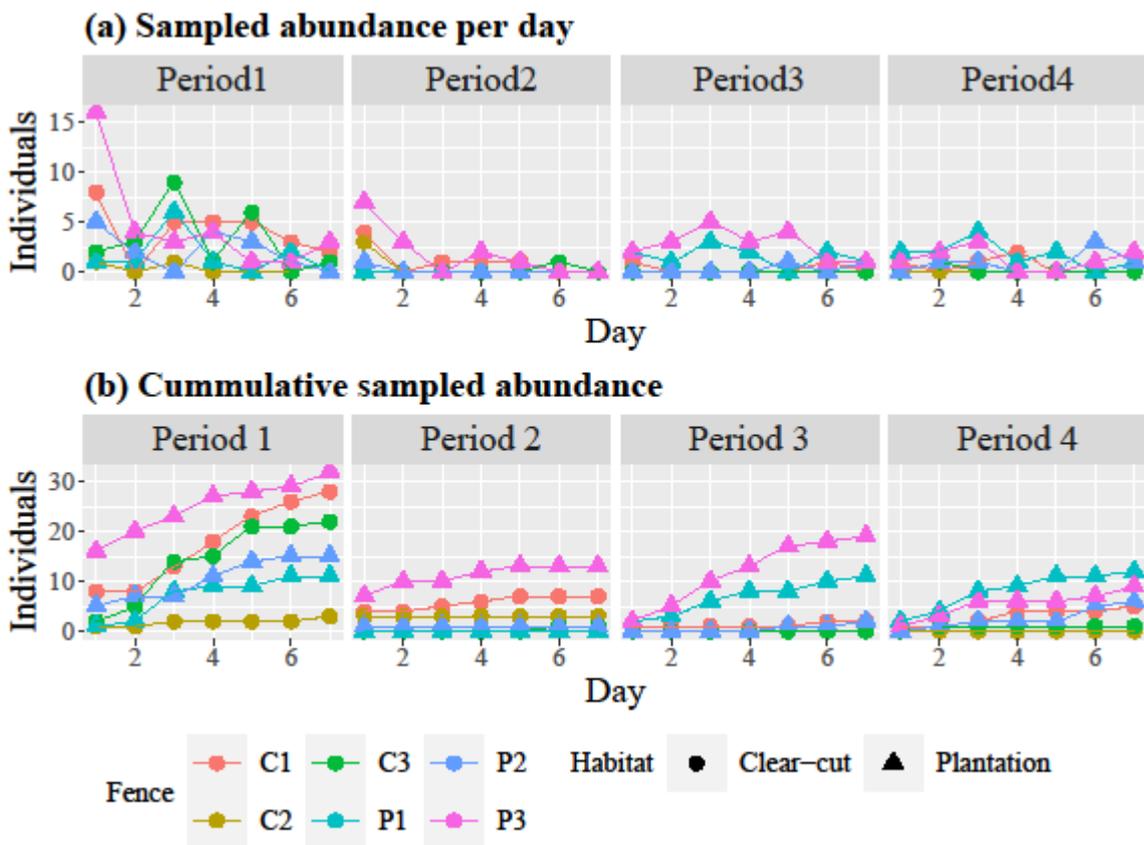
Day	Period			
	Period 1	Period 2	Period 3	Period 4
Day 1	5.5 (1-16)	2.5 (0-7)	0.8 (0-2)	0.7 (0-2)
Day 2	1.7 (0-4)	0.5 (0-3)	0.7 (0-3)	1.0 (0-2)
Day 3	4.0 (0-9)	0.2 (0-1)	1.3 (0-5)	1.5 (0-4)
Day 4	2.5 (0-5)	0.5 (0-2)	0.8 (0-3)	0.5 (0-2)

Day 5	2.5 (0-6)	0.3 (0-1)	0.8 (0-4)	0.3 (0-2)
Day 6	1.2 (0-3)	0.2 (0-1)	0.7 (0-2)	0.7 (0-3)
Day 7	1.2 (0-3)	0.0 (0-0)	0.5 (0-1)	0.8 (0-2)

224 Values indicate the mean value of daily sampled abundance and using 9 pitfall traps.

225 Values with parenthesis indicate the max and minimum values of the sampled abundance.

226



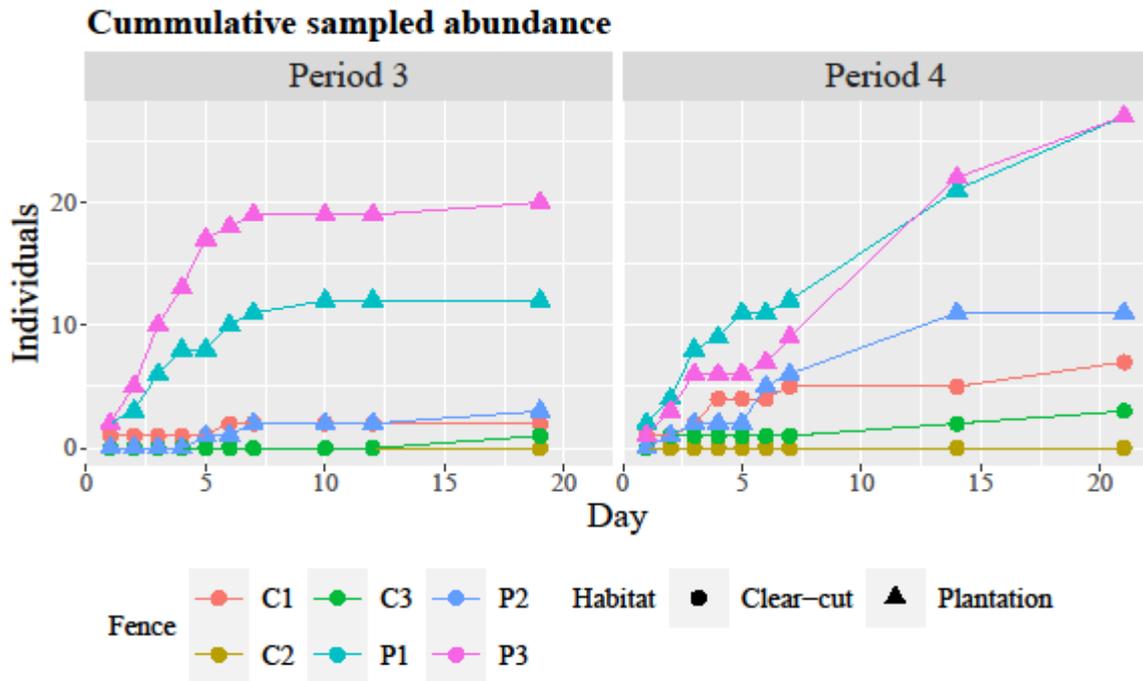
227

228 Figure S4-1 Sampled abundance (a) and cumulative sampled abundance (b) of *P.*

229 *thunbergi*

230 Circles and triangles indicate clear-cut and unharvested plantation sites, respectively.

231 Different colors indicate different fences.



232

233 Figure S4-2 Cumulative sampled abundance of *P. thunbergi* on days 1-19 and 1-21 in  
 234 periods 3 and 4

235 Circles and triangles indicate clear-cut and unharvested plantation sites, respectively.

236 Different colors indicate different fences. In periods 3 and 4, the survey was continued  
 237 after the 7th day, and samples were collected on days 10, 12, and 19 in period 3 and on  
 238 days 14 and 21 in period 4. We collected 5 and 42 individuals of *P. thunbergi*  
 239 additionally, respectively.

240

241 **Appendix 5 Weather factors (soil and air temperature and precipitation)**

242 Daily mean values of soil temperature were higher in the clear-cut site than at  
 243 the unharvested plantation site, and standard deviations were also relatively higher in  
 244 the clear-cut site (Table S5). The ranges of daily mean and maximum values of soil  
 245 temperature were overlapped among survey periods, although the minimum values in  
 246 period 3 were relatively lower than those in other periods (Figure S5-1). The daily mean  
 247 values of soil temperature were correlated with its maximum and minimum values,  
 248 respectively (Figure S5-2).

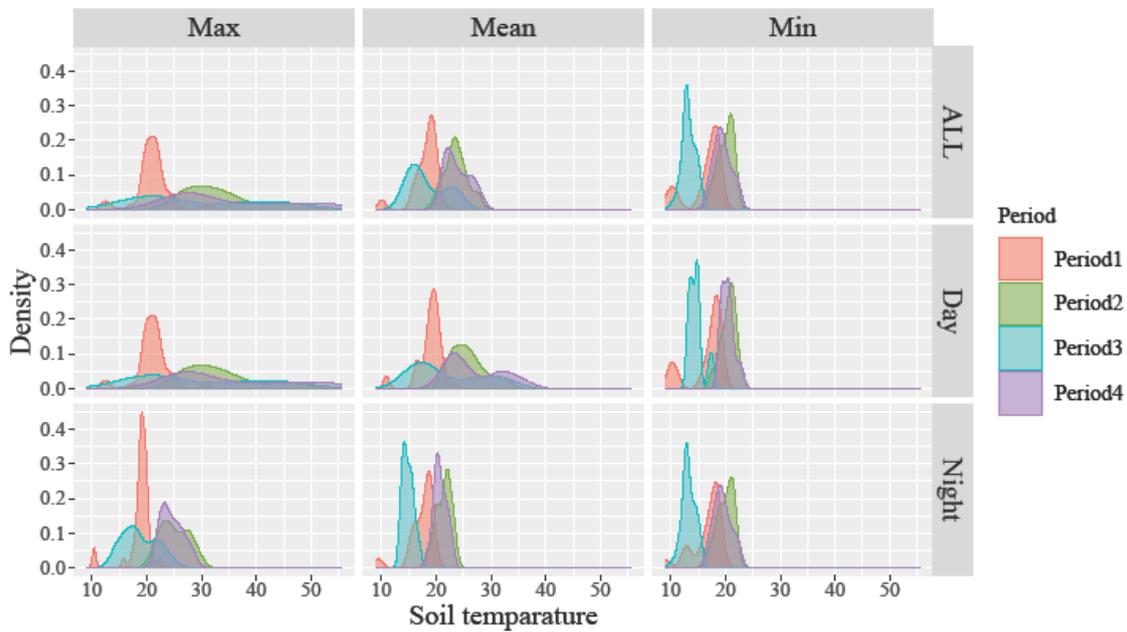
249 Comparing the mean values of soil temperature among the all-day, day, and  
 250 night time, the difference in soil temperature in the all-day time reflected the soil  
 251 temperature in day time (from 6:00 to 18:00), and the difference among fences was  
 252 small in night time (from 18:00 to 6:00) (Figure S5-3). We also examined air  
 253 temperature at several fences. The daily mean values of air temperature had similar  
 254 patterns with those of soil temperature (Figure S5-4).

255

256 Table S5 Daily mean values and its standard deviation of soil temperature

Period	Clear-cut		Plantation	
	Mean	SD	Mean	SD
Period1 2018 June	18.90	1.33	17.38	2.73
Period2 2018 July	25.04	1.54	22.15	1.34
Period3 2019 June	20.95	2.98	15.74	1.14
Period4 2019 July	25.79	1.48	22.00	0.67

257



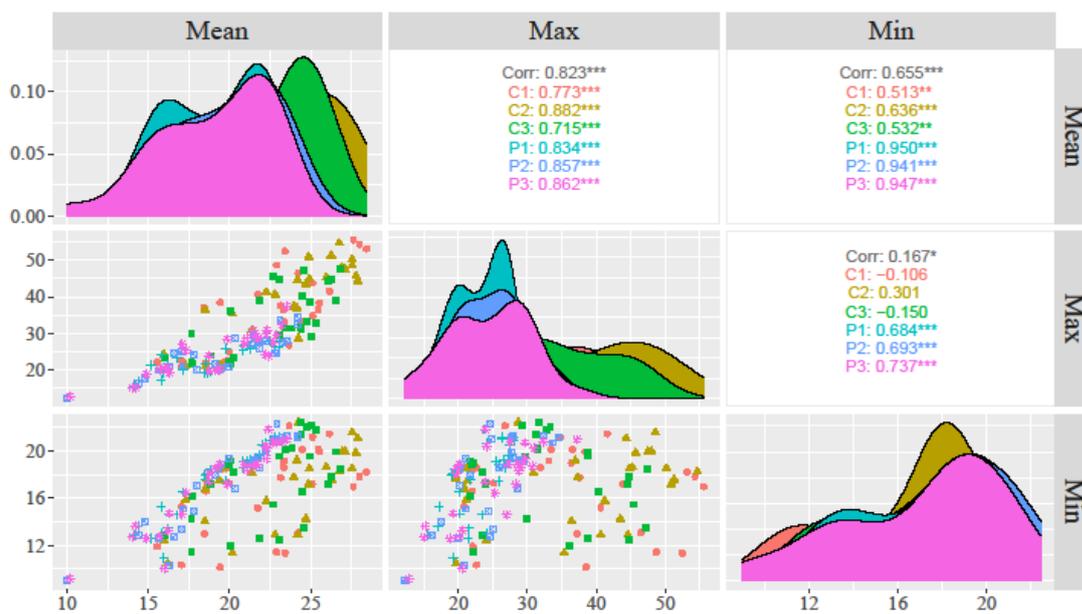
258

259 Figure S5-1 Distribution of mean values of soil temperature

260 All: all-day time (24h), Day: day time (6:00-18:00, 12h), Night: night time (18:00-6:00,

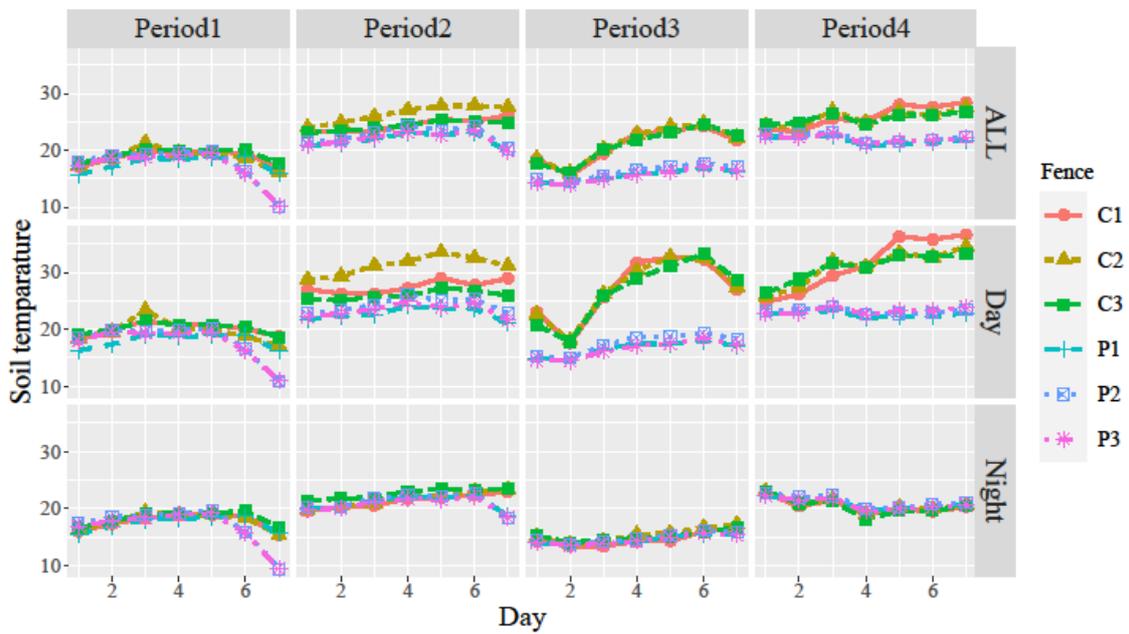
261 12h)

262

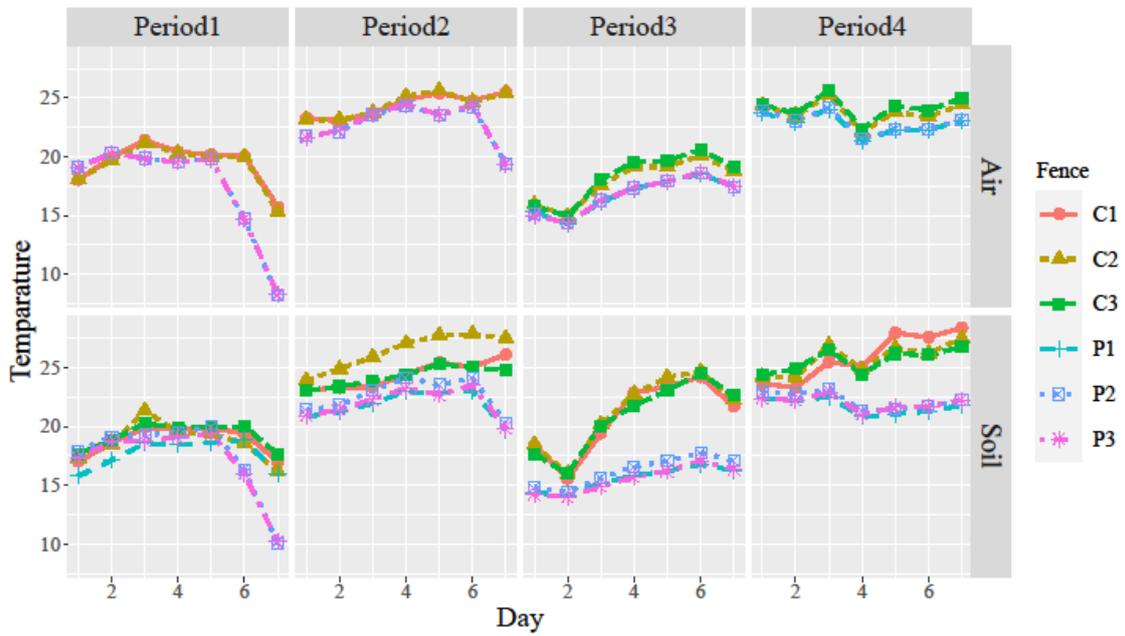


263

264 Figure S5-2 Correlation of mean values of soil temperature in the all-day time with  
 265 maximum and minimum values  
 266 Different symbols indicate each fence. The values at the right panels indicate the  
 267 correlation coefficients.  
 268



269  
 270 Figure S5-3 Mean values of soil temperature at the all-day, day, and night time.  
 271 All: all-day time (24h), Day: day time (6:00-18:00, 12h), Night: night time (18:00-6:00,  
 272 12h). Different symbols indicate each fence.  
 273

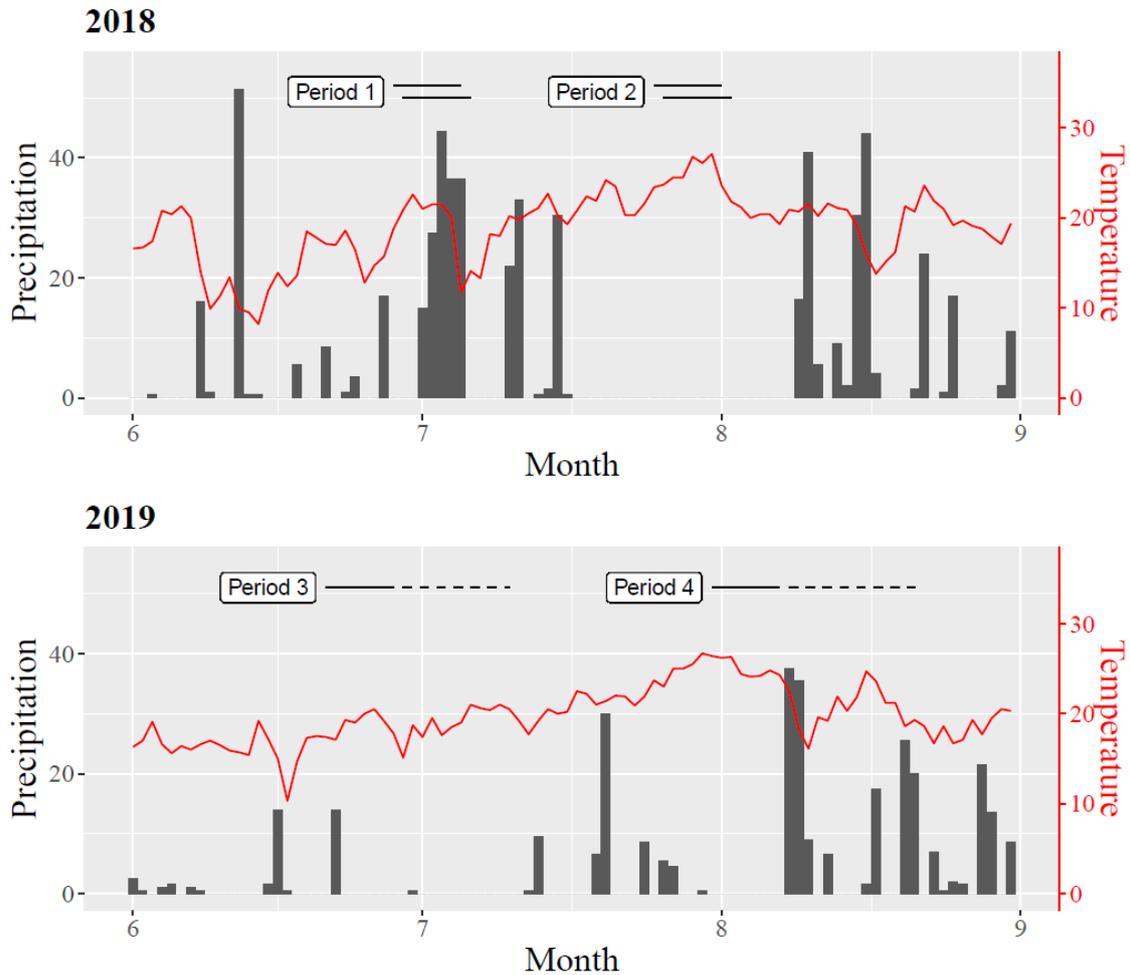


274

275 Figure S5-4 Daily mean values of air and soil temperature

276 Values indicate daily mean values. Different symbols indicate each fence.

277



278

279 Figure S5-5 Daily total precipitation and mean temperature of the study area during the  
 280 surveyed period

281 Black bars and red lines indicate daily total precipitation and mean temperature,  
 282 respectively. These data were obtained from the Ashibetsu observatory (Japan  
 283 Meteorological Agency 2021), which is the closest study area. Horizontal solid and  
 284 dashed lines indicate 7-day continuous sampling (at periods 1-4) and extended sampling  
 285 (at only periods 3-4), respectively. In period 1, sampling at grids C1, C2, C3, and P1  
 286 started on 28 June 2018 and ended on 5 July 2018, and sampling at grids P2 and P3  
 287 started on 29 June 2018 and ended on 6 July 2018. In period 2, sampling at Clear-cut  
 288 site (C1, C2, and C3) and Plantation site (P1, P2, and P3) started on 25 and 26 July

289 2018 and ended on 1 and 2 August 2018, respectively. In period 3, sampling started on  
290 21 June 2019 and ended on 10 July 2019. In period 4, sampling started on 31 July 2019  
291 and ended on 21 August 2019.

292

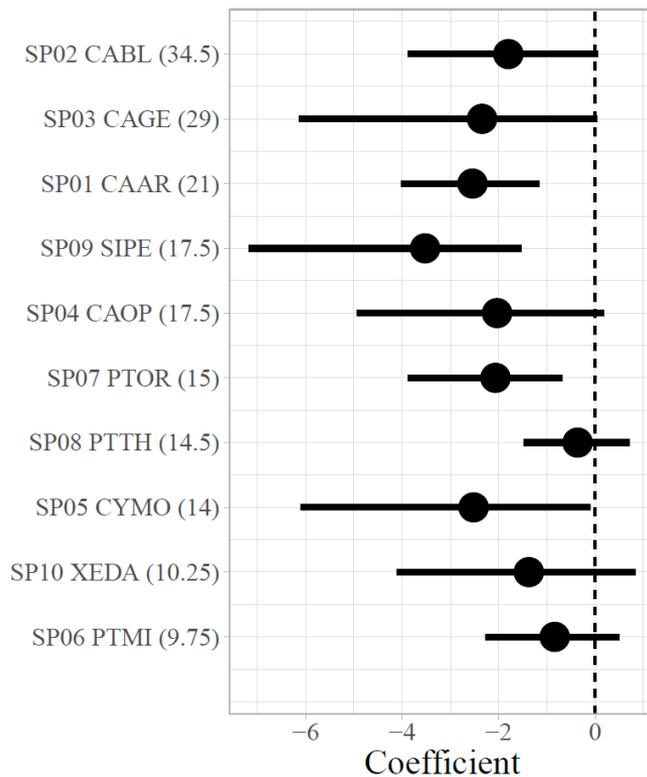
293 Reference

294 Japan Meteorological Agency (2021) Historical weather data,  
295 <http://www.data.jma.go.jp/obd/stats/etrn/index.php>.

296

297 **Appendix 6 Relationship between body size and coefficients of the effects of fences**  
298 **on sampled abundance**

299



300

301 Figure S6 body size and coefficients of the effects of fences on sampled abundance of  
302 each species

303 Y-axis indicates each species, sorted from top to bottom by body size (numbers in  
304 parentheses next to species names represent the mean body length). X-axis indicates  
305 coefficient of the effects of fences on sampled abundance. Black circles indicate mean  
306 values, and horizontal lines represent 95% credible intervals of the coefficients. Mean  
307 body length of each species was calculated by averaging the maximum and minimum  
308 sizes that reported by Ueno, Kurosawa and Sato (1985).

309

310 Reference

311 Ueno, S., Kurosawa, Y. & Sato, M. (1985) *The Coleoptera of Japan in Color, Vol. II (in*  
312 *Japanese)*. Hoikusha, Osaka.  
313

## 314 **Appendix 7 Results of removal sampling models**

315

316 1. the coefficients of removal sampling models with different model-fitting methods

317       The `ubms` and `RStan` models converged; however, a poor model fit for all

318 `ubms` models was suggested based on Pareto smoothed importance-sampling leave-

319 one-out cross-validation (PSIS-LOO). Most of the estimated coefficients were similar

320 among the different model-fitting methods (Figure S7-1 and S7-3). As well as

321 `unmarked` models in the main text, in other models in periods 3 and 4, the estimated

322 removal probability and the initial abundance estimates were overly low and high,

323 respectively (Figure S7-1c and S7-2). This suggests that these models in these periods

324 also could not estimate the removal probability and initial abundance precisely. Unlike

325 other models, in the `unmarked` Poisson model, the effect of soil temperature was

326 negative for the whole period data (Figure S7-1a) likely due to the coefficient of period

327 3.

328

329 2. Comparing classic removal sampling models

330       In the result of classic removal sampling models, estimated coefficients did not

331 differ among model-fitting methods (`unmarked`, `ubms`, and `RStan`) (Figure S7-4).

332 Moreover, the coefficients did not largely differ from that of the abovementioned

333 removal sampling models with the whole period data (Figure S7-3 and S7-4). This

334 suggests that effects of daily soil temperature on removal probability were not large

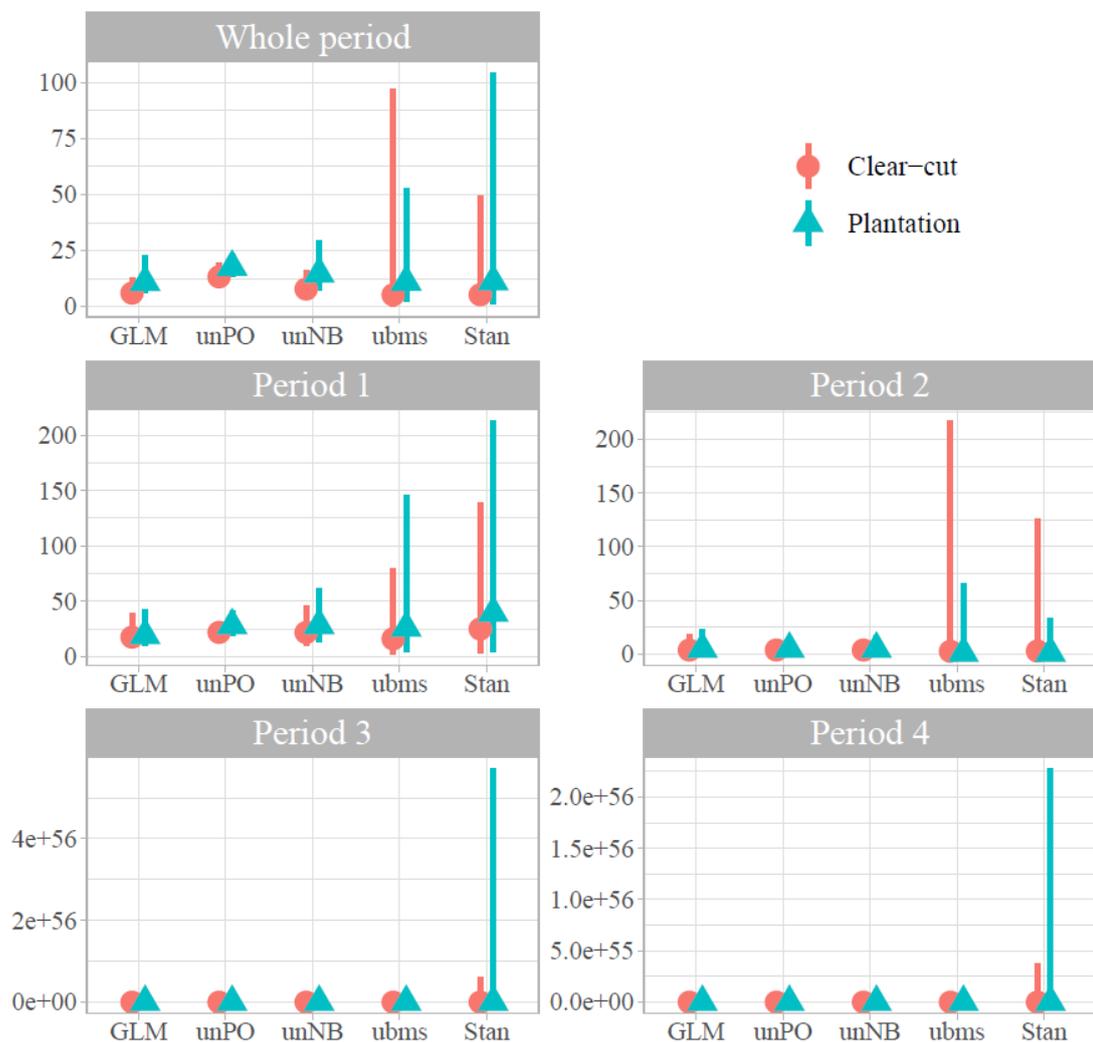
335 when considering over the whole survey periods, and this was consistent with our main

336 conclusion.

337



342 The mean values of estimates are denoted by circles or triangles, and their 95% CIs are  
343 denoted by vertical bars. (a) Effects of daily soil temperature on removal probability, (b)  
344 estimated per-day removal probability (given the mean soil temperature), and (c)  
345 abundance estimates are shown separately for different methods and distributions.  
346 Abbreviation: unmarked Poisson (unPO) and negative binomial (unNB), ubms  
347 (ubms), and RStan model (Stan). It is noted that unNB is the unmarked models in  
348 the main text. In periods 3 and 4, some mean values and their CIs were overly large, and  
349 thus, the figures show estimates with a limited range on the y-axis (0-200) to clarify the  
350 difference in estimates among models. Figures with the complete range of the y-axis are  
351 shown in Figure S7-2.



352

353 Figure S7-2 Initial abundances estimated by removal sampling models and expected

354 sampled abundance by GLMs (with full-range y-axis)

355 The estimates of GLM indicate the expected sampled abundance estimated by GLMs.

356 Other estimates (unPO, unNB, ubms, and Stan) indicate the initial abundance estimated

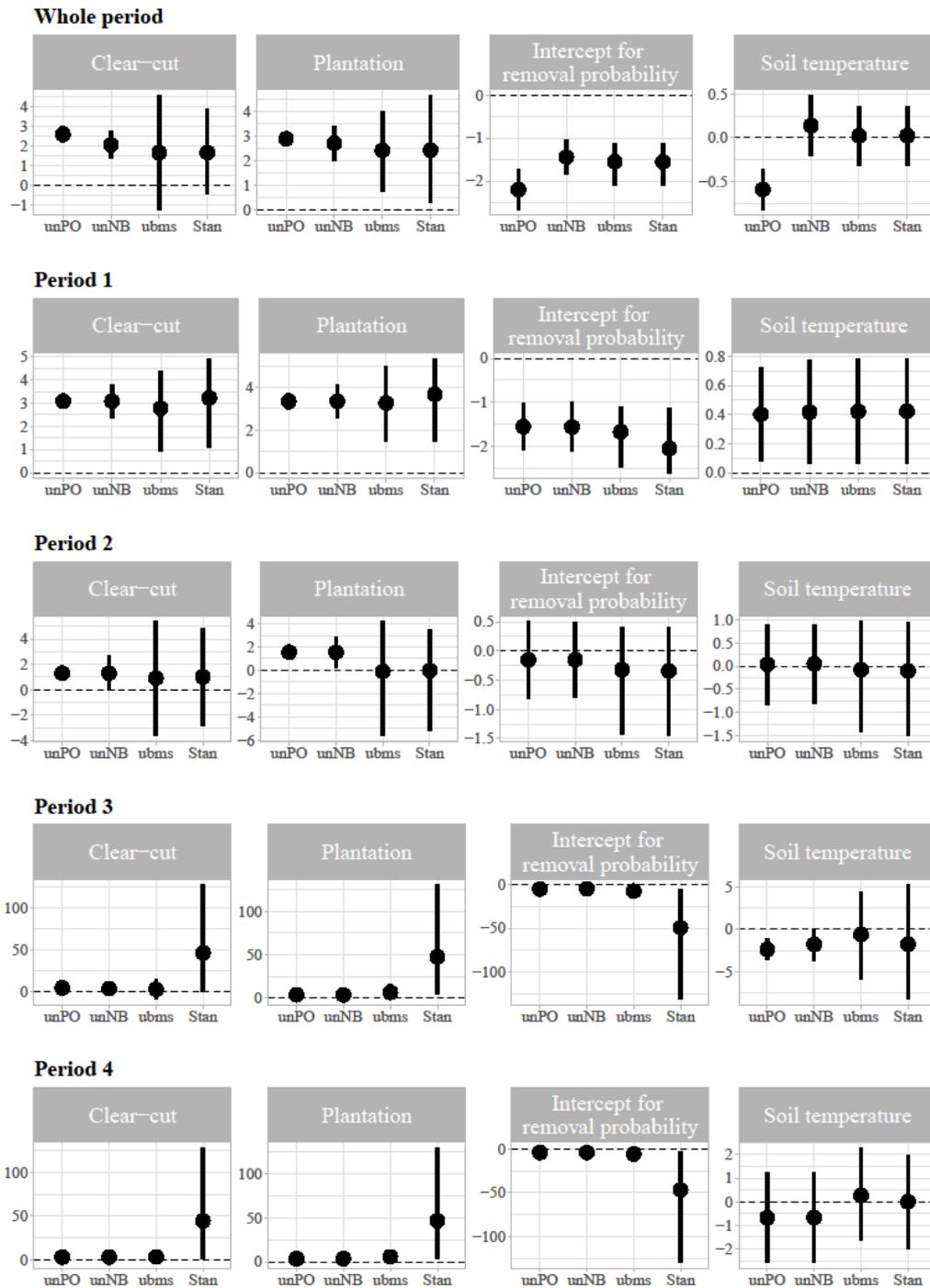
357 by each removal sampling model. unPO and unNB indicate unmarked model with a

358 Poisson and negative binomial distribution, respectively. Red circles and blue triangles

359 indicate the mean values of clear-cut and unharvested plantation sites, respectively.

360 Vertical bars show 95% CIs of estimates for GLMs and unmarked models or ubms  
 361 and RStan models, respectively.

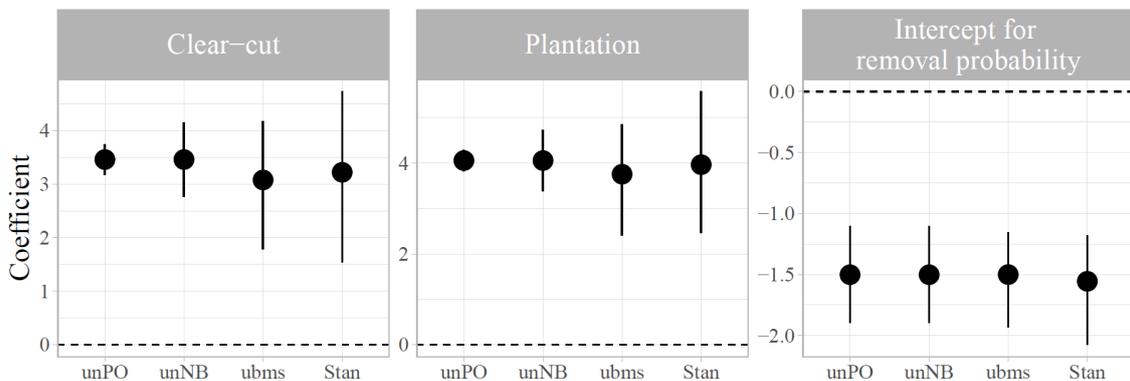
362



363

364 Figure S7-3 Estimated coefficients of parameters of removal sampling models

365 Black circles indicate mean values of the coefficients. Vertical bars show 95%  
 366 confidence intervals or credible intervals of the coefficients for unmarked models or  
 367 ubms and RStan models, respectively. Clear-cut and plantation indicate the initial  
 368 abundance of clear-cut and unharvested plantation sites, respectively. Soil temperature  
 369 indicates the coefficients of the effects of daily soil temperature on removal probability.  
 370 unPO and unNB indicate unmarked models with a Poisson and negative binomial  
 371 distribution, respectively.  
 372



373  
 374 Figure S7-4 Estimated coefficients of parameters of classic removal sampling models  
 375 (without considering the daily changes in removal probability)  
 376 Black circles indicate mean values of the coefficients. Vertical bars show 95% CIs of the  
 377 coefficients for unmarked models or ubms and RStan models, respectively. Clear-  
 378 cut and plantation indicate the initial abundance of clear-cut and unharvested plantation  
 379 sites, respectively. unPO and unNB indicate unmarked models with a Poisson and  
 380 negative binomial distribution, respectively.  
 381  
 382