

Title (a maximum of 20 words)

Examining the bias of pitfall traps with enclosure experiments and removal sampling

Running title (a maximum of 40 characters including spaces)

Bias of pitfall traps

Abstract (249/250 words)

1. Pitfall trapping is among the most widely used sampling methods for ground-dwelling beetles. However, the sampled ground-dwelling beetle abundance in pitfall traps may be biased, which can lead to difficulties when comparing the sampled abundance between different habitats.
2. To better understand the comparability of the sampled abundance, we focused on two sampling processes: a temporary migration process (referring to the individual movements into and out of the effective trap area) and a removal process (determining the number of individuals collected given the number of individuals within the trap area).
3. To examine the effects of the temporary migration process, we compared the sampled abundance of ten ground-dwelling beetle species inside and outside enclosures in two habitats (clear-cut and forested sites) and found that installing the enclosure decreased the sampled abundance in both habitats, but the effects did not differ among habitats.
4. To examine the effects of the removal process, we examined the effects of daily soil temperature on the daily sampled abundance (removal probability) of *Pterostichus thunbergi* using removal sampling within enclosures. We found that soil

temperature increased the removal probability, but this effect was limited and not consistent across survey periods.

5. The results suggest that the effects of the temporary migration process may not be habitat dependent and that the effects of the removal process may be small. Therefore, the traditional use of pitfall traps is, to some extent, a robust and comparable measure of the sampled abundance of ground-dwelling beetles among different habitats.

Keywords (maximum of 10 keywords)

activity density, Carabidae, catchability, detection probability, forest, population density, sampling bias, soil temperature, trap efficiency

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52 **5551 Words (abstract: 249, references: 1345)**

53 (Max 7000 words without Abstract and References - Insect Conservation and Diversity)

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Introduction

Pitfall trapping is one of the most widely used techniques for ground invertebrate sampling (Woodcock 2005; Henderson & Southwood 2016). As such, pitfall trap sampling has been used in various studies of ground-dwelling beetles, including the examination of their regional community and habitat preferences (e.g., Niemela & Halme 1992; Jouveau *et al.* 2019) and the anthropogenic impact on this group (e.g., Koivula *et al.* 2019; Yamanaka *et al.* 2021). However, the abundance sampled using pitfall traps is determined through the sampling process, which is influenced by several factors, including the sampling design and targeted site- and species-specific characteristics. Sampled abundance, therefore, does not equal the “true” abundance existing in a certain area during an arbitrary time period (hereafter, initial abundance; see Table 1 for glossary), and it should be interpreted with caution when comparing sampled abundance among different sampling designs and sites and periods with different environmental conditions (Adis 1979; Lang 2000; Woodcock 2005). For example, if the catchability of pitfall traps differs greatly between sampling designs, sites, or periods, the sampled abundance will be determined by the differences in the catchability rather than differences in the initial abundance, making it difficult to infer differences in the initial abundance between sites and periods.

The factors influencing the catchability of pitfall traps can be divided into two groups: (1) sampling design and (2) site- and species-specific characteristics that affect the movement activity of target species. There have been many studies on the effects of sampling design on catchability, for example, trap design (Luff 1975; Koivula *et al.* 2003), preservation liquid types (Koivula *et al.* 2003), the number and arrangement of traps installed (Lövei & Magura 2011; Engel *et al.* 2017), and the sampling period

(Schirmel *et al.* 2010; Lövei & Magura 2011). The differences in the sampling design can lead to differences in the sampled abundance even if the initial abundance is the same between sites. However, this problem can be addressed by using the same sampling method across target sites, that is, by disseminating a standard design (Brown & Matthews 2016).

For factors influencing catchability through changes in the movement activity of ground-dwelling beetles, there are extrinsic factors, such as temperature (Honek 1988; Engel *et al.* 2017), vegetation complexity (Honek 1988; Lang 2000; Koivula *et al.* 2003), and soil disturbances (e.g., digging-in effects; Schirmel *et al.* 2010). In addition, there are intrinsic factors, such as body size (Lang 2000; Engel *et al.* 2017), habitat preferences (Koivula *et al.* 2003; Růžicková *et al.* 2021), seasonal abundance/behavior changes (Baars 1979), and the hunger level of individuals (Raworth & Choi 2001). The influence of these factors can be expected to vary widely among target sites and species. Therefore, it is difficult to control differences in sampled abundance caused by changes in movement activity, which can become a serious problem, particularly when comparing the sampled abundance between different sites and/or periods. For example, even if the initial abundance is the same in two habitats, such as forest and grassland, the sampled abundance may be different between the habitats if the movement activity is higher in one habitat than the other. Similarly, differences in the movement activity may mask differences in the initial abundance even if the abundance truly differs between the two habitats.

In this study, to better understand the influence of differences in the movement activity of ground-dwelling beetles on the comparability of the sampled abundance, we deconstruct the pitfall trap sampling process into two subprocesses, namely, (i)

temporary migration and (ii) removal processes (Iknayan *et al.* 2014; Yamaura & Royle 2017). The temporary migration process is defined as individual movement into and out of the effective trap area. In this process, the movement activity increases/decreases the number of individuals moving into/out of the trap area. Differences in the movement activity between sites can lead to differences in the potential number of individuals who can be caught by traps between sites. The removal process determines the number of individuals collected based on the removal probability (Table 1) of the traps given the number of individuals exposed to the sampling within the trapping area. In this process, the movement activity increases/decreases the removal probability, and differences in the movement activity between sites can change the sampled abundance of each site.

The purpose of this study is to address the two sampling processes (temporary emigration and removal process) separately and examine whether sampled abundance can be compared among different habitats. First, to examine the effect of the temporary migration process on the sampled abundance, we installed enclosure fences (hereafter, fences) preventing the emigration of ground-dwelling beetles into/out of the trapping area and compared their sampled abundance among areas with and without fences and different habitats (clear-cut and forested areas). We hypothesized that the fences prevent individuals from moving into the effective trap areas, and thus, the sampled abundance within fences decreases more than that without fences. We also examined whether the fence effect differed between habitats. If the effect differs between habitats, the comparison of sampled abundance may be difficult because initial abundances are likely to differ between sites, even if the sampled abundances are the same. In fact, the movement activity of ground-dwelling beetles can change depending on its habitat (e.g., Niemela & Halme 1992; Allema *et al.* 2015; Růžicková *et al.* 2021).

Second, to examine the effect of the removal sampling process on sampled abundance, we conducted removal sampling within fences and examined the relationships between daily removal probability and sampled abundance. By installing fences, we excluded the effects of the temporary migration process on the sampled abundance and focused on the removal process. In this study, we focused on the effect of soil temperature on removal probability because temperature is one of the fundamental factors determining insect activity (e.g., Messenger 1959). High temperature can increase movement activity and sampled abundance of ground-dwelling beetles (Honek 1988; Honek 1997; Raworth & Choi 2001; Saska *et al.* 2013), and temperature is highly variable on a short time scale and is easy to monitor. If soil temperature affects removal probability, it may be problematic because sampled abundances would differ between sites and/or periods with different soil temperatures even if initial abundances are the same. We hypothesized that a high daily soil temperature increases the daily removal probability and indirectly increases the sampled abundance on each sampling day.

Materials and methods

Study area and experimental design

The study area is in the Irumukeppu Mountain area of Hokkaido in northern Japan (43°34'37" to 39°26'N, 142°05'27" to 09°33'E). The mean annual temperature and precipitation in this area are 7.4°C and 1141.6 mm, respectively, according to the data obtained from 1991 to 2020 by the Ashibetsu Observatory, which is the closest to the study area (Japan Meteorological Agency 2021). This area is dominated by conifer plantations, which mostly consist of Sakhalin fir (*Abies sachalinensis* [F. Schmidt]

Mast.). Harvesting significantly changes forest structure, including forest floor vegetation and soil conditions, and these changes can affect the initial abundance of each species and may change their movement activity. Therefore, we surveyed two forest stands, a clear-cut plantation site and an unharvested plantation site (Figure 1a; Appendix 1), to examine the effects of the two sampling processes on sampled abundance. At the clear-cut site, a mature plantation of Sakhalin fir was last harvested in 2015; hence, three to four years had passed by the time of the survey. The unharvested plantation site mainly consisted of mature Sakhalin fir, and its forest age is approximately 55 years (Akashi *et al.* 2017).

We installed three fences with galvanized iron panels on each site (clear-cut site: C1, C2, and C3; plantation site: P1, P2, and P3). The fences were established as a 5×5 m square with a height of 0.8 m above ground level, and their footings were buried 0.1 m deep under the surface (Figure 1b). We installed a grid of nine pitfall traps (3×3 traps) within each fence. Each pitfall trap consisted of a plastic cup with a diameter of 95 mm and a depth of 124 mm and contained 100% propylene glycol for preservation. The distances among traps and between traps and the nearest fence were 1 m and 1.5 m, respectively (Figure 1c). We also installed the same trap grid outside of the fences at least 5 m away from each fence. In all, we set 12 grids (2 sites \times 3 fences \times 2 grids; Figure 1a). Fences were installed in May 2018, and sampling was conducted four times during the following 14 months (period 1 [from June 28 or 29 to July 5 or 6, 2018], period 2 [from July 25 or 26 to August 1 or 2, 2018], period 3 [from June 21 to July 10, 2019], and period 4 [from July 31 to August 21, 2019]: Appendix 5). In each period, we collected ground-dwelling beetles each day for seven days. In periods 3 and 4, we conducted additional sampling after day seven. We collected samples on days 10, 12,

and 19 in period 3 and on days 14 and 21 in period 4 to examine the number of individuals who could be collected beyond the formal 7-day sampling period. Before each sampling period, we removed two of the panels, each 1.5 m in length, from two sides of each fence and allowed ground-dwelling beetles to move in and out of the fences to eliminate the effects of the fences on the initial abundance in each grid. In 2018 (periods 1 and 2) and 2019 (periods 3 and 4), there were 18 and 20 days between the two sampling periods, respectively (Appendix 5).

To examine the soil temperature across sites and periods, we installed a temperature data logger (Onset, HOBO UA-001-64) 1 cm underground at the center of each fenced area. Then, to examine whether soil temperature is correlated with air temperature, another type of temperature data logger (Onset, HOBO U23-004) with a solar radiation shield (Onset, HOBO, RS3-B) was set to observe air temperature. The logger was set in the corner of the fences 1 m above the ground surface to avoid the lower airflow caused by fences. The loggers were set at fences C1, C2, P2, and P3 in periods 1 and 2; C2, C3, P1, P2, and P3 in period 3; and C2, C3, P1, and P2 in period 4.

We identified the species in the collected samples based on the work of Ueno *et al.* (1985) and counted the number of individuals of each species collected on each sampling day and at each grid. In this study, we categorized each species as a macropterous, apterous, or brachypterous species by checking their wing types according to Shibuya *et al.* (2018) and used only apterous and brachypterous species (species having no, very short, or rudimentary wings) in the analysis. This is because our fences had no roof, and we could not exclude the possibility of movement of individuals in species that can fly over fences.

Data Analysis

Effects of the temporary migration process on sampled abundance

To examine the effects of the temporary migration process (i.e., installing fences) on the sampled abundance of pitfall traps, we examined whether the sampled abundance of ground-dwelling beetle species differed between areas with and without fences and whether the effect of fences differed between the two habitats. Here, we constructed abundance-based hierarchical community models (HCMs) according to the method of Yamaura *et al.* (2016). HCMs can be used to estimate multiple hierarchical parameters, allowing for simultaneous estimation of both community-level responses (common responses of targeted species) and species-level responses.

In this analysis, we focused on ten ground-dwelling beetles (nine carabid species and one carrion species: see Results section). For the HCMs, we pooled the sampled abundance for each species from the whole period (7 sampling days \times 4 periods) and used the total sampled abundance of each species as response variables following a negative binomial distribution. As explanatory variables for the species-level response variables, we used habitat type (binary variable where 0 denotes clear-cut and 1 denotes plantation), installed fence (binary variable where 0 denotes absent and 1 denotes present), and the interaction of the habitat type and installed fence variables. We also considered the log-transformed number of valid traps as an offset term because several traps without fences were disturbed by mammals. In addition, for the intercept and coefficients of each species, we set hyperparameters that were shared by species parameters to estimate species- and community-level responses (see details in Appendix 2).

To build the HCMs, we used R version 4.0.5 (R Core Team 2021) and the R

package “RStan” version 2.21.2 (Stan Development Team 2020). For the model, we used four chains that had 20,000 iterations following 5,000 burn-in with 20 thinning rate. We used the “R-hat” and “N_eff/N” statistics ($R\text{-hat} < 1.1$ and $N\text{-eff}/N > 0.1$) to assess the convergence of the model according to Gabry and Modrák (2019).

Effects of the removal process on sampled abundance

Removal sampling is a method in which individuals are sequentially removed from a population to estimate population size, assuming that if the population is closed and the removal probability remains at a certain level, the number of individuals removed during each sampling event will gradually decrease. While classic removal sampling models assume that the probability is constant over a given period and/or across different sampling sites (Rodriguez de Rivera & McCrea 2021), recently, several models have been presented that can take into account the changes in removal probability with site-dependent factors (Kéry & Royle 2015; Kellner *et al.* 2022).

To examine the effects of the removal process (i.e., the changes in the removal probability) on the sampled abundance, we examined whether the daily soil temperature increases the daily removal probability using “unmarked” removal sampling models (Fiske & Chandler 2011). In the analysis, we focused on the most dominant species, *Pterostichus thunbergi*. We used the daily number of individuals collected within the fences during the four survey periods consisting of seven continuous sampling days. We used two types of data: (1) the whole period data, which covered the four survey periods, and (2) each period data, which included data from each survey period because the effects of soil temperature may differ across survey periods, i.e., we built four separate models for the four periods.

In the `unmarked` models, we used the daily number of individuals collected as the response variable, assuming a negative binomial distribution, and categorical habitat type (clear-cut vs. unharvested plantation) and standardized daily mean soil temperature as explanatory variables for initial abundance and removal probability, respectively. That is, the models assumed that the initial abundance of each grid was determined by habitat type and that the daily removal probability (and the daily number of individuals collected) depended on daily soil temperature. For the construction of the models, we used the “`gmultmix`” function in the R package “`unmarked`” (Fiske & Chandler 2011). We compared (1) the effects of soil temperature on removal probability across survey periods, (2) the estimated removal probability across survey periods, and (3) the initial abundance across survey periods and between habitat types.

Several removal sampling models have been presented in addition to the `unmarked` models, including `ubms` models (Kellner *et al.* 2022). Therefore, we additionally constructed several models, including the classic models that assume the removal probability remains constant across survey sites and periods, with different model-fitting methods and/or different distributions that were followed by response variables (`unmarked` Poisson, `ubms`, and `Rstan` models; Appendix 3). We compared the results with those of the abovementioned `unmarked` models.

Moreover, we constructed generalized linear models (GLMs) that assumed that the response variables follow a negative binomial distribution. In the models, we used the total sampled abundance of *P. thunbergi* in each period and at each grid as response variables and habitat type as explanatory variables. Unlike the abovementioned `unmarked` models, the GLMs do not account for the removal process, and therefore, the abundance estimates would include the bias caused by the removal process. We

271 compared the effects of habitat type on the abundance estimates between GLMs and the
272 unmarked models to examine whether potential bias can impede the comparison of
273 abundance estimates between different habitats. If the effects of habitat type on the
274 abundance estimates were consistent between the GLMs and unmarked models, we
275 would interpret this as evidence that models that do not account for the removal process
276 could be used to obtain differences in the initial abundance between habitat types. For
277 the GLMs, the “glm.nb” function of the R package MASS was used.

278 To build the removal sampling models and GLMs, we used R version 4.0.5 (R
279 Core Team 2021). In all analyses, we used the 95% confidence interval or credible
280 interval (which is used in the Bayesian approach instead of a confidence interval) to
281 judge whether each variable was significant (hereafter, 95% CI). When the 95% CI of
282 the coefficient of each variable or estimated value did not contain 0, we considered the
283 variables to be significant.

Results

We collected a total of 606 individuals, consisting of ten apterous and brachypterous ground-dwelling beetle species (nine forest species and one nonforest species; Table 2), in four survey periods composed of seven continuous sampling days (not including extended surveys in periods 3 and 4). We collected 203 individuals of the most dominant species, *Pterostichus thunbergi*, within fences in four survey periods composed of seven continuous sampling days. We collected 111, 25, 33, and 34 individuals of *P. thunbergi* in period 1, period 2, period 3, and period 4, respectively (Appendix 4). Extension sampling in periods 3 and 4 collected an additional 5 and 42 individuals of *P. thunbergi*, respectively (Appendix 4). The daily mean values of soil temperature were relatively higher in the clear-cut site than in the unharvested plantation site, and the daily air temperature of each grid exhibited a trend similar to that of soil temperature (Appendix 5).

Effects of the temporary migration process on sampled abundance

We constructed HCMs using the sampled abundance of ten species (Table 2). For community-level hyperparameters, the coefficient of fence was significantly negative, and the 95% CI of the coefficients of habitat type and the interaction between habitat type and fences included zero (Figure 2). This suggested that, as the average response, the sampled abundance did not differ between habitat types but decreased within fences and that the fence effect did not differ between habitat types. At the species level, the coefficients of habitat type and fences varied among species. The sampled abundance of four species significantly differed between habitat types. Specifically, the abundance of *Carabus blaptoides rugipennis* (SP2) and *Xestagonum*

daisetsuzanus (SP10) was higher in the unharvested plantation site, while the abundance of *Carabus arcensis hokkaidoensis* (SP1) and *Pterostichus microcephalus* (SP6) was higher in the clear-cut site (Figure 2). The coefficients of fences were significantly negative for four species (*C. arcensis hokkaidoensis* [SP1], *Cychrus morawitzi morawitzi* [SP5], *Pterostichus orientalis antiquus* [SP7], and *Silpha perforata* [SP9]; Figure 2). Similar to the hyperparameter, the interaction term for habitat type and fences was not significant for any species (Figure 2). In addition, there was no apparent relationship between the coefficients of fences and the body size of each species (Appendix 6).

Effects of the removal process on sampled abundance

The estimated coefficients of removal sampling models were similar among the different model-fitting methods and the model with a different distribution (Appendix 7). In addition, in periods 3 and 4, the mean values and ranges of most coefficients were overly high and wide, respectively (Figure 3c and Appendix 7), indicating that the models in these periods could not estimate the removal probability and initial abundance precisely.

In unmarked models, positive effects of soil temperature on removal probability were found only in period 1; that is, such positive effects were not observed in other periods (Figure 3a). In addition, the estimated daily removal probability differed among the survey periods (Figure 3b). During the whole period and during periods 1 and 2, the daily removal probability was relatively high, and the range of mean values was from 10 to 50% (Figure 3b). Conversely, the estimated probability was overly low for periods 3 and 4, and the mean values were close to zero (Figure 3b). The 95% CI of the estimated initial abundance of *P. thunbergi* at clear-cut and plantation

sites overlapped in all periods (Figure 3c), suggesting that the initial abundance did not differ between habitat types.

Comparing the effects of habitat types on the abundance estimate between unmarked models and GLMs, in most cases, the effects were constant, that is, the abundance estimates of *P. thunbergi* were slightly higher at the plantation site than at the clear-cut site (Figure 3c). Thus, we concluded that the GLMs that do not account for the removal process could also show the differences in the initial abundance between habitat types. In addition, for the classic removal sampling models that do not account for the daily changes in removal probability, the estimated coefficients did not largely differ from those of unmarked models with the whole period data (Appendix 7). This also suggests that the effect of the removal process (effect of soil temperature on removal probability) was not substantial, especially when considered over the entire survey period.

Discussion

Effects of the temporary migration process on sampled abundance

We found that the sampled abundance of ground-dwelling beetles in pitfall traps decreased within fences. This result is consistent with those of previous studies demonstrating that the movement of carabids into and out of effective trap areas (i.e., the temporary migration process) can increase their sampled abundance (Desender & Maelfait 1986; Andersen 1995; Holland & Smith 1999). Furthermore, we did not find that the effects were dependent on habitat type. Therefore, it could be argued that the temporary migration process did not impede the comparison of the sampled abundance among different habitats.

In addition, the results presented above also suggest that the movement activity of each species did not differ between clear-cut and unharvested forests, although previous studies demonstrated that the movement activity of forest carabids was generally larger in open land than in woodland (Allema *et al.* 2015; Růžicková *et al.* 2021). At our clear-cut site, there was a high degree of logging residue (slash) (Appendix 1). Slash increases the microhabitat complexity of harvested areas, which potentially provides a hiding site for forest ground-dwelling beetles and thus could maintain the community composition (e.g., Gunnarsson *et al.* 2004; Nittérus *et al.* 2007). Therefore, harvested areas with plenty of slash could be a suitable habitat for forest ground-dwelling beetles compared to other open lands, such as arable lands, and this may result in the small difference in the movement activity between the two habitats.

Previous studies suggested or assumed that large species have higher movement activity and are likely to be captured more than small species (Lang 2000;

Engel *et al.* 2017). However, our results also revealed that installing fences decreased the sampled abundance of each species, but the degree of the decrease was not correlated with the body size of the species (Appendix 5). Therefore, the negative fence effects were not explained by differences in body size. This was inconsistent with the findings of previous studies that pitfall trap bias can be corrected by using the body size of each species (Hancock & Legg 2012; Engel *et al.* 2017). This may be because our targeted species were relatively large (mean size: 9.75-34.5 mm), and most of the species had high movement activity rates.

Effects of daily soil temperature on the removal probability

Saska *et al.* (2013) conducted a meta-analysis on the effect of weather on the sampled abundance of carabids in several regions and demonstrated that while temperature generally increased the sampled abundance, the effects varied among species. We found that soil temperature increased the removal probability of pitfall traps and indirectly increased the sampled abundance of *Pterostichus thunbergi* but that positive effects were found in only one survey period (period 1: June 2018). With respect to *P. thunbergi*, therefore, it could be argued that the effects of soil temperature on the removal probability were not consistent throughout the survey periods. In period 1, the daily mean values of soil temperature and their ranges were not substantially different from those of the other periods (Appendix 5), but the precipitation was higher than that in the other periods (Appendix 5). Although other weather factors, such as precipitation and soil humidity, can affect the removal probability (Mitchell 1963; Saska *et al.* 2013), the effect of soil temperature appeared to be small at least in our study, which focused on a restricted area and season (i.e., hot summer season). In fact, Saska *et*

al. (2013) suggested that temperature bias may be small on a short time scale and/or in the absence of distinct temporal trends in temperature.

In addition, because we did not examine differences in vegetation structure among sites, which can affect soil temperature, we could not clearly distinguish between the effects of soil temperature and vegetation structures on the removal probability. However, we dealt with daily changes in soil temperature, which can change greatly on a short time scale compared to vegetation structure; therefore, we concluded that our analysis mainly examined the effects of soil temperature changes on sampled abundance rather than those of vegetation.

We could not estimate the removal probability and initial abundance precisely for periods 3 and 4 (June and July 2019; Figure 3bc). In these periods, the total number of collected individuals of *P. thunbergi* was relatively small (Appendix 4). Moreover, the number of collected individuals was highest on the third day of periods 3 and 4 (Appendix 4), despite the general expectation that the daily number of collected individuals would be high on the first day and decrease on subsequent days. It is possible that the small total number and irregular patterns of collected individuals in these periods made estimating the coefficients imprecise. Comparing the results of several removal sampling models, we did not find substantial differences in the coefficients among the models. Developing customized complex models, such as the RStan models (Appendix 3), required additional skill, knowledge, and effort to write code. Thus, `unmarked` and `ubms` models are recommended for estimating the initial abundance and removal probability, at least in our case.

Comparing abundance estimates among different habitats

Comparing the abundance estimates between the GLMs and unmarked models, we found that the habitat type effects on the abundance of *P. thunbergi* were consistent (Figure 3c), suggesting that the effects of the removal process on sampled abundance were small. In addition, we found that the temporary migration process did not depend on habitat differences (Figure 2). Therefore, we could compare the sampled abundance without considering the removal process to represent the difference in the initial abundance of *P. thunbergi*.

Furthermore, our results infer abundance changes in ground-dwelling beetles after harvest. Studies reported that the sampled abundance of some forest carabids did not decrease for several years after harvest (e.g., Work *et al.* 2010; Yamanaka *et al.* 2021). There are two possible hypotheses for this finding. One, the changes in the movement activity of carabids in harvested areas masked the decreased initial abundance, and two, the initial abundance did not decrease immediately after the harvest. Our results support the latter hypothesis, suggesting that the initial abundance of these forest carabids did not change after harvest and that the beetles could survive in clear-cut areas for at least a few years.

Conclusion

We focused on two processes potentially affecting pitfall trap sampling and found that the sampled abundance of each species tended to increase in the temporary migration process, but its effect did not differ among habitat types. We also found that in the removal process, soil temperature could affect the removal probability and sampled abundance, but its effect was inconsistent and relatively small over the entire survey period. Therefore, we concluded that the traditional use of pitfall traps is, to

some extent, a robust measure, particularly when comparing the sampled abundance of each species among neighboring habitats, such as in the case of our study.

Our study also has several limitations. First, in our removal sampling models, we only examined one species' response and the effect of soil temperature. Removal probability and its influencing factors can vary among species (Saska *et al.* 2013; Růžicková *et al.* 2021) and may depend on individual and habitat conditions, such as starvation and prey density distribution (Raworth & Choi 2001). Second, we could not account for the possibility that vegetation density could hamper the movement of ground-dwelling beetles and decrease the removal probability (Melbourne 1999; Thomas *et al.* 2006). Third, we could not examine the reason that the removal probability and effect of soil temperature changed across the survey periods. Future studies should focus on multispecies responses, the effects of structural complexity, and different spatial-temporal ranges to compensate for our study's limitations and to gain a full understanding of the limitations and effective usage of pitfall traps.

Author contributions

Conceptualization: SY, KO; Methodology: SY; Investigation: SY; Formal analysis: SY, YY; Writing - original draft: SY, YY; Writing - review & editing: SY, YY, KO; Funding acquisition: SY, KO.

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Conflict of Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

Figure captions

Figure 1 Study sites and enclosure experiment

(a) Location of study sites and sampling grids. C1-3 and P1-3 indicate sampling grids in clear-cut and unharvested plantation sites, respectively. (b) Enclosure fences installed. (c) Spatial arrangement of a sampling grid.

Figure 2 Coefficients of hyperparameters and species parameters of the HCMs

Black circles and horizontal bars indicate mean values and 95% CIs of coefficients, respectively. Dashed lines indicate zero values. Intercept: coefficient indicating log-scale sampled abundance at clear-cut site; Habitat type: changes in sampled abundance at unharvested plantation compared to clear-cut site; Fence: the changes in sampled abundance due to installing fences; Habitat type \times Fence: the interaction between habitat type and fences. When the 95% CIs of the coefficients of Habitat types, Fence, and Habitat \times Fence were above (below) zero, we considered that the abundance significantly increased (decreased) in the plantation site, within fences, and within fences in the plantation site, respectively. Hyperparameter: the community-level response, SP01 CAAR: *Carabus arcensis hokkaidoensis*, SP02 CABL: *Carabus blaptoides rugipennis*, SP03 CAGE: *Carabus gehinii gehinii*, SP04 CAOP: *Carabus opaculus opaculus*, SP05 CYMO: *Cychrus morawitzi morawitzi*, SP06 PTMI: *Pterostichus microcephalus*, SP07 PTOR: *Pterostichus orientalis antiquus*, SP08 PTTH: *Pterostichus thunbergi*, SP09 SIPE: *Silpha perforata*, SP10 XEDA: *Xestagonum daisetsuzanus*.

Figure 3 Results of removal sampling models (unmarked model) with the whole period data (Whole period) and each period data (Periods 1, 2, 3, and 4) in the analysis of *P. thunbergi*

The mean values of estimates are denoted by circles or triangles, and their 95% CIs are denoted by vertical bars. (a) Effects of daily soil temperature on removal probability, (b) estimated per-day removal probability (given the mean soil temperature), and (c) abundance estimates are shown separately for different survey periods. In (c), we compared the abundance estimates between unmarked models (unmarked) and generalized linear models that do not consider the removal process (GLM).

Table captions

Table 1 Glossary of terms used in this study

Table 2 Mean value of the sampled abundance of each species per day at clear-cut and unharvested plantation sites

Values indicate the mean value of sampled abundance per day and per 1 pitfall trap. Values with parentheses indicate the standard deviations. These values show the sampled abundance of 7 continuous days in four survey periods (not including extended sampling in periods 3 and 4). Mean values were calculated with the data of whole periods (each 7-sampling day in June and July 2018 and 2019). The scientific name of each species was in accordance with Suzuki (2022). Carabid species are categorized into forest and nonforest species according to the literature (Hori 2001; Hori 2003; Hori 2012; Yamanaka *et al.* 2021). Silphid species (*Silpha perforata*) are categorized as

521 forest species according to Katakura *et al.* (1986). *Pterostichus microcephalus* is a wing
522 polymorphic species (Shibuya *et al.* 2018), but most of the individuals we captured
523 were brachypterous (only 2 of 18 individuals were macropterous). Thus, *P.*
524 *microcephalus* was included in the analysis.
525

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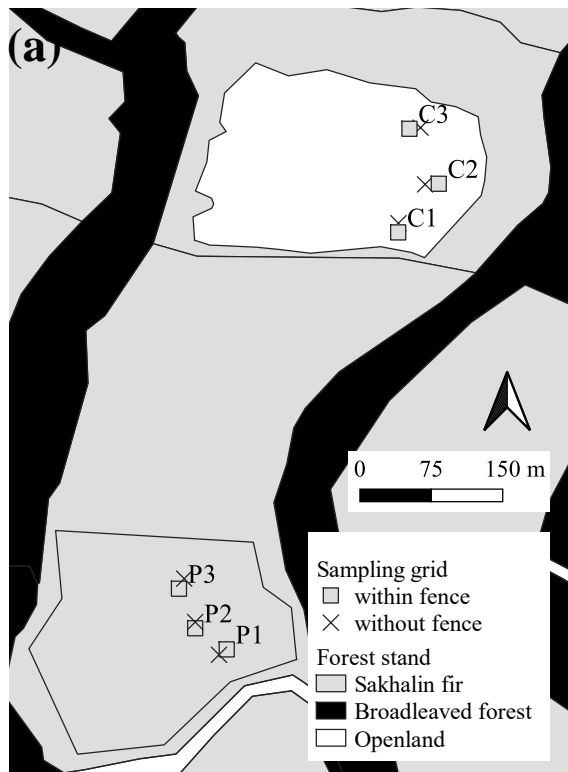
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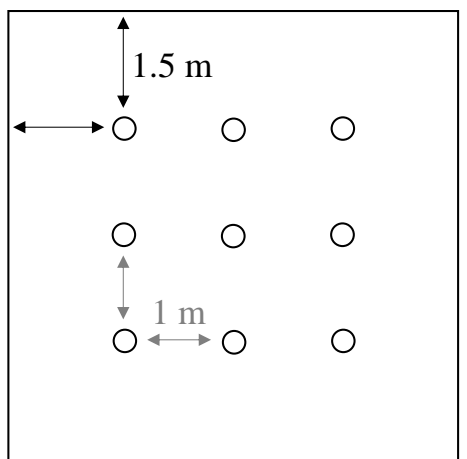
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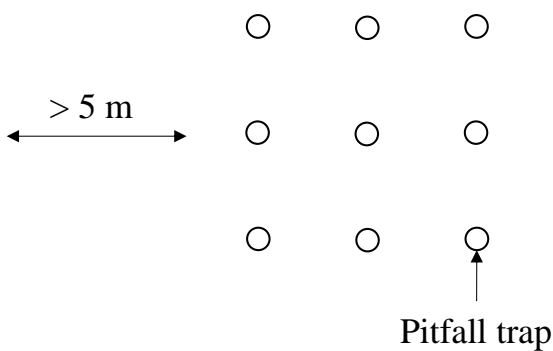


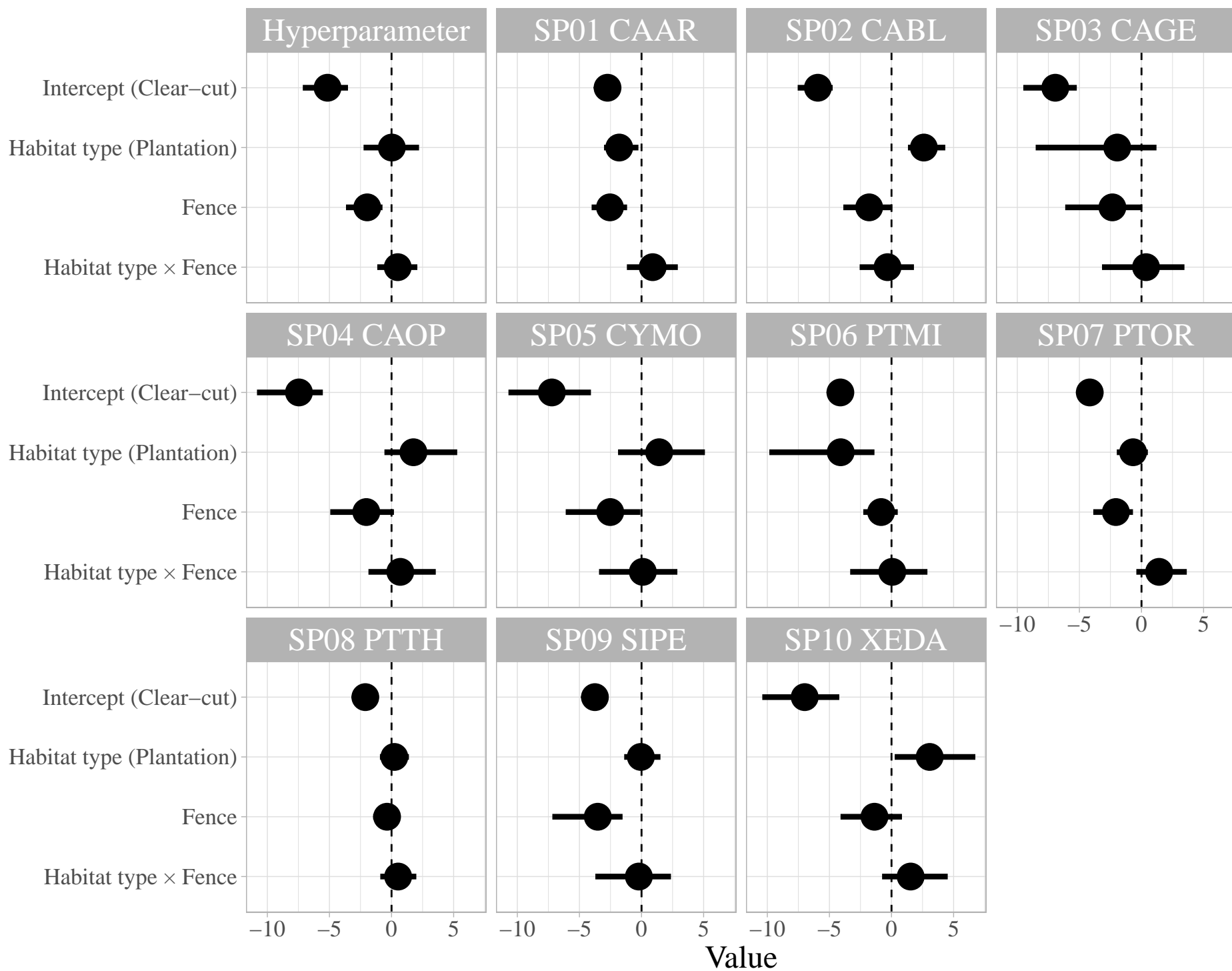
(c)

Sampling grid within fence

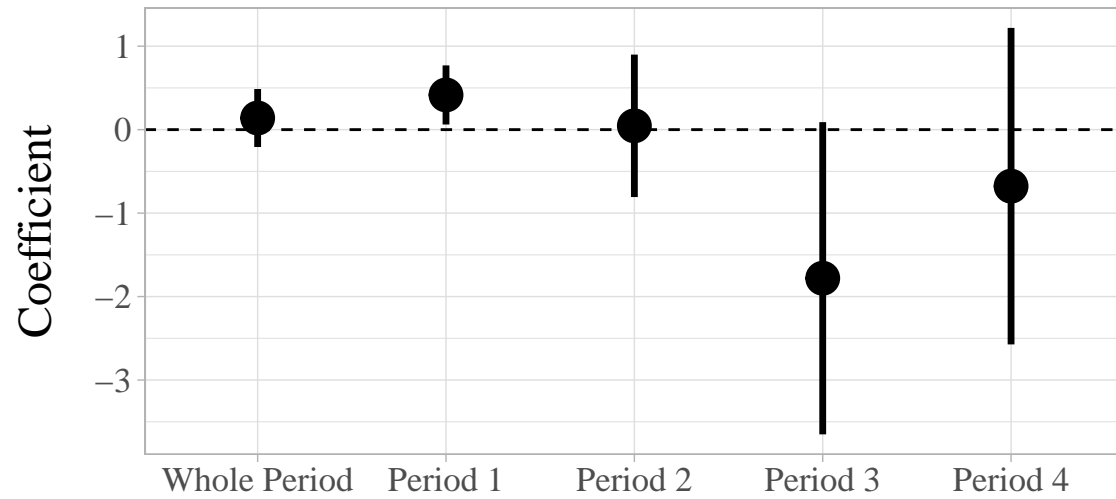


Sampling grid without fence

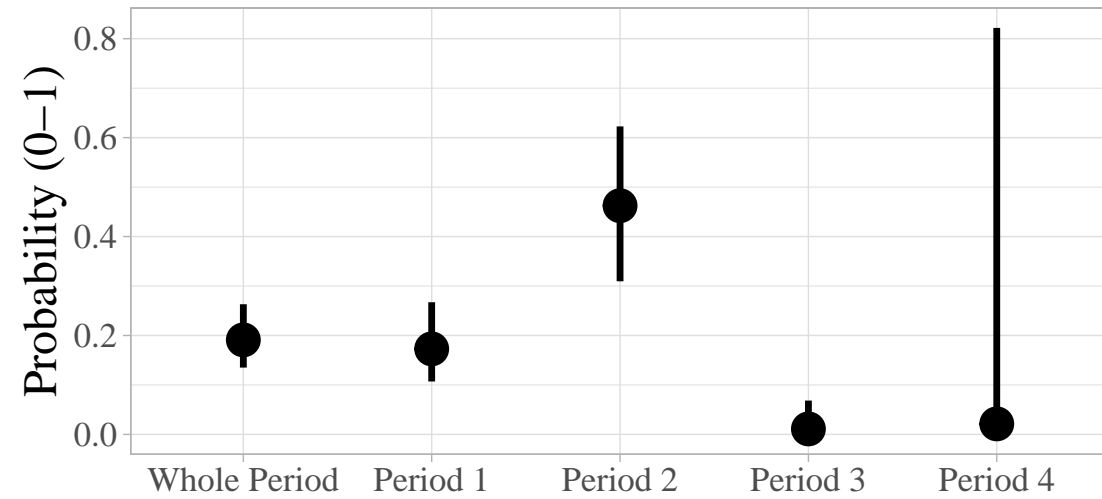




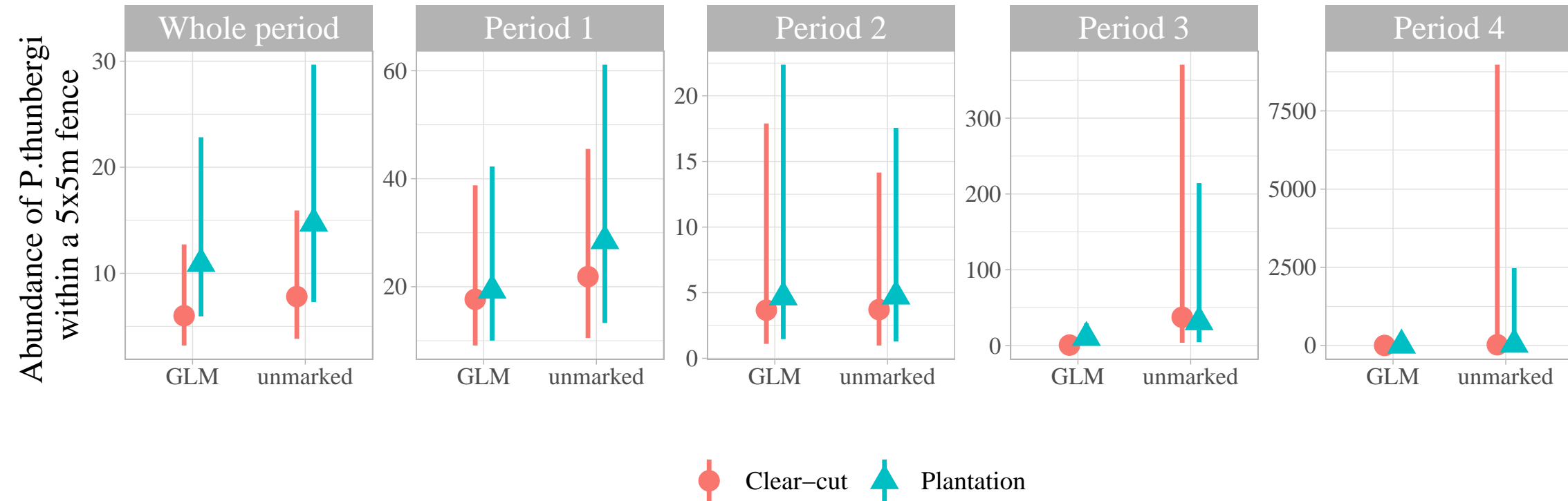
(a) Effect of soil temperature on removal probability



(b) Estimated removal probability



(c) Abundance estimates



1 **Table**

2 Table 1 Glossary of terms used in this study.

Terms	Explanation
Initial abundance	<ul style="list-style-type: none"> The number of individuals present in a certain area during an arbitrary time. In this study, we regarded initial abundance as the number of individuals present within the fences when the fences were shut down, and we estimated it using removal sampling.
Sampled abundance	<ul style="list-style-type: none"> The number of individuals actually collected by pitfall traps. It is determined through the sampling process of pitfall traps, which is influenced by various factors (sampling design and targeted site- and species-specific characteristics).
Removal probability	<ul style="list-style-type: none"> The probability that individuals within effective trap area are caught and removed by pitfall traps during a certain period. It depends on a sampling design and the movement activity of each species. In this study, we regard the removal probability as the probability that an individual is collected by traps per day in an enclosure experiment in the absence of a temporary migration process.
Movement activity	<ul style="list-style-type: none"> It represents the degree of activity of each species, which is dependent on intrinsic factors (e.g., their body size, species-specific movement distance or frequency) and extrinsic factors (e.g., soil temperature, vegetation structures). It affects both the temporary migration and removal processes.
Sampled abundance within fence	<ul style="list-style-type: none"> The total number of individuals collected within fences. It is determined by the initial abundance and removal probability within fences, and thus it is expected to be lower than the initial abundance but close to the initial abundance with a longer survey period.

	<ul style="list-style-type: none"> • The movement of individuals over fences is prevented, and thus sampled abundance within fence is not affected by temporary migration process.
Sampled abundance without fence	<ul style="list-style-type: none"> • The total number of individuals collected without fences. • It is the common measure of comparison in pitfall trap survey.

4 Table 2 Mean value of the sampled abundance of each species per day at clear-cut and unharvested plantation sites

5

ID	Species name	Abbreviation	Habitat	Taxa	Clear-cut		Plantation	
					Within fence	Without fence	Within fence	Without fence
1	<i>Carabus arcensis</i> <i>hokkaidoensis</i>	CAAR	Forest	Carabid	0.004 (0.0072)	0.0757 (0.0625)	0.0026 (0.0092)	0.0093 (0.0143)
2	<i>Carabus blaptoides</i> <i>rugipennis</i>	CABL	Forest	Carabid	0.0013 (0.0046)	0.0013 (0.0046)	0.004 (0.0072)	0.0384 (0.0386)
3	<i>Carabus gehinii gehinii</i>	CAGE	Forest	Carabid	0 (0)	0.0013 (0.0047)	0 (0)	0 (0)
4	<i>Carabus opaculus opaculus</i>	CAOP	Forest	Carabid	0 (0)	0 (0)	0.0013 (0.0046)	0.004 (0.0072)
5	<i>Cychrus morawitzi morawitzi</i>	CYMO	Forest	Carabid	0 (0)	0 (0)	0 (0)	0.004 (0.0137)
6	<i>Pterostichus microcephalus</i>	PTMI	NonForest	Carabid	0.0093 (0.0197)	0.0146 (0.023)	0 (0)	0 (0)
7	<i>Pterostichus orientalis</i> <i>antiquus</i>	PTOR	Forest	Carabid	0.0013 (0.0046)	0.0186 (0.0309)	0.0053 (0.0078)	0.0066 (0.0106)
8	<i>Pterostichus thunbergi</i>	PTTH	Forest	Carabid	0.0952 (0.1463)	0.1067 (0.1492)	0.1733 (0.1401)	0.1455 (0.0593)
9	<i>Silpha perforata</i>	SIPE	Forest	Silphid	0 (0)	0.0274 (0.0448)	0 (0)	0.0238 (0.0441)

10	<i>Xestagonum daisetsuzanus</i>	XEDA	Forest	Carabid	0 (0)	0 (0)	0.0278 (0.0479)	0.0132 (0.0223)
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6 Values indicate the mean value of sampled abundance per day and per 1 pitfall trap. Values with parentheses indicate the standard
7 deviations. These values show the sampled abundance of 7 continuous days in four survey periods (not including extended sampling in
8 periods 3 and 4). Mean values were calculated with the data of whole periods (each 7-sampling day in June and July 2018 and 2019).
9 The scientific name of each species was in accordance with Suzuki (2022). Carabid species are categorized into forest and nonforest
10 species according to the literature (Hori 2001; Hori 2003; Hori 2012; Yamanaka et al. 2021). Silphid species (*Silpha perforata*) are
11 categorized as forest species according to Katakura et al. (1986). *Pterostichus microcephalus* is a wing polymorphic species (Shibuya et
12 al. 2018), but most of the individuals we captured were brachypterous (only 2 of 18 individuals were macropterous). Thus, P.
13 microcephalus was included in the analysis.

Appendix

Examining the bias of pitfall traps with enclosure experiments and removal sampling

Satoshi YAMANAKA, Yuichi YAMAURA, Kenichi OZAKI

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Appendix 5 Weather factors (soil and air temperature and precipitation)

Appendix 6 Relationship between body size and coefficients of the effects of fences on
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Data sources and scripts for analysis

(stored in different files)

Data source

1. Data of sampled abundance of each species (“Data1.txt”)

2. Data of sampled abundance of *Pterostichus thunbergi* and soil temperature during

26 7 days (“Data2.txt”)

27 3. Data of body size of each species (“Splist.txt”)

28

29 Script for analysis

30 1. Script for HCMs (“Rcode1.R”)

31 2. Script for removal sampling models (“Rcode2.R”)

32 3. Script for classic removal sampling models (“Rcode3.R”)

33 4. Stan code for HCMs (“Stancode1.stan”)

34 5. Stan code for removal sampling models with the whole period data

35 (“Stancode2.stan”)

36 6. Stan code for removal sampling models with each period data

37 (“Stancode3.stan”)

38 7. Stan code for classic removal sampling models (“Stancode4.stan”)

39

Appendix 1 Study site

Our survey was conducted a clear-cut site and a plantation site (Figure S1). At the clear-cut site, a mature plantation of Sakhalin fir was harvested in 2015, and 3 or 4 years had passed at the time of the survey. In clear-cut site, logging residue produced by the harvesting remained (Figure S1a). The unharvested plantation site mainly consists of mature Sakhalin fir.



Figure S1 Study sites

(a) clear-cut site. (b) plantation site.

Appendix 2 Abundance-based hierarchical community models

In order to examine the effects of temporary emigration process (i.e., fence installation) on sampled abundance of each species, we constructed the abundance-based hierarchical community models (HCMs) according to Yamaura *et al.* (2016). HCMs can estimate a community-level response based on species-level responses.

Here, we pooled the sampled abundance of each species at 4 survey periods. We assumed that the total sampled abundance of species i at grid j (Y_{ij}) follows a negative binomial distribution:

$$Y_{ij} \sim \text{NegativeBinomial}(\lambda_{ij}, \sigma_i), \quad (1)$$

where λ_{ij} is an expected sampled abundance of species i in grid j , and σ_i is a dispersion parameter of species i . We modeled λ_{ij} as a function of an intercept (α_i), habitat type (binary variable H_j ; 0: clear-cut, 1: plantation), installed fence (binary variable F_j ; 0: absent, 1: present), and the interaction of habitat type and installed fence variables. We also included the log-transformed number of the valid traps as an offset term because several traps without fences were disturbed by mammals:

$$\log(\lambda_{ij}) = \alpha_i + \beta_{h_i} \times H_j + \beta_{f_i} \times F_j + \beta_{hf_i} \times H_j \times F_j + \log(T_j), \quad (2)$$

where β_{h_i} , β_{f_i} , and β_{hf_i} are coefficients of habitat type, installed fence, and the interaction of habitat type and installed fence variables, respectively. These coefficients indicate the changes in sampled abundance in unharvested plantation site compared to clear-cut site, abundance changes due to fence installation, and abundance changes due to fence installation in plantation site, respectively. T_j is the number of valid traps in grid j , and $\log(T_j)$ is the offset term.

In order to estimate species- and community-level responses, we assumed that the species level parameters (α_i , β_{h_i} , β_{f_i} , and β_{hf_i}) follow normal distributions shared by all species,

$$\begin{aligned} \alpha_i &\sim \text{Normal}[\mu_\alpha, \sigma_\alpha^2], \\ \beta_{h_i} &\sim \text{Normal}[\mu_{\beta_h}, \sigma_{\beta_h}^2], \dots \end{aligned} \quad (3)$$

where, μ_α is the mean value of α_i , showing the community-level response (i.e., hyperparameter), and σ_α is its standard deviation.

We used normal distributions (0, 100²) for μ_α , μ_{β_h} , μ_{β_f} and $\mu_{\beta_{hf}}$, and half-Cauchy distributions (0, 5) for the standard deviation of each coefficient (σ_α , σ_{β_h} , σ_{β_f} , and $\sigma_{\beta_{hf}}$) and dispersion parameter (σ_i) as weakly informative prior distributions (Gelman 2006). We used four chains that had 20,000 iterations following 5,000 burn-in with 20 thinning rate.

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Appendix 3 Explanation of removal sampling models

1. Removal sampling models considering the changes in removal probability

We constructed and fitted removal sampling models using three methods: (1) the R package “unmarked” version 1.0.1 (Fiske & Chandler 2011) with a Poisson distribution (hereafter, unmarked Poisson model), (2) the R package “ubms” version 1.0.2 (Kellner et al. 2022), and (3) the custom model code in the R package “RStan” (Stan Development Team 2020). As well as unmarked models in the main text, we used two types of data: (1) the whole period data, which covered the four survey periods, and (2) each period’s data.

1-1. unmarked Poisson models

We used the daily number of individuals collected as the response variable, following Poisson distributions, and categorical habitat type (clear-cut vs. unharvested plantation) and standardized daily mean soil temperature as explanatory variables for initial abundance and removal probability, respectively. We used the “gmultmix” function of the R package “unmarked” using maximum likelihood estimation (Fiske & Chandler 2011).

1-2. ubms models

We used the “stan_multinomPois” function of the R package “ubms” using Bayesian estimation and included random effects in the models (Kellner et al. 2022). The response variables and explanatory variables are the same as unmarked models; however, as random effects, we added grid ID and period ID to explanatory variables for the initial abundance in the model with the whole period data and grid ID

to an explanatory variable in the models with each period data. We assumed that the response variables followed a Poisson distribution because “ubms” does not support a negative binomial distribution for removal sampling models. For each model, we used four chains that had 4,000 iterations with no burn-in and one thinning rate.

1-3. RStan model

We developed custom codes according to Kéry and Royle (2015). In the models, the response variables followed a Poisson distribution, and we included grid and/or period IDs as in the ubms models. It is noted that the RStan models had the same structure as the ubms models. For the model using the data set of the whole survey period, we modeled expected value of the initial abundance of *Pterostichus thunbergi* at grid i and period j (λ_{ij}) as a function of habitat types and random effects of sampling grids and survey periods:

$$\log(\lambda_{ij}) = \alpha_i + \varepsilon_{Grid_i} + \varepsilon_{Period_j}, \quad (1)$$

where, α_i is a categorical variable showing the initial abundance in clear-cut or plantation sites. $\varepsilon_{Transect_i}$ and ε_{Period_j} are random effects to account for the variation among sampling grids and survey periods and were assumed to follow a normal distribution:

$$\begin{aligned} \varepsilon_{Grid_i} &\sim \text{Normal}(0, \sigma_{Grid}^2), \\ \varepsilon_{Period_j} &\sim \text{Normal}(0, \sigma_{Period}^2), \end{aligned} \quad (2)$$

where, σ indicates the standard deviation.

We assumed that the removal probability on day t in grid i at period j is determined by the standardized daily mean value of soil temperature on day t in grid i at period j (S_{ijt}):

$$\text{logit}(p_{ijt}) = \alpha_p + \beta_p \times S_{ijt}, \quad (3)$$

where, p_{ijt} is the removal probability on day t in grid i at period j , and its value is ranged from 0 to 1. α_p and β_p indicate an intercept and a coefficient of soil temperature, respectively. We calculated the conditional cell probabilities of multinomial models on each sampling day (π_t). The conditional cell probability on a certain day depends on the probability that the individual was not removed before that day and the removal probability on that day. For example, the conditional cell probability on day 2 (π_{ij2}) depends on the probability that the individuals were not removed on day 1 ($1 - p_{ij1}$) and the removal probability on day 2 (p_{ij2}):

$$\begin{aligned}\pi_{ij1} &= p_{ij1}, \\ \pi_{ij2} &= (1 - p_{ij1}) \times p_{ij2}, \dots, \\ \pi_{ij7} &= (1 - p_{ij1}) \times (1 - p_{ij2}) \times (1 - p_{ij3}) \times (1 - p_{ij4}) \times (1 - p_{ij5}) \times \\ &\quad (1 - p_{ij6}) \times p_{ij7}.\end{aligned}\tag{4}$$

We summed the conditional cell probability of each sampling day to calculate the total removal probability ($\sum_{t=1}^7 \pi_{ijt}$), which indicates the net probability that the individual in grid i at period j was removed in any of 7 sampling days. In addition, π_{ij0} is the probability that the individual in grid i at period j was not removed during sampling days and is shown as $1 - \sum_{t=1}^7 \pi_{ijt}$.

We assumed that sampled abundance on day t in grid i at period j (Y_{ijt}) is determined by the total sampled abundance (n_{ij}) and rate of collected individuals on day t out of the total sampled abundance ($\pi_{ijt}^c = \pi_{ijt}/(1 - \pi_{ij0})$). We also assume that total sampled abundance (n_{ij}) depends on the initial abundance (N_{ij}) and total removal probability ($1 - \pi_{ij0}$), and that the initial abundance follows a Poisson distribution:

$$\begin{aligned}Y_{ijt}|n_{ij} &\sim \text{Multinom}(n_{ij}, \pi_{ijt}^c), \\ n_{ij} &\sim \text{Binomial}(N_{ij}, 1 - \pi_{ij0}), \\ N_{ij} &\sim \text{Poisson}(\lambda_{ij}).\end{aligned}\tag{5}$$

As prior distributions, we used normal distributions $(0, 100^2)$ for α_i , α_p , and β_p . We also used half-Cauchy distributions $(0, 5)$ for any variance of coefficients as weakly informative prior distributions (Gelman 2006). For each model, we used four chains that had 50,000 iterations following 25,000 burn-in with 20 thinning rate. We used the “R-hat” and “N_eff/N” statistics (R-hat < 1.1 and N_eff/N > 0.1) to assess model convergence.

We also constructed models using the data set of each survey period. In the models, we assumed that the expected initial abundance of *P. thunbergi* at grid i (λ_i) was a function of habitat types and random grid effects:

$$\log(\lambda_i) = \alpha_i + \varepsilon_{Transect_i}. \quad (1)'$$

2. Classic removal sampling models

We also constructed classic removal sampling models with four model-fitting methods (unmarked with a Poisson and negative-binomial distribution, ubms, and RStan models). These models were simpler than the above-mentioned models because the classic models do not consider effects of daily soil temperature on removal probability (“Rcode3.R” and “Stancode4.stan” in Script for analysis), meaning that in these models the daily removal probability is assumed to be constant through all survey periods and among sites. We pooled the sampled abundance across survey periods, and the total of sampled abundance on each sampling day and at each grid was used as the response variables.

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Appendix 4 Sampled abundance of *Pterostichus thunbergi* within fences

We collected totally 203 individuals of *P. thunbergi* within fences during each 7-sampling day within 4 periods (Figure S4-1). The sampled abundance was highest at period 1, June 2018 (111 individuals), and the abundance at the other periods was relatively small. In fact, 34, 33 and 25 individuals were caught at period 4, period 3, and period 2, respectively (Figure S4-1). At period 1 and 2, the daily sampled abundance was highest on the first day tended to decrease after the following days. On the other hand, at period 3 and 4, the sampled abundance was highest on the third day and maintained after the following days (Table S4 and Figure S4-1).

In period 3 and period 4, we conducted extended sampling after day 7 and additionally collected a total of 5 and 42 individuals of *P. thunbergi*, respectively (Figure S4-2), suggesting that in these periods (especially period 4), the 7-day sampling did not remove all the individuals within fences. These results were consistent with the results of the removal sampling models showing that the estimated removal probability was overly low in periods 3 and 4 (see Figure 3b and S7-1b in Appendix 7).

Table S4 Mean value of daily sampled abundance of *Pterostichus thunbergi* at each period

Day	Period			
	Period 1	Period 2	Period 3	Period 4
Day 1	5.5 (1-16)	2.5 (0-7)	0.8 (0-2)	0.7 (0-2)
Day 2	1.7 (0-4)	0.5 (0-3)	0.7 (0-3)	1.0 (0-2)
Day 3	4.0 (0-9)	0.2 (0-1)	1.3 (0-5)	1.5 (0-4)
Day 4	2.5 (0-5)	0.5 (0-2)	0.8 (0-3)	0.5 (0-2)

Day 5	2.5 (0-6)	0.3 (0-1)	0.8 (0-4)	0.3 (0-2)
Day 6	1.2 (0-3)	0.2 (0-1)	0.7 (0-2)	0.7 (0-3)
Day 7	1.2 (0-3)	0.0 (0-0)	0.5 (0-1)	0.8 (0-2)

Values indicate the mean value of daily sampled abundance and using 9 pitfall traps.

Values with parenthesis indicate the max and minimum values of the sampled abundance.

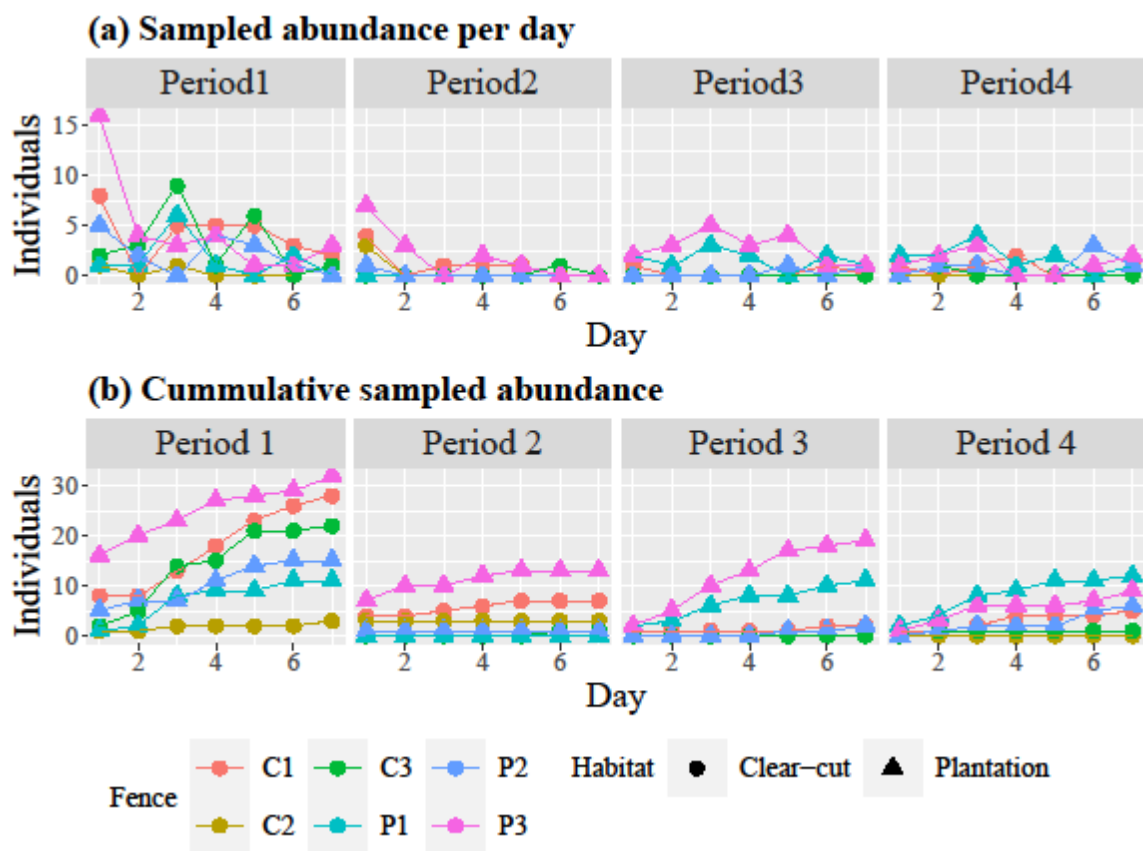


Figure S4-1 Sampled abundance (a) and cumulative sampled abundance (b) of *P.*

thunbergi

Circles and triangles indicate clear-cut and unharvested plantation sites, respectively.

Different colors indicate different fences.

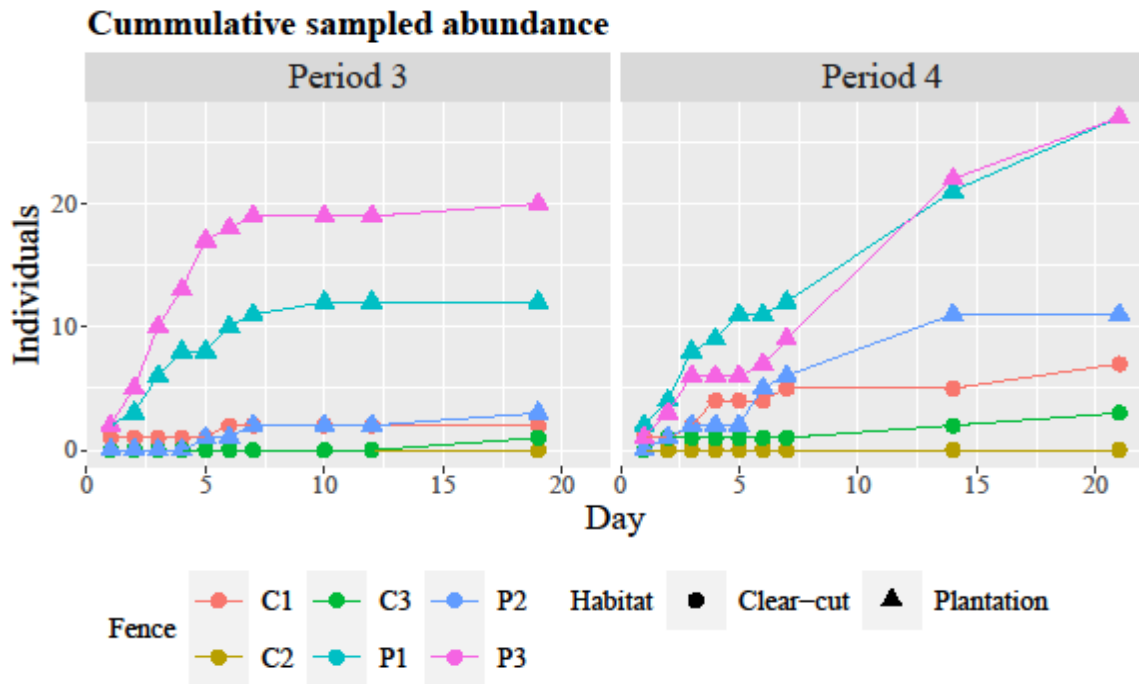


Figure S4-2 Cumulative sampled abundance of *P. thunbergi* on days 1-19 and 1-21 in periods 3 and 4

Circles and triangles indicate clear-cut and unharvested plantation sites, respectively.

Different colors indicate different fences. In periods 3 and 4, the survey was continued after the 7th day, and samples were collected on days 10, 12, and 19 in period 3 and on days 14 and 21 in period 4. We collected 5 and 42 individuals of *P. thunbergi* additionally, respectively.

Appendix 5 Weather factors (soil and air temperature and precipitation)

Daily mean values of soil temperature were higher in the clear-cut site than at the unharvested plantation site, and standard deviations were also relatively higher in the clear-cut site (Table S5). The ranges of daily mean and maximum values of soil temperature were overlapped among survey periods, although the minimum values in period 3 were relatively lower than those in other periods (Figure S5-1). The daily mean values of soil temperature were correlated with its maximum and minimum values, respectively (Figure S5-2).

Comparing the mean values of soil temperature among the all-day, day, and night time, the difference in soil temperature in the all-day time reflected the soil temperature in day time (from 6:00 to 18:00), and the difference among fences was small in night time (from 18:00 to 6:00) (Figure S5-3). We also examined air temperature at several fences. The daily mean values of air temperature had similar patterns with those of soil temperature (Figure S5-4).

Table S5 Daily mean values and its standard deviation of soil temperature

Period	Clear-cut		Plantation	
	Mean	SD	Mean	SD
Period1 2018 June	18.90	1.33	17.38	2.73
Period2 2018 July	25.04	1.54	22.15	1.34
Period3 2019 June	20.95	2.98	15.74	1.14
Period4 2019 July	25.79	1.48	22.00	0.67

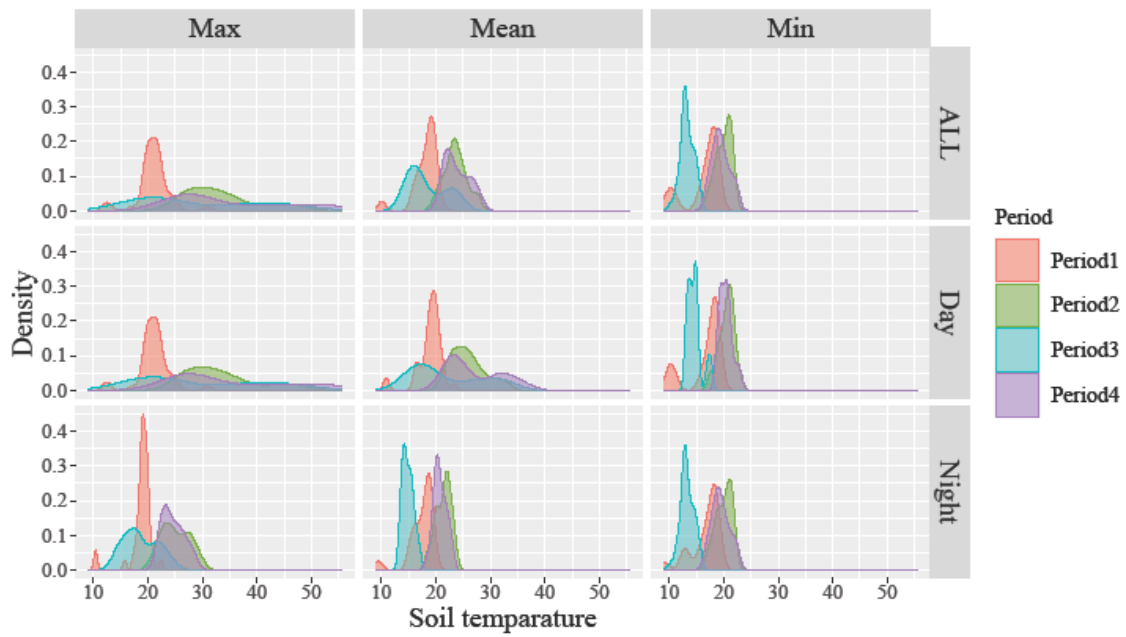


Figure S5-1 Distribution of mean values of soil temperature

All: all-day time (24h), Day: day time (6:00-18:00, 12h), Night: night time (18:00-6:00, 12h)

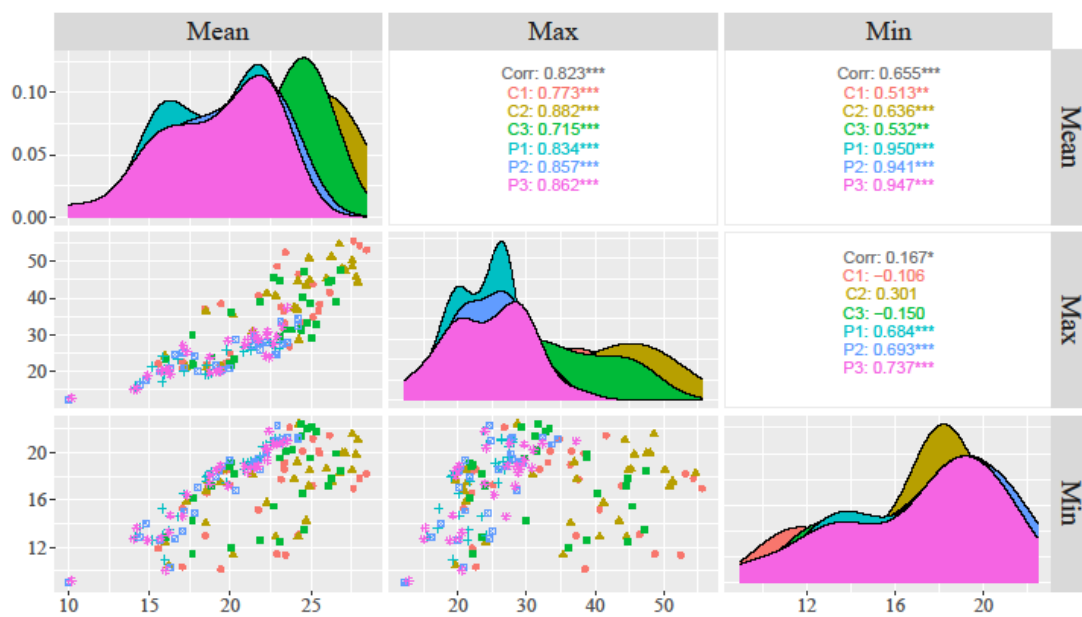


Figure S5-2 Correlation of mean values of soil temperature in the all-day time with maximum and minimum values

Different symbols indicate each fence. The values at the right panels indicate the correlation coefficients.

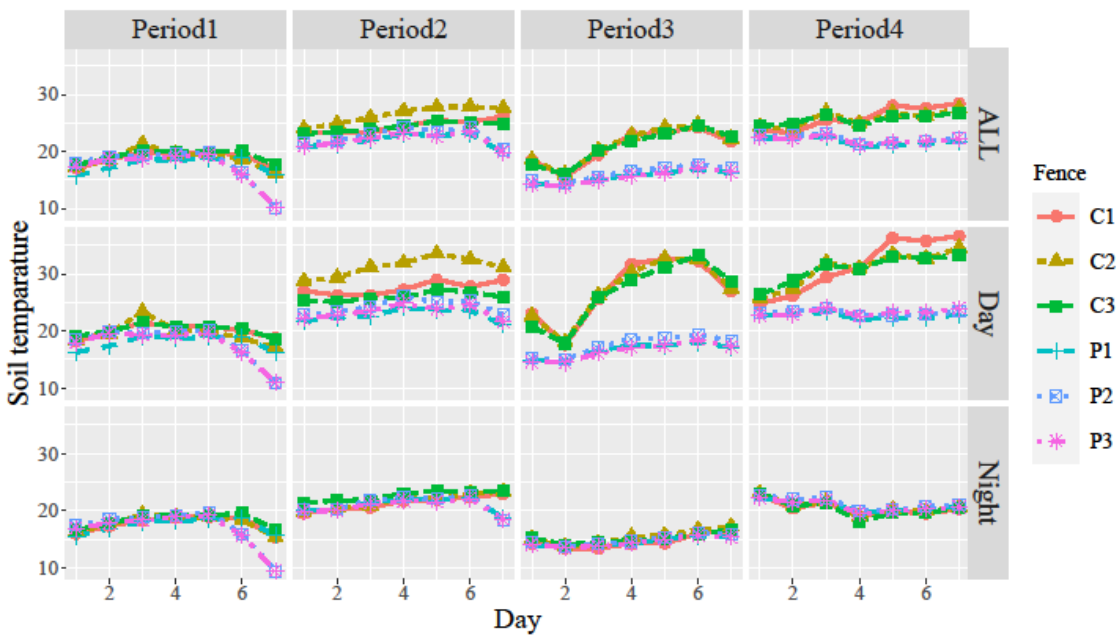


Figure S5-3 Mean values of soil temperature at the all-day, day, and night time.

All: all-day time (24h), Day: day time (6:00-18:00, 12h), Night: night time (18:00-6:00, 12h). Different symbols indicate each fence.

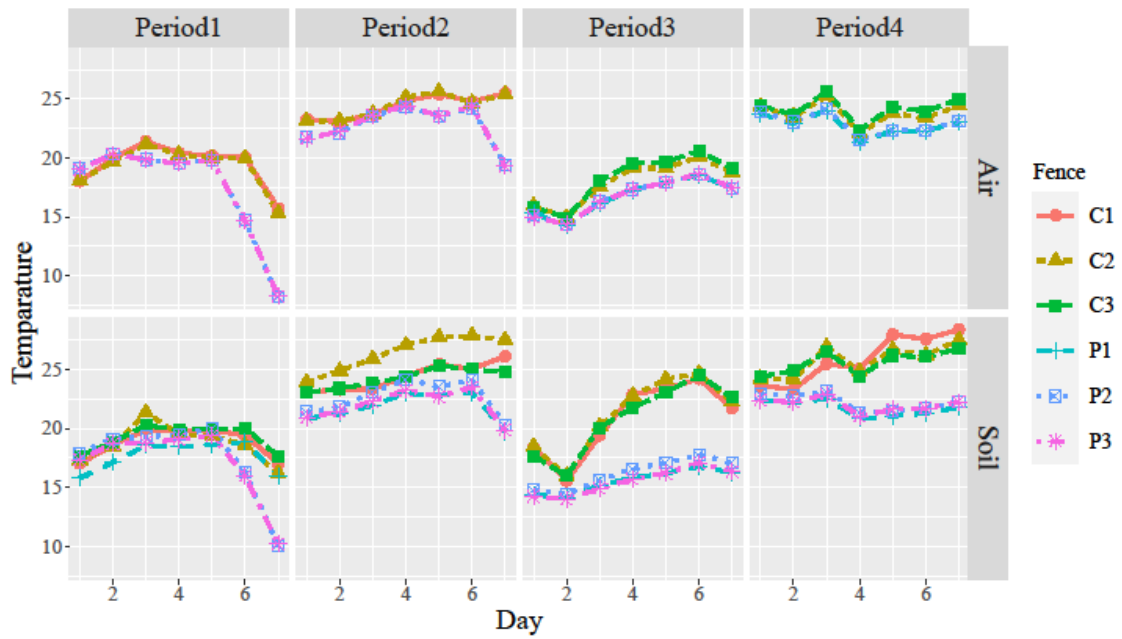


Figure S5-4 Daily mean values of air and soil temperature

Values indicate daily mean values. Different symbols indicate each fence.

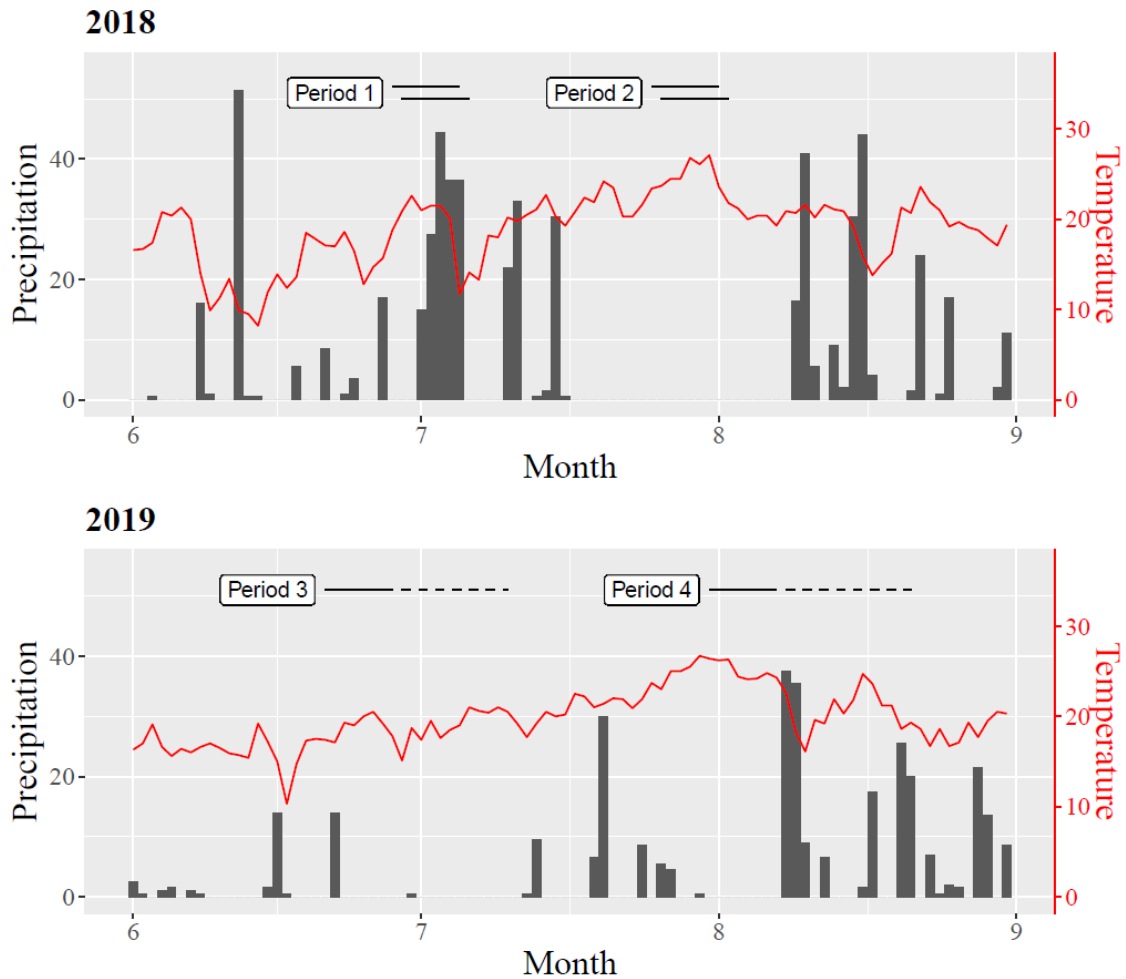


Figure S5-5 Daily total precipitation and mean temperature of the study area during the surveyed period

Black bars and red lines indicate daily total precipitation and mean temperature, respectively. These data were obtained from the Ashibetsu observatory (Japan Meteorological Agency 2021), which is the closest study area. Horizontal solid and dashed lines indicate 7-day continuous sampling (at periods 1-4) and extended sampling (at only periods 3-4), respectively. In period 1, sampling at grids C1, C2, C3, and P1 started on 28 June 2018 and ended on 5 July 2018, and sampling at grids P2 and P3 started on 29 June 2018 and ended on 6 July 2018. In period 2, sampling at Clear-cut site (C1, C2, and C3) and Plantation site (P1, P2, and P3) started on 25 and 26 July

289 2018 and ended on 1 and 2 August 2018, respectively. In period 3, sampling started on
290 21 June 2019 and ended on 10 July 2019. In period 4, sampling started on 31 July 2019
291 and ended on 21 August 2019.

292

293 Reference

294 Japan Meteorological Agency (2021) Historical weather data,
295 <http://www.data.jma.go.jp/obd/stats/etrn/index.php>.

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Appendix 6 Relationship between body size and coefficients of the effects of fences on sampled abundance

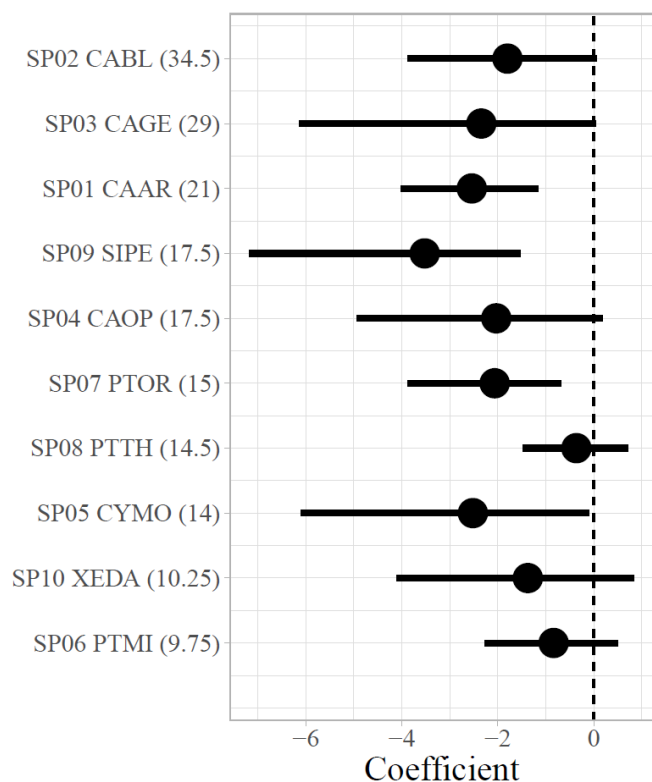


Figure S6 body size and coefficients of the effects of fences on sampled abundance of each species

Y-axis indicates each species, sorted from top to bottom by body size (numbers in parentheses next to species names represent the mean body length). X-axis indicates coefficient of the effects of fences on sampled abundance. Black circles indicate mean values, and horizontal lines represent 95% credible intervals of the coefficients. Mean body length of each species was calculated by averaging the maximum and minimum sizes that reported by Ueno, Kurosawa and Sato (1985).

311 Ueno, S., Kurosawa, Y. & Sato, M. (1985) *The Coleoptera of Japan in Color, Vol. II (in*
312 *Japanese)*. Hoikusha, Osaka.
313

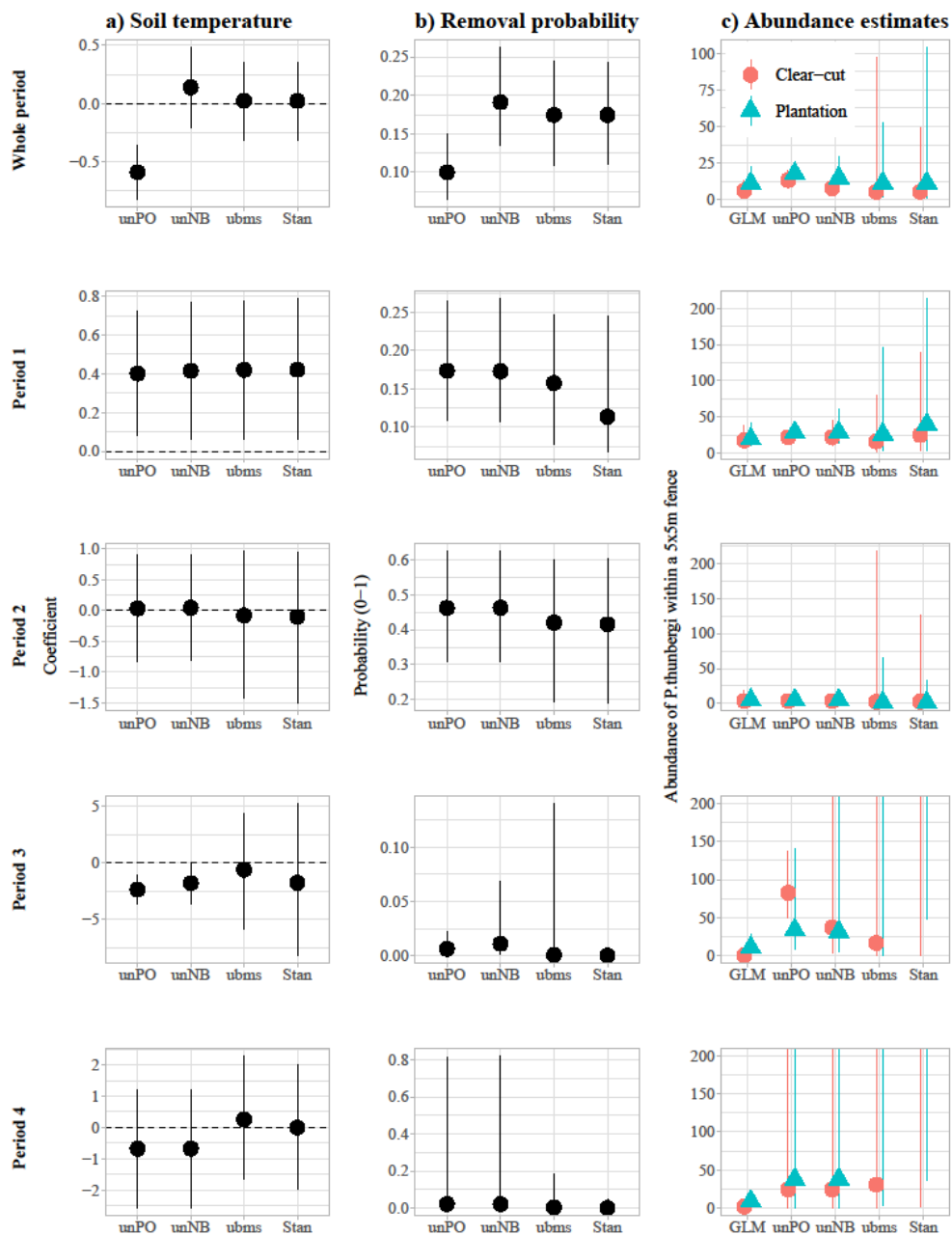
Appendix 7 Results of removal sampling models

1. the coefficients of removal sampling models with different model-fitting methods

The `ubms` and `RStan` models converged; however, a poor model fit for all `ubms` models was suggested based on Pareto smoothed importance-sampling leave-one-out cross-validation (PSIS-LOO). Most of the estimated coefficients were similar among the different model-fitting methods (Figure S7-1 and S7-3). As well as `unmarked` models in the main text, in other models in periods 3 and 4, the estimated removal probability and the initial abundance estimates were overly low and high, respectively (Figure S7-1c and S7-2). This suggests that these models in these periods also could not estimate the removal probability and initial abundance precisely. Unlike other models, in the `unmarked` Poisson model, the effect of soil temperature was negative for the whole period data (Figure S7-1a) likely due to the coefficient of period 3.

2. Comparing classic removal sampling models

In the result of classic removal sampling models, estimated coefficients did not differ among model-fitting methods (`unmarked`, `ubms`, and `RStan`) (Figure S7-4). Moreover, the coefficients did not largely differ from that of the abovementioned removal sampling models with the whole period data (Figure S7-3 and S7-4). This suggests that effects of daily soil temperature on removal probability were not large when considering over the whole survey periods, and this was consistent with our main conclusion.



338

339 Figure S7-1 Comparison of estimates of four removal sampling models and GLMs with
 340 the whole period data (Whole period) and each period data (Periods 1, 2, 3, and 4) in the
 341 analysis of *P. thunbergi*

The mean values of estimates are denoted by circles or triangles, and their 95% CIs are denoted by vertical bars. (a) Effects of daily soil temperature on removal probability, (b) estimated per-day removal probability (given the mean soil temperature), and (c) abundance estimates are shown separately for different methods and distributions. Abbreviation: unmarked Poisson (unPO) and negative binomial (unNB), ubms (ubms), and RStan model (Stan). It is noted that unNB is the unmarked models in the main text. In periods 3 and 4, some mean values and their CIs were overly large, and thus, the figures show estimates with a limited range on the y-axis (0-200) to clarify the difference in estimates among models. Figures with the complete range of the y-axis are shown in Figure S7-2.

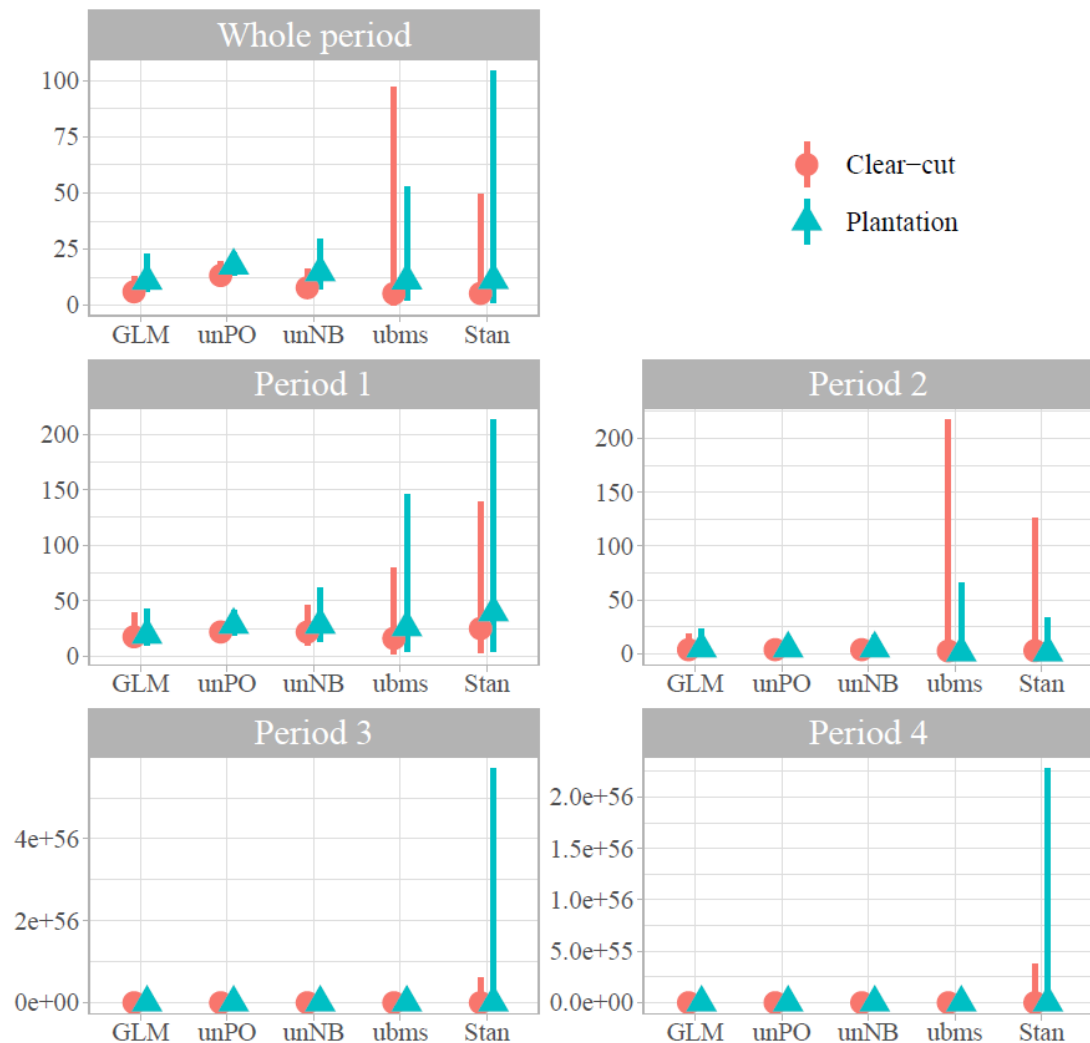
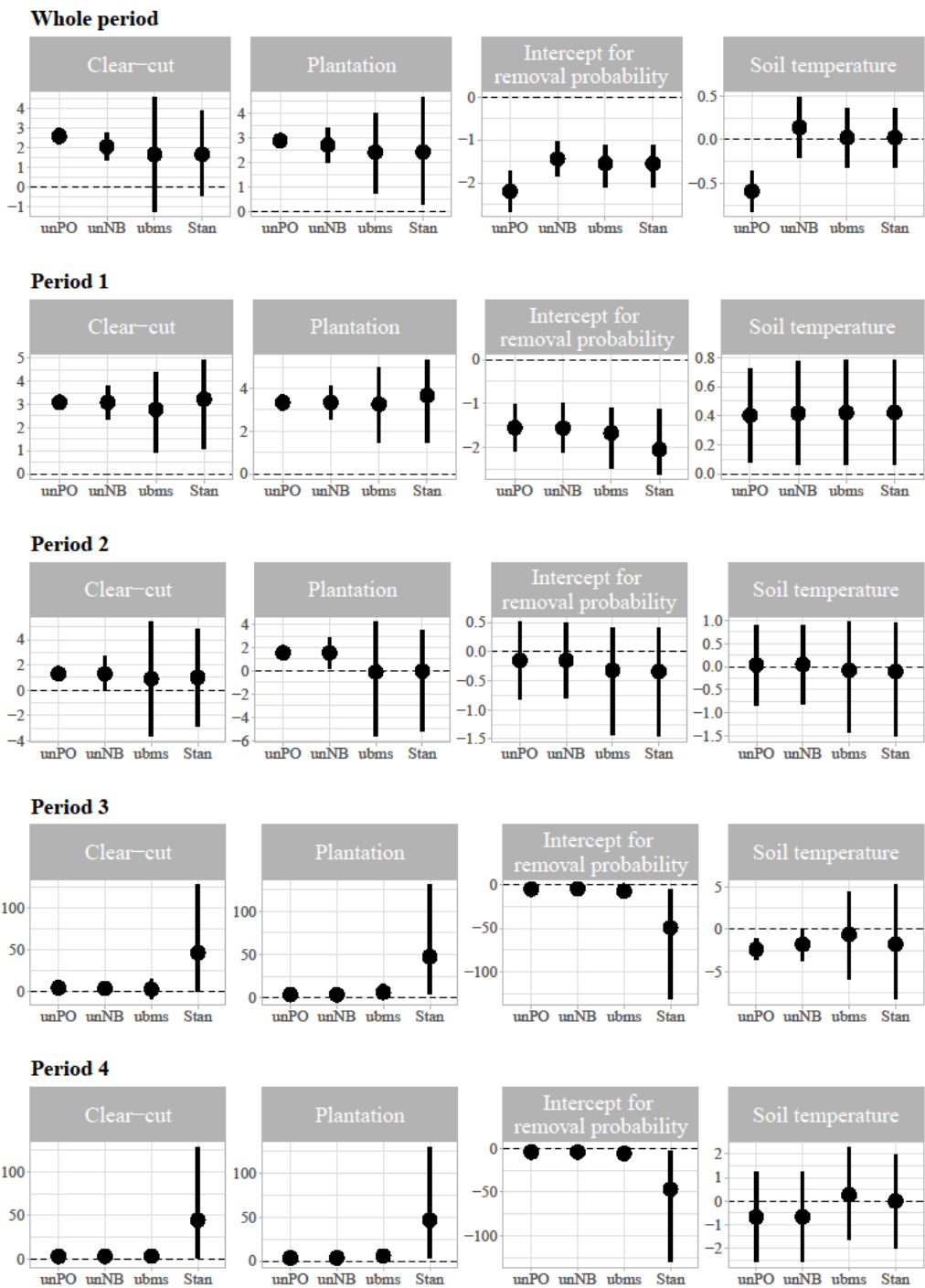


Figure S7-2 Initial abundances estimated by removal sampling models and expected sampled abundance by GLMs (with full-range y-axis)

The estimates of GLM indicate the expected sampled abundance estimated by GLMs. Other estimates (unPO, unNB, ubms, and Stan) indicate the initial abundance estimated by each removal sampling model. unPO and unNB indicate unmarked model with a Poisson and negative binomial distribution, respectively. Red circles and blue triangles indicate the mean values of clear-cut and unharvested plantation sites, respectively.

360 Vertical bars show 95% CIs of estimates for GLMs and unmarked models or ubms
 361 and RStan models, respectively.

362



363

364 Figure S7-3 Estimated coefficients of parameters of removal sampling models

Black circles indicate mean values of the coefficients. Vertical bars show 95% confidence intervals or credible intervals of the coefficients for unmarked models or ubms and RStan models, respectively. Clear-cut and plantation indicate the initial abundance of clear-cut and unharvested plantation sites, respectively. Soil temperature indicates the coefficients of the effects of daily soil temperature on removal probability. unPO and unNB indicate unmarked models with a Poisson and negative binomial distribution, respectively.

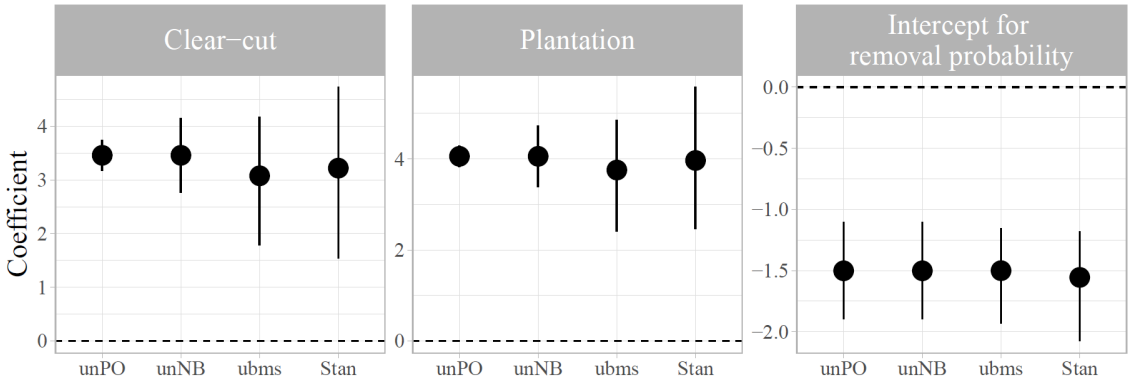


Figure S7-4 Estimated coefficients of parameters of classic removal sampling models (without considering the daily changes in removal probability)

Black circles indicate mean values of the coefficients. Vertical bars show 95% CIs of the coefficients for unmarked models or ubms and RStan models, respectively. Clear-cut and plantation indicate the initial abundance of clear-cut and unharvested plantation sites, respectively. unPO and unNB indicate unmarked models with a Poisson and negative binomial distribution, respectively.